## Appendix

### Model

The number of cases, Y, for each observation, i, for every hospital, h, follows a Poisson distribution with mean lambda[i,h].

 $Y[i,h] \sim Poisson(lambda[i,h])$ 

The log of the mean Poisson rate, with patient-days as the offset, is modelled as the sum of a random year effect, random period effect, and a hospital effect. The hierarchical modelling of the year and hospital effects allow for the year effect to vary by hospital with a variance denoted by sigma.year.

log(lambda[i,h]) = log(patient - days[i,h]) + mean[i,h]

mean[i,h] = year.hospital[s,h] + period.effect[p]

year.hospital[s,h]~Normal(year.effect[s] + hospital.effect[h], sigma.year)

The mean of the hospital effect is modelled as a function of an overall intercept, the type of hospital (A; non-teaching without ICU, B; non-teaching with ICU, and C; teaching with ICU) and the number of beds in the hospital, as well as a random effect relating the calendar year the hospital entered BACTOT, where g is a pointer to hospital h's first year in the BACTOT.

hospital.effect[h] = intercept + beta.typeB \* typeB[h] + beta.typeC \* typeC[h] + beta.beds \* beds[h] + entry[g[h]]

Below are the prior distributions for the random effects.

entry[g] ~ Normal(0, sigma. entry)
period.effect[p] ~ Normal(0, sigma. period)
year.effect[s] ~ Normal(0, sigma2. year)

#### R code

Below is the R code used to fit the model using JAGS. *nhospital* refers to the total number of hospitals. *group.hosp* refers to the year in which a hospital entered BACTOT. *n[h]* refers to the total number of observations available for each hospital. *nyear* refers to the total number of surveillance years for which a hospital has participated in BACTOT. *Togroup* refers to the number of possible years in which new hospitals entered BACTOT. *toperiod* refers to the maximum number of surveillance periods, and *toyear* refers to the maximum number of surveillance years.

model {

```
## Likelihood function
         for(h in 1:nhospital){
                  level[h] <- beta.c + b.beds*beds[h] + phi[group.hosp[h]] + b.typeB*typeB[h] + b.typeC*typeC[h]
                  risk.hosp[h] <- exp(level[h])
                           for(i in 1:n[h])
      y[i,h] \sim dpois(lambda[i,h])
      log(lambda[i,h]) <- offset[i,h] + mean[i,h]
      mean[i,h] \leq gamma[vear[i,h],h] + delta.c[period[i,h]]
      y.fitted[i,h] ~ dpois(lambda[i,h])
  }
 ## Prior specification
   for(h in 1:nhospital){
     for(ii in 1:nyear[h]){
      gamma[ii,h] ~ dnorm(level[h] + gamma.c[ii], prec.gamma)
      risk.year.hosp[ii,h] <- exp(gamma[ii,h])
     }
   for(i in 1:togroup){
     phi[i] ~ dnorm(0,prec.group)
     risk.group[i] <- exp(phi[i])
    for(p in 1:toperiod){
     delta.c[p] \sim dnorm(0, prec.delta.c)
     risk.period[p] <- exp(delta.c[p])
   for(p in 2:toperiod){
     rr.period[p] \le exp(delta.c[p])/exp(delta.c[1])
  for(i in 1:toyear){
     gamma.c[i] \sim dnorm(0, prec.gamma.c)
     risk.year[i] <- exp(gamma.c[i])
  for(i in 2:toyear){
     rr.year[i] <- exp(gamma.c[i])/exp(gamma.c[1])
  for(i in 2:togroup){
   rr.group[i] <- exp(phi[i])/exp(phi[1])
  }
  beta.c \sim dnorm(0,0.1)
```

```
prec.gamma.c ~ dgamma(2,0.01)
```

sigma.gamma.c <- 1/prec.gamma.c prec.gamma ~ dgamma(2,0.01) sigma.gamma <- 1/prec.gamma prec.delta.c ~ dgamma(2,0.01) sigma.delta.c <- 1/prec.delta.c prec.group ~ dgamma(2,0.01) sigma.group<- 1/prec.group

b.beds ~ dnorm(0,0.1) b.typeB ~ dnorm(0,0.1) b.typeC ~ dnorm(0,0.1)

rr.beds <- exp(b.beds) rr.typeB <- exp(b.typeB) rr.typeC <- exp(b.typeC) Table S1. Healthcare-associated Bloodstream Infection Cases, Patient-days and Pooled Incidence Rates for Each BACTOT

Surveillance Year, Including Hospitals With no Cases.

<u>Surveillance</u> <u>year</u>	<u>1</u>	2	<u>3</u>	<u>4</u>	<u>5</u>	<u>6</u>	7	<u>8</u>	<u>9</u>	<u>10</u>	<u>Overall</u>
<u>Hospitals</u>	<u>79</u>	<u>79</u>	<u>79</u>	<u>67</u>	<u>56</u>	<u>53</u>	<u>51</u>	<u>51</u>	<u>46</u>	<u>40</u>	<u>79</u>
<u>Cases</u>	<u>2169</u>	<u>2248</u>	<u>2038</u>	<u>1942</u>	<u>1788</u>	<u>1563</u>	<u>1479</u>	<u>1356</u>	<u>1323</u>	<u>1275</u>	<u>17181</u>
Patient-days	<u>4210728</u>	4242444	<u>4197734</u>	<u>3618735</u>	<u>3258880</u>	<u>2986131</u>	2762369	2738324	<u>2506465</u>	<u>2278894</u>	<u>32800704</u>
<b>Pooled</b>	<u>5.15</u>	<u>5.30</u>	<u>4.86</u>	<u>5.37</u>	<u>5.49</u>	<u>5.23</u>	<u>5.35</u>	<u>4.95</u>	<u>5.28</u>	<u>5.59</u>	<u>5.24</u>
Incidence Rate	<u>(4.94-5.37)</u>	<u>(5.08-5.52)</u>	<u>(4.65-5.07)</u>	<u>(5.13-5.61)</u>	<u>(5.24-5.75)</u>	<u>(4.98-5.5)</u>	<u>(5.09-5.63)</u>	<u>(4.7-5.22)</u>	<u>(5.00-5.57)</u>	<u>(5.30-5.91)</u>	<u>(5.16-5.32)</u>
<u>(95% CI)</u>											

95% CI; 95% confidence interval



Figure S1. Posterior summaries of the incidence rate ratios of healthcare-associated bloodstream infections and its most frequent subtypes in surveillance years 2-10 relative to surveillance year 1, from all included hospitals. Dots represent the mean and the lines represent the 95% credible interval. HABSI; Healthcare-associated bloodstream infection, CA-BSI; catheter-associated bloodstream

# infection, NCA-BSI; non-catheter associated primary bloodstream infection, BSI-UTI; bloodstream infection secondary to a urinary

## tract infection.



**Figure S2.** Posterior summaries of the period component of healthcare-associated bloodstream infections and its most frequent subtypes incidence rates from all included hospitals. Dots represent the mean and the lines represent the 95% credible interval. <u>HABSI; Healthcare-associated bloodstream infection, CA-BSI; catheter-associated bloodstream infection, NCA-BSI; non-catheter associated primary bloodstream infection, BSI-UTI; bloodstream infection secondary to a urinary tract infection.</u>