

Supplementary Appendix to  
“Closed-Form Identification of Dynamic Discrete Choice Models  
with Proxies for Unobserved State Variables”

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This supplementary appendix presents proofs for the identification results under extended models (Sections B.1–B.4) and discuss the consistency of the analog estimator for the law of state transition (Section C).

## B Additional Proofs

### B.1 Proof of Theorem 2

Our closed-form identification includes four steps.

**Step 1: Closed-form identification of the transition rule  $f_3(x_t^*|d_{t-1}, w_{t-1}, x_{t-1}^*)$ :** First, we show the identification of the parameters and the distributions in transition of  $x_t^*$ . Since

$$\begin{aligned} x_t &= x_t^* + \varepsilon_t = \sum_d \mathbb{1}\{d_{t-1} = d\}[\alpha^d + \beta^d w_{t-1} + \gamma^d x_{t-1}^* + \eta_t^d] + \varepsilon_t \\ &= \sum_d \mathbb{1}\{d_{t-1} = d\}[\alpha^d + \beta^d w_{t-1} + \gamma^d x_{t-1} + \eta_t^d - \gamma^d \varepsilon_{t-1}] + \varepsilon_t \end{aligned}$$

we obtain the following equalities for each  $d$ :

$$\begin{aligned}
\mathbb{E}[x_t \mid d_{t-1} = d] &= \alpha^d + \beta^d \mathbb{E}[w_{t-1} \mid d_{t-1} = d] + \gamma^d \mathbb{E}[x_{t-1} \mid d_{t-1} = d] \\
&\quad - \mathbb{E}[\gamma^d \varepsilon_{t-1} \mid d_{t-1} = d] + \mathbb{E}[\eta_t^d \mid d_{t-1} = d] + \mathbb{E}[\varepsilon_t \mid d_{t-1} = d] \\
&= \alpha^d + \beta^d \mathbb{E}[w_{t-1} \mid d_{t-1} = d] + \gamma^d \mathbb{E}[x_{t-1} \mid d_{t-1} = d] \\
\mathbb{E}[x_t w_{t-1} \mid d_{t-1} = d] &= \alpha^d \mathbb{E}[w_{t-1} \mid d_{t-1} = d] + \beta^d \mathbb{E}[w_{t-1}^2 \mid d_{t-1} = d] + \gamma^d \mathbb{E}[x_{t-1} w_{t-1} \mid d_{t-1} = d] \\
&\quad - \mathbb{E}[\gamma^d \varepsilon_{t-1} w_{t-1} \mid d_{t-1} = d] + \mathbb{E}[\eta_t^d w_{t-1} \mid d_{t-1} = d] + \mathbb{E}[\varepsilon_t w_{t-1} \mid d_{t-1} = d] \\
&= \alpha^d \mathbb{E}[w_{t-1} \mid d_{t-1} = d] + \beta^d \mathbb{E}[w_{t-1}^2 \mid d_{t-1} = d] + \gamma^d \mathbb{E}[x_{t-1} w_{t-1} \mid d_{t-1} = d] \\
\mathbb{E}[x_t w_t \mid d_{t-1} = d] &= \alpha^d \mathbb{E}[w_t \mid d_{t-1} = d] + \beta^d \mathbb{E}[w_{t-1} w_t \mid d_{t-1} = d] + \gamma^d \mathbb{E}[x_{t-1} w_t \mid d_{t-1} = d] \\
&\quad - \mathbb{E}[\gamma^d \varepsilon_{t-1} w_t \mid d_{t-1} = d] + \mathbb{E}[\eta_t^d w_t \mid d_{t-1} = d] + \mathbb{E}[\varepsilon_t w_t \mid d_{t-1} = d] \\
&= \alpha^d \mathbb{E}[w_t \mid d_{t-1} = d] + \beta^d \mathbb{E}[w_{t-1} w_t \mid d_{t-1} = d] + \gamma^d \mathbb{E}[x_{t-1} w_t \mid d_{t-1} = d]
\end{aligned}$$

by the independence and zero mean assumptions for  $\eta_t^d$  and  $\varepsilon_t$ . From these, we have the linear equation

$$\begin{bmatrix} \mathbb{E}[x_t \mid d_{t-1} = d] \\ \mathbb{E}[x_t w_{t-1} \mid d_{t-1} = d] \\ \mathbb{E}[x_t w_t \mid d_{t-1} = d] \end{bmatrix} = \begin{bmatrix} 1 & \mathbb{E}[w_{t-1} \mid d_{t-1} = d] & \mathbb{E}[x_{t-1} \mid d_{t-1} = d] \\ \mathbb{E}[w_{t-1} \mid d_{t-1} = d] & \mathbb{E}[w_{t-1}^2 \mid d_{t-1} = d] & \mathbb{E}[x_{t-1} w_{t-1} \mid d_{t-1} = d] \\ \mathbb{E}[w_t \mid d_{t-1} = d] & \mathbb{E}[w_{t-1} w_t \mid d_{t-1} = d] & \mathbb{E}[x_{t-1} w_t \mid d_{t-1} = d] \end{bmatrix} \begin{bmatrix} \alpha^d \\ \beta^d \\ \gamma^d \end{bmatrix}$$

Provided that the matrix on the right-hand side is non-singular, we can identify the parameters

$(\alpha^d, \beta^d, \gamma^d)$  by

$$\begin{bmatrix} \alpha^d \\ \beta^d \\ \gamma^d \end{bmatrix} = \begin{bmatrix} 1 & \mathbb{E}[w_{t-1} \mid d_{t-1} = d] & \mathbb{E}[x_{t-1} \mid d_{t-1} = d] \\ \mathbb{E}[w_{t-1} \mid d_{t-1} = d] & \mathbb{E}[w_{t-1}^2 \mid d_{t-1} = d] & \mathbb{E}[x_{t-1} w_{t-1} \mid d_{t-1} = d] \\ \mathbb{E}[w_t \mid d_{t-1} = d] & \mathbb{E}[w_{t-1} w_t \mid d_{t-1} = d] & \mathbb{E}[x_{t-1} w_t \mid d_{t-1} = d] \end{bmatrix}^{-1} \begin{bmatrix} \mathbb{E}[x_t \mid d_{t-1} = d] \\ \mathbb{E}[x_t w_{t-1} \mid d_{t-1} = d] \\ \mathbb{E}[x_t w_t \mid d_{t-1} = d] \end{bmatrix}$$

Next, we show identification of the distributions of  $\varepsilon_t$  and  $\eta_t^d$  for each  $d$ . Observe that

$$\begin{aligned}
&\mathbb{E}[\exp(is_1 x_{t-1} + is_2 x_t) \mid d_{t-1} = d] \\
&= \mathbb{E}[\exp(is_1 (x_{t-1}^* + \varepsilon_{t-1}) + is_2 (\alpha^d + \beta^d w_{t-1} + \gamma^d x_{t-1}^* + \eta_t^d + \varepsilon_t)) \mid d_{t-1} = d] \\
&= \mathbb{E}[\exp(i(s_1 x_{t-1}^* + s_2 \alpha^d + s_2 \beta^d w_{t-1} + s_2 \gamma^d x_{t-1}^*)) \mid d_{t-1} = d] \\
&\quad \times \mathbb{E}[\exp(is_1 \varepsilon_{t-1})] \mathbb{E}[\exp(is_2 (\eta_t^d + \varepsilon_t))]
\end{aligned}$$

follows from the independence assumptions for  $\eta_t^d$  and  $\varepsilon_t$ . Taking the derivative with respect to  $s_2$  yields

$$\begin{aligned}
& \left[ \frac{\partial}{\partial s_2} \ln \mathbb{E} [\exp (i s_1 x_{t-1} + i s_2 x_t) \mid d_{t-1} = d] \right]_{s_2=0} \\
&= \frac{\mathbb{E} [i(\alpha^d + \beta^d w_{t-1} + \gamma^d x_{t-1}^*) \exp (i s_1 x_{t-1}^*) \mid d_{t-1} = d]}{\mathbb{E} [\exp (i s_1 x_{t-1}^*) \mid d_{t-1} = d]} \\
&= i\alpha^d + \beta^d \frac{\mathbb{E}[i w_{t-1} \exp(i s_1 x_{t-1}^*) \mid d_{t-1} = d]}{\mathbb{E}[\exp(i s_1 x_{t-1}^*) \mid d_{t-1} = d]} + \gamma^d \frac{\partial}{\partial s_1} \ln \mathbb{E} [\exp (i s_1 x_{t-1}^*) \mid d_{t-1} = d] \\
&= i\alpha^d + \beta^d \frac{\mathbb{E}[i w_{t-1} \exp(i s_1 x_{t-1}) \mid d_{t-1} = d]}{\mathbb{E}[\exp(i s_1 x_{t-1}) \mid d_{t-1} = d]} + \gamma^d \frac{\partial}{\partial s_1} \ln \mathbb{E} [\exp (i s_1 x_{t-1}^*) \mid d_{t-1} = d]
\end{aligned}$$

where the switch of the differential and integral operators is permissible provided that there exists  $h \in L^1(F_{w_{t-1} x_{t-1}^* | d_{t-1}=d})$  such that  $|i(\alpha^d + \beta^d w_{t-1} + \gamma^d x_{t-1}^*) \exp (i s_1 x_{t-1}^*)| < h(w_{t-1}, x_{t-1}^*)$  holds for all  $(w_{t-1}, x_{t-1}^*)$ , which follows from the bounded conditional moment given in Assumption 10, and the denominators are nonzero as the conditional characteristic function of  $x_t^*$  given  $d_t$  does not vanish on the real line under Assumption 10. Therefore,

$$\begin{aligned}
\mathbb{E} [\exp (i s x_{t-1}^*) \mid d_{t-1} = d] &= \exp \left[ \int_0^s \left[ \frac{1}{\gamma^d} \frac{\partial}{\partial s_2} \ln \mathbb{E} [\exp (i s_1 x_{t-1} + i s_2 x_t) \mid d_{t-1} = d] \right]_{s_2=0} ds_1 \right. \\
&\quad \left. - \int_0^s \frac{i\alpha^d}{\gamma^d} ds_1 - \int_0^s \frac{\beta^d \mathbb{E}[i w_{t-1} \exp(i s_1 x_{t-1}) \mid d_{t-1} = d]}{\gamma^d \mathbb{E}[\exp(i s_1 x_{t-1}) \mid d_{t-1} = d]} ds_1 \right] \\
&= \exp \left[ \int_0^s \frac{\mathbb{E} [i(x_t - \alpha^d - \beta^d w_{t-1}) \exp (i s_1 x_{t-1}) \mid d_{t-1} = d]}{\gamma^d \mathbb{E} [\exp (i s_1 x_{t-1}) \mid d_{t-1} = d]} ds_1 \right].
\end{aligned}$$

From the proxy model and the independence assumption for  $\varepsilon_t$ ,

$$\mathbb{E} [\exp (i s x_{t-1}) \mid d_{t-1} = d] = \mathbb{E} [\exp (i s x_{t-1}^*) \mid d_{t-1} = d] \mathbb{E} [\exp (i s \varepsilon_{t-1})].$$

We then obtain the following result using any  $d$ .

$$\begin{aligned}
\mathbb{E} [\exp (i s \varepsilon_{t-1})] &= \frac{\mathbb{E} [\exp (i s x_{t-1}) \mid d_{t-1} = d]}{\mathbb{E} [\exp (i s x_{t-1}^*) \mid d_{t-1} = d]} \\
&= \frac{\mathbb{E} [\exp (i s x_{t-1}) \mid d_{t-1} = d]}{\exp \left[ \int_0^s \frac{\mathbb{E} [i(x_t - \alpha^d - \beta^d w_{t-1}) \exp(i s_1 x_{t-1}) \mid d_{t-1}=d]}{\gamma^d \mathbb{E}[\exp(i s_1 x_{t-1}) \mid d_{t-1}=d]} ds_1 \right]}.
\end{aligned}$$

This argument holds for all  $t$  so that we can identify the characteristic function of  $\varepsilon_t$  with

$$\mathbb{E}[\exp(is\varepsilon_t)] = \frac{\mathbb{E}[\exp(isx_t) | d_t = d]}{\exp\left[\int_0^s \frac{\mathbb{E}[i(x_{t+1} - \alpha^d - \beta^d w_t) \exp(is_1 x_t) | d_t = d]}{\gamma^d \mathbb{E}[\exp(is_1 x_t) | d_t = d]} ds_1\right]} \quad (\text{B.1})$$

using any  $d$ .

In order to identify  $f_{\eta_t^d}$  for each  $d$ , consider

$$x_t + \gamma^d \varepsilon_{t-1} = \alpha^d + \beta^d w_{t-1} + \gamma^d x_{t-1} + \varepsilon_t + \eta_t^d,$$

and thus

$$\begin{aligned} \mathbb{E}[\exp(isx_t) | d_{t-1} = d] \mathbb{E}[\exp(is\gamma^d \varepsilon_{t-1})] &= \mathbb{E}[\exp(is(\alpha^d + \beta^d w_{t-1} + \gamma^d x_{t-1})) | d_{t-1} = d] \\ &\quad \times \mathbb{E}[\exp(is\eta_t^d)] \mathbb{E}[\exp(is\varepsilon_t)] \end{aligned}$$

follows by the independence assumptions for  $\eta_t^d$  and  $\varepsilon_t$ . Therefore, by the formula (B.1), the characteristic function of  $\eta_t^d$  can be expressed by

$$\begin{aligned} \mathbb{E}[\exp(is\eta_t^d)] &= \frac{\mathbb{E}[\exp(isx_t) | d_{t-1} = d] \cdot \mathbb{E}[\exp(is\gamma^d \varepsilon_{t-1})]}{\mathbb{E}[\exp(is(\alpha^d + \beta^d w_{t-1} + \gamma^d x_{t-1})) | d_{t-1} = d] \mathbb{E}[\exp(is\varepsilon_t)]} \\ &= \frac{\mathbb{E}[\exp(isx_t) | d_{t-1} = d] \cdot \exp\left[\int_0^s \frac{\mathbb{E}[i(x_{t+1} - \alpha^d - \beta^d w_t) \exp(is_1 x_t) | d_t = d]}{\gamma^d \mathbb{E}[\exp(is_1 x_t) | d_t = d]} ds_1\right]}{\mathbb{E}[\exp(is(\alpha^d + \beta^d w_{t-1} + \gamma^d x_{t-1})) | d_{t-1} = d] \cdot \mathbb{E}[\exp(isx_t) | d_t = d]} \times \\ &\quad \frac{\mathbb{E}[\exp(is\gamma^d x_{t-1}) | d_{t-1} = d]}{\exp\left[\int_0^{s\gamma^d} \frac{\mathbb{E}[i(x_t - \alpha^d - \beta^d w_{t-1}) \exp(is_1 x_{t-1}) | d_{t-1} = d]}{\gamma^d \mathbb{E}[\exp(is_1 x_{t-1}) | d_{t-1} = d]} ds_1\right]}. \end{aligned}$$

The denominator on the right-hand side is non-zero, as the conditional and unconditional characteristic functions do not vanish on the real line under Assumption 10. Letting  $\mathcal{F}$  denote the operator defined by

$$(\mathcal{F}\phi)(\xi) = \frac{1}{2\pi} \int e^{-is\xi} \phi(s) ds \quad \text{for all } \phi \in L^1(\mathbb{R}) \text{ and } \xi \in \mathbb{R},$$

we identify  $f_{\eta_t^d}$  by

$$f_{\eta_t^d}(\eta) = \left(\mathcal{F}\phi_{\eta_t^d}\right)(\eta) \quad \text{for all } \eta,$$

where the characteristic function  $\phi_{\eta_t^d}$  is given by

$$\phi_{\eta_t^d}(s) = \frac{\mathbb{E}[\exp(isx_t) | d_{t-1} = d] \cdot \exp\left[\int_0^s \frac{\mathbb{E}[i(x_{t+1} - \alpha^d - \beta^d w_t) \exp(is_1 x_t) | d_t = d]}{\gamma^d \mathbb{E}[\exp(is_1 x_t) | d_t = d]} ds_1\right]}{\mathbb{E}[\exp(is(\alpha^d + \beta^d w_{t-1} + \gamma^d x_{t-1})) | d_{t-1} = d] \cdot \mathbb{E}[\exp(isx_t) | d_t = d]} \times \frac{\mathbb{E}[\exp(is\gamma^d x_{t-1}) | d_{t-1} = d]}{\exp\left[\int_0^{s\gamma^d} \frac{\mathbb{E}[i(x_t - \alpha^d - \beta^d w_{t-1}) \exp(is_1 x_{t-1}) | d_{t-1} = d]}{\gamma^d \mathbb{E}[\exp(is_1 x_{t-1}) | d_{t-1} = d]} ds_1\right]}.$$

We can use this identified density in turn to identify the transition rule  $f_3(x_t^* | d_{t-1}, w_{t-1}, x_{t-1}^*)$  with

$$f_3(x_t^* | d_{t-1}, w_{t-1}, x_{t-1}^*) = \sum_d \mathbb{1}\{d_{t-1} = d\} f_{\eta_t^d}(x_t^* - \alpha^d - \beta^d w_{t-1} - \gamma^d x_{t-1}^*).$$

In summary, we obtain the closed-form expression

$$\begin{aligned} f_3(x_t^* | d_{t-1}, w_{t-1}, x_{t-1}^*) &= \sum_d \mathbb{1}\{d_{t-1} = d\} (\mathcal{F}\phi_{\eta_t^d})(x_t^* - \alpha^d - \beta^d w_{t-1} - \gamma^d x_{t-1}^*) \\ &= \sum_d \frac{\mathbb{1}\{d_{t-1} = d\}}{2\pi} \int \exp(-is(x_t^* - \alpha^d - \beta^d w_{t-1} - \gamma^d x_{t-1}^*)) \times \\ &\quad \frac{\mathbb{E}[\exp(isx_t) | d_{t-1} = d] \cdot \exp\left[\int_0^s \frac{\mathbb{E}[i(x_{t+1} - \alpha^{d'} - \beta^{d'} w_t) \exp(is_1 x_t) | d_t = d']}{\gamma^{d'} \mathbb{E}[\exp(is_1 x_t) | d_t = d']]} ds_1\right]}{\mathbb{E}[\exp(is(\alpha^d + \beta^d w_{t-1} + \gamma^d x_{t-1})) | d_{t-1} = d] \cdot \mathbb{E}[\exp(isx_t) | d_t = d]} \times \\ &\quad \frac{\mathbb{E}[\exp(is\gamma^d x_{t-1}) | d_{t-1} = d]}{\exp\left[\int_0^{s\gamma^d} \frac{\mathbb{E}[i(x_t - \alpha^{d'} - \beta^{d'} w_{t-1}) \exp(is_1 x_{t-1}) | d_{t-1} = d']}{\gamma^{d'} \mathbb{E}[\exp(is_1 x_{t-1}) | d_{t-1} = d']]} ds_1\right]} ds. \end{aligned}$$

using any  $d'$ . This completes Step 1.

**Step 2: Closed-form identification of the proxy model  $f_4(x_t | x_t^*)$ :** Given (B.1), we can write the density of  $\varepsilon_t$  by

$$f_{\varepsilon_t}(\varepsilon) = (\mathcal{F}\phi_{\varepsilon_t})(\varepsilon) \quad \text{for all } \varepsilon,$$

where the characteristic function  $\phi_{\varepsilon_t}$  is defined by (B.1) as

$$\phi_{\varepsilon_t}(s) = \frac{\mathbb{E}[\exp(isx_t) | d_t = d]}{\exp\left[\int_0^s \frac{\mathbb{E}[i(x_{t+1} - \alpha^d - \beta^d w_t) \exp(is' x_t) | d_t = d]}{\gamma^d \mathbb{E}[\exp(is' x_t) | d_t = d]} ds'\right]}.$$

Provided this identified density of  $\varepsilon_t$ , we nonparametrically identify the proxy model

$$f_4(x_t | x_t^*) = f_{\varepsilon_t}(x_t - x_t^*)$$

In summary, we obtain the closed-form expression

$$\begin{aligned} f_4(x_t | x_t^*) &= (\mathcal{F}\phi_{\varepsilon_t})(x_t - x_t^*) \\ &= \frac{1}{2\pi} \int \frac{\exp(-is(x_t - x_t^*)) \cdot \mathbb{E}[\exp(isx_t) | d_t = d]}{\exp\left[\int_0^s \frac{\mathbb{E}[i(x_{t+1} - \alpha^d - \beta^d w_t) \exp(is_1 x_t) | d_t = d]}{\gamma^d \mathbb{E}[\exp(is_1 x_t) | d_t = d]} ds_1\right]} ds \end{aligned}$$

using any  $d$ . This completes Step 2.

**Step 3: Closed-form identification of the transition rule  $f_2(w_t | d_{t-1}, w_{t-1}, x_{t-1}^*)$ :** Consider the joint density expressed by the convolution integral

$$f(x_{t-1}, w_t | d_{t-1}, w_{t-1}) = \int f_{\varepsilon_{t-1}}(x_{t-1} - x_{t-1}^*) f(x_{t-1}^*, w_t | d_{t-1}, w_{t-1}) dx_{t-1}^*$$

We can thus obtain a closed-form expression of  $f(x_{t-1}^*, w_t | d_{t-1}, w_{t-1})$  by the deconvolution.

To see this, observe

$$\begin{aligned} \mathbb{E}[\exp(is_1 x_{t-1} + is_2 w_t) | d_{t-1}, w_{t-1}] &= \mathbb{E}[\exp(is_1 x_{t-1}^* + is_1 \varepsilon_{t-1} + is_2 w_t) | d_{t-1}, w_{t-1}] \\ &= \mathbb{E}[\exp(is_1 x_{t-1}^* + is_2 w_t) | d_{t-1}, w_{t-1}] \mathbb{E}[\exp(is_1 \varepsilon_{t-1})] \end{aligned}$$

by the independence assumption for  $\varepsilon_t$ , and so

$$\begin{aligned} \mathbb{E}[\exp(is_1 x_{t-1}^* + is_2 w_t) | d_{t-1}, w_{t-1}] &= \frac{\mathbb{E}[\exp(is_1 x_{t-1} + is_2 w_t) | d_{t-1}, w_{t-1}]}{\mathbb{E}[\exp(is_1 \varepsilon_{t-1})]} \\ &= \frac{\mathbb{E}[\exp(is_1 x_{t-1} + is_2 w_t) | d_{t-1}, w_{t-1}] \cdot \exp\left[\int_0^{s_1} \frac{\mathbb{E}[i(x_t - \alpha^d - \beta^d w_{t-1}) \exp(is'_1 x_{t-1}) | d_{t-1} = d]}{\gamma^d \mathbb{E}[\exp(is'_1 x_{t-1}) | d_{t-1} = d]} ds'_1\right]}{\mathbb{E}[\exp(is_1 x_{t-1}) | d_{t-1} = d]} \end{aligned}$$

follows. Letting  $\mathcal{F}_2$  denote the operator defined by

$$(\mathcal{F}_2\phi)(\xi_1, \xi_2) = \frac{1}{4\pi^2} \int \int e^{-is_1 \xi_1 - is_2 \xi_2} \phi(s_1, s_2) ds_1 ds_2 \quad \text{for all } \phi \in L^1(\mathbb{R}^2) \text{ and } (\xi_1, \xi_2) \in \mathbb{R}^2,$$

we can express the conditional density as

$$f(x_{t-1}^*, w_t | d_{t-1}, w_{t-1}) = \left( \mathcal{F}_2 \phi_{x_{t-1}^*, w_t | d_{t-1}, w_{t-1}} \right) (w_t, x_{t-1}^*)$$

where the characteristic function is defined by

$$\begin{aligned} & \phi_{x_{t-1}^*, w_t | d_{t-1}, w_{t-1}}(s_1, s_2) \\ &= \frac{\mathbb{E}[\exp(is_1 x_{t-1} + is_2 w_t) | d_{t-1}, w_{t-1}] \cdot \exp\left[\int_0^{s_1} \frac{\mathbb{E}[i(x_t - \alpha^d - \beta^d w_{t-1}) \exp(is'_1 x_{t-1}) | d_{t-1} = d]}{\gamma^d \mathbb{E}[\exp(is'_1 x_{t-1}) | d_{t-1} = d]} ds'_1\right]}{\mathbb{E}[\exp(is_1 x_{t-1}) | d_{t-1} = d]} \end{aligned}$$

with any  $d$ . Using this conditional density, we can nonparametrically identify the transition rule  $f_2(w_t | d_{t-1}, w_{t-1}, x_{t-1}^*)$  with

$$f_2(w_t | d_{t-1}, w_{t-1}, x_{t-1}^*) = \frac{f(x_{t-1}^*, w_t | d_{t-1}, w_{t-1})}{\int f(x_{t-1}^*, w_t | d_{t-1}, w_{t-1}) dw_t}.$$

In summary, we obtain the closed-form expression

$$\begin{aligned} f_2(w_t | d_{t-1}, w_{t-1}, x_{t-1}^*) &= \frac{\left( \mathcal{F}_2 \phi_{x_{t-1}^*, w_t | d_{t-1}, w_{t-1}} \right) (x_{t-1}^*, w_t)}{\int \left( \mathcal{F}_2 \phi_{x_{t-1}^*, w_t | d_{t-1}, w_{t-1}} \right) (x_{t-1}^*, w_t) dw_t} \\ &= \sum_d \mathbb{1}\{d_{t-1} = d\} \int \int \exp(-is_1 w_t - is_2 x_{t-1}^*) \cdot \mathbb{E}[\exp(is_1 x_{t-1} + is_2 w_t) | d_{t-1} = d, w_{t-1}] \times \\ & \quad \frac{\exp\left[\int_0^{s_1} \frac{\mathbb{E}[i(x_t - \alpha^{d'} - \beta^{d'} w_{t-1}) \exp(is'_1 x_{t-1}) | d_{t-1} = d']}{\gamma^{d'} \mathbb{E}[\exp(is'_1 x_{t-1}) | d_{t-1} = d']} ds'_1\right]}{\mathbb{E}[\exp(is_1 x_{t-1}) | d_{t-1} = d']} ds_1 ds_2 \Big/ \\ & \quad \int \int \int \exp(-is_1 w_t - is_2 x_{t-1}^*) \cdot \mathbb{E}[\exp(is_1 x_{t-1} + is_2 w_t) | d_{t-1} = d, w_{t-1}] \times \\ & \quad \frac{\exp\left[\int_0^{s_1} \frac{\mathbb{E}[i(x_t - \alpha^{d'} - \beta^{d'} w_{t-1}) \exp(is'_1 x_{t-1}) | d_{t-1} = d']}{\gamma^{d'} \mathbb{E}[\exp(is'_1 x_{t-1}) | d_{t-1} = d']} ds'_1\right]}{\mathbb{E}[\exp(is_1 x_{t-1}) | d_{t-1} = d']} ds_1 ds_2 dw_t \end{aligned}$$

using any  $d'$ . This completes Step 3.

**Step 4: Closed-form identification of the CCP  $f_1(d_t | w_t, x_t^*)$ :** Note that we have

$$\begin{aligned} \mathbb{E}[\mathbb{1}\{d_t = d\} \exp(isx_t) | w_t] &= \mathbb{E}[\mathbb{1}\{d_t = d\} \exp(isx_t^* + is\varepsilon_t) | w_t] \\ &= \mathbb{E}[\mathbb{1}\{d_t = d\} \exp(isx_t^*) | w_t] \mathbb{E}[\exp(is\varepsilon_t)] \\ &= \mathbb{E}[\mathbb{E}[\mathbb{1}\{d_t = d\} | w_t, x_t^*] \exp(isx_t^*) | w_t] \mathbb{E}[\exp(is\varepsilon_t)] \end{aligned}$$

by the independence assumption for  $\varepsilon_t$  and the law of iterated expectations. Therefore

$$\begin{aligned} \frac{\mathbb{E}[\mathbb{1}\{d_t = d\} \exp(isx_t) | w_t]}{\mathbb{E}[\exp(is\varepsilon_t)]} &= \mathbb{E}[\mathbb{E}[\mathbb{1}\{d_t = d\} | w_t, x_t^*] \exp(isx_t^*) | w_t] \\ &= \int \exp(isx_t^*) \mathbb{E}[\mathbb{1}\{d_t = d\} | w_t, x_t^*] f(x_t^* | w_t) dx_t^* \end{aligned}$$

This is the Fourier inversion of  $\mathbb{E}[\mathbb{1}\{d_t = d\} | w_t, x_t^*] f(x_t^* | w_t)$ . On the other hand, the Fourier inversion of  $f(x_t^* | w_t)$  can be found as

$$\mathbb{E}[\exp(isx_t^*) | w_t] = \frac{\mathbb{E}[\exp(isx_t) | w_t]}{\mathbb{E}[\exp(is\varepsilon_t)]}.$$

Therefore, we find the closed-form expression for CCP  $f_1(d_t | w_t, x_t^*)$  as follows.

$$\Pr(d_t = d | w_t, x_t^*) = \mathbb{E}[\mathbb{1}\{d_t = d\} | w_t, x_t^*] = \frac{\mathbb{E}[\mathbb{1}\{d_t = d\} | w_t, x_t^*] f(x_t^* | w_t)}{f(x_t^* | w_t)} = \frac{(\mathcal{F}\phi_{(d)x_t^* | w_t})(x_t^*)}{(\mathcal{F}\phi_{x_t^* | w_t})(x_t^*)}$$

where the characteristic functions are defined by

$$\begin{aligned} \phi_{(d)x_t^* | w_t}(s) &= \frac{\mathbb{E}[\mathbb{1}\{d_t = d\} \exp(isx_t) | w_t]}{\mathbb{E}[\exp(is\varepsilon_t)]} \\ &= \frac{\mathbb{E}[\mathbb{1}\{d_t = d\} \exp(isx_t) | w_t] \cdot \exp\left[\int_0^s \frac{\mathbb{E}[i(x_{t+1} - \alpha^{d'} - \beta^{d'} w_t) \exp(is_1 x_t) | d_t = d']}{\gamma^{d'} \mathbb{E}[\exp(is_1 x_t) | d_t = d']} ds_1\right]}{\mathbb{E}[\exp(isx_t) | d_t = d']} \end{aligned}$$

and

$$\begin{aligned} \phi_{x_t^* | w_t}(s) &= \frac{\mathbb{E}[\exp(isx_t) | w_t]}{\mathbb{E}[\exp(is\varepsilon_t)]} \\ &= \frac{\mathbb{E}[\exp(isx_t) | w_t] \cdot \exp\left[\int_0^s \frac{\mathbb{E}[i(x_{t+1} - \alpha^{d'} - \beta^{d'} w_t) \exp(is_1 x_t) | d_t = d']}{\gamma^{d'} \mathbb{E}[\exp(is_1 x_t) | d_t = d']} ds_1\right]}{\mathbb{E}[\exp(isx_t) | d_t = d']} \end{aligned}$$

by (B.1) using any  $d'$ . In summary, we obtain the closed-form expression

$$\begin{aligned}
\Pr(d_t = d | w_t, x_t^*) &= \frac{(\mathcal{F}\phi_{(d)x_t^*|w_t})(x_t^*)}{(\mathcal{F}\phi_{x_t^*|w_t})(x_t^*)} \\
&= \int \exp(-isx_t^*) \cdot \mathbb{E}[\mathbb{1}\{d_t = d\} \exp(isx_t) | w_t] \times \\
&\quad \frac{\exp\left[\int_0^s \frac{\mathbb{E}[i(x_{t+1} - \alpha^{d'} - \beta^{d'} w_t) \exp(is_1 x_t) | d_t = d']}{\gamma^{d'} \mathbb{E}[\exp(is_1 x_t) | d_t = d']} ds_1\right]}{\mathbb{E}[\exp(isx_t) | d_t = d']} ds \Big/ \\
&\quad \int \exp(-isx_t^*) \cdot \mathbb{E}[\exp(isx_t) | w_t] \times \\
&\quad \frac{\exp\left[\int_0^s \frac{\mathbb{E}[i(x_{t+1} - \alpha^{d'} - \beta^{d'} w_t) \exp(is_1 x_t) | d_t = d']}{\gamma^{d'} \mathbb{E}[\exp(is_1 x_t) | d_t = d']} ds_1\right]}{\mathbb{E}[\exp(isx_t) | d_t = d']} ds
\end{aligned}$$

using any  $d'$ . This completes Step 4.  $\square$

## B.2 Proof of Theorem 3

Similarly to the baseline case, our closed-form identification includes four steps.

**Step 1: Closed-form identification of the transition rule  $f_3(x_t^* | d_{t-1}, w_{t-1}, x_{t-1}^*)$ :** First, we show the identification of the parameters and the distributions in transition of  $x_t^*$ . Since

$$\begin{aligned}
x_t &= \sum_d \mathbb{1}\{d_{t-1} = d\} [\delta^d x_t^* + \varepsilon_t^d] \\
&= \sum_d \mathbb{1}\{d_{t-1} = d\} [\alpha^d \delta^d + \beta^d \delta^d w_{t-1} + \gamma^d \delta^d x_{t-1}^* + \delta^d \eta_t^d + \varepsilon_t^d] \\
&= \sum_d \sum_{d'} \mathbb{1}\{d_{t-1} = d\} \mathbb{1}\{d_{t-2} = d'\} \left[ \alpha^d \delta^d + \beta^d \delta^d w_{t-1} + \gamma^d \frac{\delta^d}{\delta^{d'}} x_{t-1} + \delta^d \eta_t^d + \varepsilon_t^d - \gamma^d \frac{\delta^d}{\delta^{d'}} \varepsilon_{t-1}^{d'} \right]
\end{aligned}$$

we obtain the following equalities for each  $d$  and  $d'$ :

$$\begin{aligned}
\mathbb{E}[x_t \mid d_{t-1} = d, d_{t-2} = d'] &= \alpha^d \delta^d + \beta^d \delta^d \mathbb{E}[w_{t-1} \mid d_{t-1} = d, d_{t-2} = d'] \\
&\quad + \gamma^d \frac{\delta^d}{\delta^{d'}} \mathbb{E}[x_{t-1} \mid d_{t-1} = d, d_{t-2} = d'] \\
\mathbb{E}[x_t w_{t-1} \mid d_{t-1} = d, d_{t-2} = d'] &= \alpha^d \delta^d \mathbb{E}[w_{t-1} \mid d_{t-1} = d, d_{t-2} = d'] \\
&\quad + \beta^d \delta^d \mathbb{E}[w_{t-1}^2 \mid d_{t-1} = d, d_{t-2} = d'] \\
&\quad + \gamma^d \frac{\delta^d}{\delta^{d'}} \mathbb{E}[x_{t-1} w_{t-1} \mid d_{t-1} = d, d_{t-2} = d'] \\
\mathbb{E}[x_t w_t \mid d_{t-1} = d, d_{t-2} = d'] &= \alpha^d \delta^d \mathbb{E}[w_t \mid d_{t-1} = d, d_{t-2} = d'] \\
&\quad + \beta^d \delta^d \mathbb{E}[w_{t-1} w_t \mid d_{t-1} = d, d_{t-2} = d'] \\
&\quad + \gamma^d \frac{\delta^d}{\delta^{d'}} \mathbb{E}[x_{t-1} w_t \mid d_{t-1} = d, d_{t-2} = d']
\end{aligned}$$

by the independence and zero mean assumptions for  $\eta_t^d$  and  $\varepsilon_t^d$ . From these, we have the linear equation

$$\begin{bmatrix} \mathbb{E}[x_t \mid d_{t-1} = d, d_{t-2} = d'] \\ \mathbb{E}[x_t w_{t-1} \mid d_{t-1} = d, d_{t-2} = d'] \\ \mathbb{E}[x_t w_t \mid d_{t-1} = d, d_{t-2} = d'] \end{bmatrix} = \begin{bmatrix} 1 & \mathbb{E}[w_{t-1} \mid d_{t-1} = d, d_{t-2} = d'] & \mathbb{E}[x_{t-1} \mid d_{t-1} = d, d_{t-2} = d'] \\ \mathbb{E}[w_{t-1} \mid d_{t-1} = d, d_{t-2} = d'] & \mathbb{E}[w_{t-1}^2 \mid d_{t-1} = d, d_{t-2} = d'] & \mathbb{E}[x_{t-1} w_{t-1} \mid d_{t-1} = d, d_{t-2} = d'] \\ \mathbb{E}[w_t \mid d_{t-1} = d, d_{t-2} = d'] & \mathbb{E}[w_{t-1} w_t \mid d_{t-1} = d, d_{t-2} = d'] & \mathbb{E}[x_{t-1} w_t \mid d_{t-1} = d, d_{t-2} = d'] \end{bmatrix} \begin{bmatrix} \alpha^d \delta^d \\ \beta^d \delta^d \\ \gamma^d \frac{\delta^d}{\delta^{d'}} \end{bmatrix}$$

Provided that the matrix on the right-hand side is non-singular, we can identify the composite parameters  $\left( \alpha^d \delta^d, \beta^d \delta^d, \gamma^d \frac{\delta^d}{\delta^{d'}} \right)$  by

$$\begin{bmatrix} \alpha^d \delta^d \\ \beta^d \delta^d \\ \gamma^d \frac{\delta^d}{\delta^{d'}} \end{bmatrix} = \begin{bmatrix} 1 & \mathbb{E}[w_{t-1} \mid d_{t-1} = d, d_{t-2} = d'] & \mathbb{E}[x_{t-1} \mid d_{t-1} = d, d_{t-2} = d'] \\ \mathbb{E}[w_{t-1} \mid d_{t-1} = d, d_{t-2} = d'] & \mathbb{E}[w_{t-1}^2 \mid d_{t-1} = d, d_{t-2} = d'] & \mathbb{E}[x_{t-1} w_{t-1} \mid d_{t-1} = d, d_{t-2} = d'] \\ \mathbb{E}[w_t \mid d_{t-1} = d, d_{t-2} = d'] & \mathbb{E}[w_{t-1} w_t \mid d_{t-1} = d, d_{t-2} = d'] & \mathbb{E}[x_{t-1} w_t \mid d_{t-1} = d, d_{t-2} = d'] \end{bmatrix}^{-1} \times \begin{bmatrix} \mathbb{E}[x_t \mid d_{t-1} = d, d_{t-2} = d'] \\ \mathbb{E}[x_t w_{t-1} \mid d_{t-1} = d, d_{t-2} = d'] \\ \mathbb{E}[x_t w_t \mid d_{t-1} = d, d_{t-2} = d'] \end{bmatrix}.$$

Once the composite parameters  $\gamma^d \frac{\delta^d}{\delta^0}$  and  $\gamma^d = \gamma^d \frac{\delta^d}{\delta^d}$  are identified by the above formula, we can in turn identify

$$\delta^d = \frac{\gamma^d \frac{\delta^d}{\delta^0}}{\gamma^d \frac{\delta^d}{\delta^d}}$$

for each  $d$  by the normalization assumption  $\delta^0 = 1$ . It in turn can be used to identify  $(\alpha^d, \beta^d, \gamma^d)$  for each  $d$  from the identified composite parameters  $(\alpha^d \delta^d, \beta^d \delta^d, \gamma^d \frac{\delta^d}{\delta^0})$  by

$$(\alpha^d, \beta^d, \gamma^d) = \frac{1}{\delta^d} \left( \alpha^d \delta^d, \beta^d \delta^d, \gamma^d \frac{\delta^d}{\delta^0} \right).$$

Next, we show identification of the distributions of  $\varepsilon_t^d$  and  $\eta_t^d$  for each  $d$ . Observe that

$$\begin{aligned} & \text{E} [\exp (i s_1 x_{t-1} + i s_2 x_t) | d_{t-1} = d, d_{t-2} = d'] \\ &= \text{E} \left[ \exp \left( i s_1 \left( \delta^{d'} x_{t-1}^* + \varepsilon_{t-1}^{d'} \right) + i s_2 \left( \alpha^d \delta^d + \beta^d \delta^d w_{t-1} + \gamma^d \delta^d x_{t-1}^* + \delta^d \eta_t^d + \varepsilon_t^d \right) \right) | d_{t-1} = d, d_{t-2} = d' \right] \\ &= \text{E} \left[ \exp \left( i \left( s_1 \delta^{d'} x_{t-1}^* + s_2 \alpha^d \delta^d + s_2 \beta^d \delta^d w_{t-1} + s_2 \gamma^d \delta^d x_{t-1}^* \right) \right) | d_{t-1} = d, d_{t-2} = d' \right] \\ & \quad \times \text{E} \left[ \exp \left( i s_1 \varepsilon_{t-1}^{d'} \right) \right] \text{E} \left[ \exp \left( i s_2 \left( \delta^d \eta_t^d + \varepsilon_t^d \right) \right) \right] \end{aligned}$$

follows for each pair  $(d, d')$  from the independence assumptions for  $\eta_t^d$  and  $\varepsilon_t^d$  for each  $d$ . We may then use the Kotlarski's identity

$$\begin{aligned} & \left[ \frac{\partial}{\partial s_2} \ln \text{E} [\exp (i s_1 x_{t-1} + i s_2 x_t) | d_{t-1} = d, d_{t-2} = d'] \right]_{s_2=0} \\ &= \frac{\text{E} [i(\alpha^d \delta^d + \beta^d \delta^d w_{t-1} + \gamma^d \delta^d x_{t-1}^*) \exp (i s_1 \delta^{d'} x_{t-1}^*) | d_{t-1} = d, d_{t-2} = d']}{\text{E} [\exp (i s_1 \delta^{d'} x_{t-1}^*) | d_{t-1} = d, d_{t-2} = d']} \\ &= i \alpha^d \delta^d + \beta^d \delta^d \frac{\text{E}[i w_{t-1} \exp(i s_1 \delta^{d'} x_{t-1}^*) | d_{t-1} = d, d_{t-2} = d']}{\text{E}[\exp(i s_1 \delta^{d'} x_{t-1}^*) | d_{t-1} = d, d_{t-2} = d']} \\ & \quad + \gamma^d \frac{\delta^d}{\delta^{d'}} \frac{\partial}{\partial s_1} \ln \text{E} \left[ \exp \left( i s_1 \delta^{d'} x_{t-1}^* \right) | d_{t-1} = d, d_{t-2} = d' \right] \\ &= i \alpha^d \delta^d + \beta^d \delta^d \frac{\text{E}[i w_{t-1} \exp(i s_1 x_{t-1}) | d_{t-1} = d, d_{t-2} = d']}{\text{E}[\exp(i s_1 x_{t-1}) | d_{t-1} = d, d_{t-2} = d']} \\ & \quad + \gamma^d \frac{\delta^d}{\delta^{d'}} \frac{\partial}{\partial s_1} \ln \text{E} \left[ \exp \left( i s_1 \delta^{d'} x_{t-1}^* \right) | d_{t-1} = d, d_{t-2} = d' \right] \end{aligned}$$

Therefore,

$$\begin{aligned}
& \mathbb{E} \left[ \exp \left( is\delta^{d'} x_{t-1}^* \right) \mid d_{t-1} = d, d_{t-2} = d' \right] \\
&= \exp \left[ \int_0^s \left[ \frac{\delta^{d'}}{\gamma^d \delta^d} \frac{\partial}{\partial s_2} \ln \mathbb{E} \left[ \exp (is_1 x_{t-1} + is_2 x_t) \mid d_{t-1} = d, d_{t-2} = d' \right] \right]_{s_2=0} ds_1 \right. \\
&\quad \left. - \int_0^s \frac{i\alpha^d \delta^{d'}}{\gamma^d} ds_1 - \int_0^s \frac{\beta^d \delta^{d'}}{\gamma^d} \frac{\mathbb{E}[iw_{t-1} \exp(is_1 x_{t-1}) \mid d_{t-1} = d, d_{t-2} = d']}{\mathbb{E}[\exp(is_1 x_{t-1}) \mid d_{t-1} = d, d_{t-2} = d']} ds_1 \right] \\
&= \exp \left[ \int_0^s \frac{\mathbb{E} \left[ i \left( \frac{\delta^{d'}}{\delta^d} x_t - \alpha^d \delta^{d'} - \beta^d \delta^{d'} w_{t-1} \right) \exp(is_1 x_{t-1}) \mid d_{t-1} = d, d_{t-2} = d' \right]}{\gamma^d \mathbb{E} [\exp(is_1 x_{t-1}) \mid d_{t-1} = d, d_{t-2} = d']} ds_1 \right].
\end{aligned}$$

From the proxy model and the independence assumption for  $\varepsilon_t$ ,

$$\mathbb{E} [\exp(isx_{t-1}) \mid d_{t-1} = d, d_{t-2} = d'] = \mathbb{E} \left[ \exp \left( is\delta^{d'} x_{t-1}^* \right) \mid d_{t-1} = d, d_{t-2} = d' \right] \mathbb{E} \left[ \exp \left( is\varepsilon_{t-1}^{d'} \right) \right].$$

We then obtain the following result using any  $d$ .

$$\begin{aligned}
\mathbb{E} \left[ \exp \left( is\varepsilon_{t-1}^{d'} \right) \right] &= \frac{\mathbb{E} [\exp(isx_{t-1}) \mid d_{t-1} = d, d_{t-2} = d']}{\mathbb{E} \left[ \exp \left( is\delta^{d'} x_{t-1}^* \right) \mid d_{t-1} = d, d_{t-2} = d' \right]} \\
&= \frac{\mathbb{E} [\exp(isx_{t-1}) \mid d_{t-1} = d, d_{t-2} = d']}{\exp \left[ \int_0^s \frac{\mathbb{E} \left[ i \left( \frac{\delta^{d'}}{\delta^d} x_t - \alpha^d \delta^{d'} - \beta^d \delta^{d'} w_{t-1} \right) \exp(is_1 x_{t-1}) \mid d_{t-1} = d, d_{t-2} = d' \right]}{\gamma^d \mathbb{E} [\exp(is_1 x_{t-1}) \mid d_{t-1} = d, d_{t-2} = d']} ds_1 \right]}.
\end{aligned}$$

This argument holds for all  $t$  so that we can identify the characteristic function of  $\varepsilon_t^d$  for each  $d$  with

$$\mathbb{E} [\exp(is\varepsilon_t^d)] = \frac{\mathbb{E} [\exp(isx_t) \mid d_t = d', d_{t-1} = d]}{\exp \left[ \int_0^s \frac{\mathbb{E} \left[ i \left( \frac{\delta^{d'}}{\delta^d} x_{t+1} - \alpha^{d'} \delta^d - \beta^{d'} \delta^d w_t \right) \exp(is_1 x_t) \mid d_t = d', d_{t-1} = d \right]}{\gamma^{d'} \mathbb{E} [\exp(is_1 x_t) \mid d_t = d', d_{t-1} = d]} ds_1 \right]}. \quad (\text{B.2})$$

using any  $d'$ .

In order to identify  $f_{\eta_t^d}$  for each  $d$ , consider

$$\begin{aligned}
& \mathbb{E} [\exp(isx_t) \mid d_{t-1} = d, d_{t-2} = d'] \mathbb{E} \left[ \exp \left( is\gamma^d \frac{\delta^d}{\delta^{d'}} \varepsilon_{t-1}^{d'} \right) \right] \\
&= \mathbb{E} \left[ \exp \left( is \left( \alpha^d \delta^d + \beta^d \delta^d w_{t-1} + \gamma^d \frac{\delta^d}{\delta^{d'}} x_{t-1} \right) \mid d_{t-1} = d, d_{t-2} = d' \right) \right] \\
&\quad \times \mathbb{E} [\exp(is\delta^d \eta_t^d)] \mathbb{E} [\exp(is\varepsilon_t^d)]
\end{aligned}$$

by the independence assumptions for  $\eta_t^d$  and  $\varepsilon_t^d$ . Therefore,

$$\begin{aligned} \mathbb{E} [\exp (i s \delta^d \eta_t^d)] &= \frac{\mathbb{E} [\exp (i s x_t) \mid d_{t-1} = d, d_{t-2} = d']}{\mathbb{E} \left[ \exp \left( i s \left( \alpha^d \delta^d + \beta^d \delta^d w_{t-1} + \gamma^d \frac{\delta^d}{\delta^{d'}} x_{t-1} \right) \mid d_{t-1} = d, d_{t-2} = d' \right) \right]} \\ &\quad \times \frac{\mathbb{E} \left[ \exp \left( i s \gamma^d \frac{\delta^d}{\delta^{d'}} \varepsilon_{t-1}^d \right) \right]}{\mathbb{E} [\exp (i s \varepsilon_t^d)]} \end{aligned}$$

and the characteristic function of  $\eta_t^d$  can be expressed by

$$\begin{aligned} \mathbb{E} [\exp (i s \eta_t^d)] &= \frac{\mathbb{E} [\exp (i s \frac{1}{\delta^d} x_t) \mid d_{t-1} = d, d_{t-2} = d']}{\mathbb{E} \left[ \exp \left( i s \left( \alpha^d + \beta^d w_{t-1} + \gamma^d \frac{1}{\delta^{d'}} x_{t-1} \right) \mid d_{t-1} = d, d_{t-2} = d' \right) \right]} \\ &\quad \times \frac{1}{\mathbb{E} [\exp (i s \frac{1}{\delta^d} \varepsilon_t^d)] \mathbb{E} \left[ \exp \left( -i s \gamma^d \frac{1}{\delta^{d'}} \varepsilon_{t-1}^d \right) \right]} \\ &= \frac{\mathbb{E} [\exp (i s \frac{1}{\delta^d} x_t) \mid d_{t-1} = d, d_{t-2} = d']}{\mathbb{E} \left[ \exp \left( i s \left( \alpha^d + \beta^d w_{t-1} + \gamma^d \frac{1}{\delta^{d'}} x_{t-1} \right) \mid d_{t-1} = d, d_{t-2} = d' \right) \right]} \\ &\quad \times \frac{\exp \left[ \int_0^{s/\delta^d} \frac{\mathbb{E} \left[ i \left( \frac{\delta^d}{\delta^{d'}} x_{t+1} - \alpha^d \delta^d - \beta^d \delta^d w_t \right) \exp (i s_1 x_t) \mid d_t = d', d_{t-1} = d \right]}{\gamma^{d'} \mathbb{E} [\exp (i s_1 x_t) \mid d_t = d', d_{t-1} = d]} ds_1 \right]}{\mathbb{E} [\exp (i s \frac{1}{\delta^d} x_t) \mid d_t = d', d_{t-1} = d]} \\ &\quad \times \frac{\mathbb{E} \left[ \exp \left( i s \gamma^d \frac{1}{\delta^{d'}} x_{t-1} \right) \mid d_{t-1} = d, d_{t-2} = d' \right]}{\exp \left[ \int_0^{s \gamma^d / \delta^{d'}} \frac{\mathbb{E} \left[ i \left( \frac{\delta^{d'}}{\delta^d} x_t - \alpha^d \delta^{d'} - \beta^d \delta^{d'} w_{t-1} \right) \exp (i s_1 x_{t-1}) \mid d_{t-1} = d, d_{t-2} = d' \right]}{\gamma^d \mathbb{E} [\exp (i s_1 x_t) \mid d_{t-1} = d, d_{t-2} = d']} ds_1 \right]} \end{aligned}$$

by the formula (B.2). We can then identify  $f_{\eta_t^d}$  by

$$f_{\eta_t^d}(\eta) = \left( \mathcal{F} \phi_{\eta_t^d} \right) (\eta) \quad \text{for all } \eta,$$

where the characteristic function  $\phi_{\eta_t^d}$  is given by

$$\begin{aligned} \phi_{\eta_t^d}(s) &= \frac{\mathbb{E} [\exp (i s \frac{1}{\delta^d} x_t) \mid d_{t-1} = d, d_{t-2} = d']}{\mathbb{E} \left[ \exp \left( i s \left( \alpha^d + \beta^d w_{t-1} + \gamma^d \frac{1}{\delta^{d'}} x_{t-1} \right) \mid d_{t-1} = d, d_{t-2} = d' \right) \right]} \\ &\quad \times \frac{\exp \left[ \int_0^{s/\delta^d} \frac{\mathbb{E} \left[ i \left( \frac{\delta^d}{\delta^{d'}} x_{t+1} - \alpha^d \delta^d - \beta^d \delta^d w_t \right) \exp (i s_1 x_t) \mid d_t = d', d_{t-1} = d \right]}{\gamma^{d'} \mathbb{E} [\exp (i s_1 x_t) \mid d_t = d', d_{t-1} = d]} ds_1 \right]}{\mathbb{E} [\exp (i s \frac{1}{\delta^d} x_t) \mid d_t = d', d_{t-1} = d]} \\ &\quad \times \frac{\mathbb{E} \left[ \exp \left( i s \gamma^d \frac{1}{\delta^{d'}} x_{t-1} \right) \mid d_{t-1} = d, d_{t-2} = d' \right]}{\exp \left[ \int_0^{s \gamma^d / \delta^{d'}} \frac{\mathbb{E} \left[ i \left( \frac{\delta^{d'}}{\delta^d} x_t - \alpha^d \delta^{d'} - \beta^d \delta^{d'} w_{t-1} \right) \exp (i s_1 x_{t-1}) \mid d_{t-1} = d, d_{t-2} = d' \right]}{\gamma^d \mathbb{E} [\exp (i s_1 x_t) \mid d_{t-1} = d, d_{t-2} = d']} ds_1 \right]} \end{aligned}$$

We can use this identified density in turn to identify the transition rule  $f_3(x_t^* | d_{t-1}, w_{t-1}, x_{t-1}^*)$  with

$$f_3(x_t^* | d_{t-1}, x_{t-1}, x_{t-1}^*) = \sum_d \mathbb{1}\{d_{t-1} = d\} f_{\eta_t^d}(x_t^* - \alpha^d - \beta^d w_{t-1} - \gamma^d x_{t-1}^*).$$

In summary, we obtain the closed-form expression

$$\begin{aligned} f_3(x_t^* | d_{t-1}, x_{t-1}, x_{t-1}^*) &= \sum_d \mathbb{1}\{d_{t-1} = d\} \left( \mathcal{F} \phi_{\eta_t^d} \right) (x_t^* - \alpha^d - \beta^d w_{t-1} - \gamma^d x_{t-1}^*) \\ &= \sum_d \frac{\mathbb{1}\{d_{t-1} = d\}}{2\pi} \int \exp(-is(x_t^* - \alpha^d - \beta^d w_{t-1} - \gamma^d x_{t-1}^*)) \times \\ &\quad \frac{\mathbb{E} \left[ \exp \left( is \frac{1}{\delta^d} x_t \right) | d_{t-1} = d, d_{t-2} = d' \right]}{\mathbb{E} \left[ \exp \left( is \left( \alpha^d + \beta^d w_{t-1} + \gamma^d \frac{1}{\delta^d} x_{t-1} \right) \right) | d_{t-1} = d, d_{t-2} = d' \right]} \times \\ &\quad \exp \left[ \int_0^{s/\delta^d} \frac{\mathbb{E} \left[ i \left( \frac{\delta^d}{\delta^{d'}} x_{t+1} - \alpha^{d'} \delta^d - \beta^{d'} \delta^d w_t \right) \exp(is_1 x_t) | d_t = d', d_{t-1} = d \right]}{\gamma^{d'} \mathbb{E}[\exp(is_1 x_t) | d_t = d', d_{t-1} = d]} ds_1 \right] \times \\ &\quad \frac{\mathbb{E} \left[ \exp \left( is \frac{1}{\delta^d} x_t \right) | d_t = d', d_{t-1} = d \right]}{\mathbb{E} \left[ \exp \left( is \gamma^d \frac{1}{\delta^{d'}} x_{t-1} \right) | d_{t-1} = d, d_{t-2} = d' \right]} \times \\ &\quad \exp \left[ \int_0^{s\gamma^d/\delta^{d'}} \frac{\mathbb{E} \left[ i \left( \frac{\delta^{d'}}{\delta^d} x_t - \alpha^d \delta^{d'} - \beta^d \delta^{d'} w_{t-1} \right) \exp(is_1 x_{t-1}) | d_{t-1} = d, d_{t-2} = d' \right]}{\gamma^d \mathbb{E}[\exp(is_1 x_t) | d_{t-1} = d, d_{t-2} = d']} ds_1 \right] \end{aligned}$$

using any  $d'$ . This completes Step 1.

**Step 2: Closed-form identification of the proxy model  $f_4(x_t | d_{t-1}, x_t^*)$ :** Given (B.2),

we can write the density of  $\varepsilon_t^d$  by

$$f_{\varepsilon_t^d}(\varepsilon) = \left( \mathcal{F} \phi_{\varepsilon_t^d} \right) (\varepsilon) \quad \text{for all } \varepsilon,$$

where the characteristic function  $\phi_{\varepsilon_t^d}$  is defined by (B.2) as

$$\phi_{\varepsilon_t^d}(s) = \frac{\mathbb{E} \left[ \exp(isx_t) | d_t = d', d_{t-1} = d \right]}{\exp \left[ \int_0^s \frac{\mathbb{E} \left[ i \left( \frac{\delta^d}{\delta^{d'}} x_{t+1} - \alpha^{d'} \delta^d - \beta^{d'} \delta^d w_t \right) \exp(is_1 x_t) | d_t = d', d_{t-1} = d \right]}{\gamma^{d'} \mathbb{E}[\exp(is_1 x_t) | d_t = d', d_{t-1} = d]} ds_1 \right]}.$$

Provided this identified density of  $\varepsilon_t^d$ , we nonparametrically identify the proxy model

$$f_4(x_t | d_{t-1} = d, x_t^*) = f_{\varepsilon_t^d | d_{t-1} = d}(x_t - \delta^d x_t^*) = f_{\varepsilon_t^d}(x_t - \delta^d x_t^*)$$

by the independence assumption for  $\varepsilon_t^d$ . In summary, we obtain the closed-form expression

$$\begin{aligned} f_4(x_t | d_{t-1}, x_t^*) &= \sum_d \mathbb{1}\{d_{t-1} = d\} \left( \mathcal{F}\phi_{\varepsilon_t^d} \right) (x_t - \delta^d x_t^*) \\ &= \sum_d \frac{\mathbb{1}\{d_{t-1} = d\}}{2\pi} \int \frac{\exp(-is(x_t - \delta^d x_t^*)) \cdot \mathbb{E}[\exp(isx_t) | d_t = d', d_{t-1} = d]}{\exp\left[\int_0^s \frac{\mathbb{E}\left[i\left(\frac{\delta^d}{\delta^{d'}}x_{t+1} - \alpha^{d'}\delta^d - \beta^{d'}\delta^d w_t\right)\exp(is_1x_t) | d_t = d', d_{t-1} = d\right]}{\gamma^{d'} \mathbb{E}[\exp(is_1x_t) | d_t = d', d_{t-1} = d]} ds_1\right]} ds \end{aligned}$$

using any  $d'$ . This completes Step 2.

**Step 3: Closed-form identification of the transition rule  $f_2(w_t | d_{t-1}, w_{t-1}, x_{t-1}^*)$ :** Consider the joint density expressed by the convolution integral

$$f(x_{t-1}, w_t | d_{t-1}, w_{t-1}, d_{t-2} = d) = \int f_{\varepsilon_{t-1}^d}(x_{t-1} - \delta^d x_{t-1}^*) f(x_{t-1}^*, w_t | d_{t-1}, w_{t-1}, d_{t-2} = d) dx_{t-1}^*$$

We can thus obtain a closed-form expression of  $f(x_{t-1}^*, w_t | d_{t-1}, w_{t-1}, d_{t-2} = d)$  by the deconvolution. To see this, observe

$$\begin{aligned} &\mathbb{E}[\exp(is_1x_{t-1} + is_2w_t) | d_{t-1}, w_{t-1}, d_{t-2} = d] \\ &= \mathbb{E}[\exp(is_1\delta^d x_{t-1}^* + is_1\varepsilon_{t-1}^d + is_2w_t) | d_{t-1}, w_{t-1}, d_{t-2} = d] \\ &= \mathbb{E}[\exp(is_1\delta^d x_{t-1}^* + is_2w_t) | d_{t-1}, w_{t-1}, d_{t-2} = d] \mathbb{E}[\exp(is_1\varepsilon_{t-1}^d)] \end{aligned}$$

by the independence assumption for  $\varepsilon_t^d$ , and so

$$\begin{aligned} \mathbb{E}[\exp(is_1\delta^d x_{t-1}^* + is_2w_t) | d_{t-1}, w_{t-1}, d_{t-2} = d] &= \frac{\mathbb{E}[\exp(is_1x_{t-1} + is_2w_t) | d_{t-1}, w_{t-1}, d_{t-2} = d]}{\mathbb{E}[\exp(is_1\varepsilon_{t-1}^d)]} \\ &= \mathbb{E}[\exp(is_1x_{t-1} + is_2w_t) | d_{t-1}, w_{t-1}, d_{t-2} = d] \\ &\quad \times \frac{\exp\left[\int_0^{s_1} \frac{\mathbb{E}\left[i\left(\frac{\delta^d}{\delta^{d'}}x_t - \alpha^{d'}\delta^d - \beta^{d'}\delta^d w_{t-1}\right)\exp(is'_1x_{t-1}) | d_{t-1} = d', d_{t-2} = d\right]}{\gamma^{d'} \mathbb{E}[\exp(is'_1x_{t-1}) | d_{t-1} = d', d_{t-2} = d]} ds'_1\right]}{\mathbb{E}[\exp(is_1x_{t-1}) | d_{t-1} = d', d_{t-2} = d]} \end{aligned}$$

follows with any choice of  $d'$ . Rescaling  $s_1$  yields

$$\begin{aligned} & \text{E} \left[ \exp \left( i s_1 x_{t-1}^* + i s_2 w_t \right) \mid d_{t-1}, w_{t-1}, d_{t-2} = d \right] \\ = & \text{E} \left[ \exp \left( i s_1 \frac{1}{\delta^d} x_{t-1} + i s_2 w_t \right) \mid d_{t-1}, w_{t-1}, d_{t-2} = d \right] \times \\ & \frac{\exp \left[ \int_0^{s_1/\delta^d} \frac{\text{E} \left[ i \left( \frac{\delta^d}{\delta^{d'}} x_t - \alpha^{d'} \delta^d - \beta^{d'} \delta^d w_{t-1} \right) \exp(i s'_1 x_{t-1}) \mid d_{t-1}=d', d_{t-2}=d \right]}{\gamma^{d'} \text{E} \left[ \exp(i s'_1 x_{t-1}) \mid d_{t-1}=d', d_{t-2}=d \right]} ds'_1 \right]}{\text{E} \left[ \exp \left( i s_1 \frac{1}{\delta^d} x_{t-1} \right) \mid d_{t-1} = d', d_{t-2} = d \right]}. \end{aligned}$$

We can then express the conditional density as

$$f \left( x_{t-1}^*, w_t \mid d_{t-1}, w_{t-1}, d_{t-2} = d \right) = \left( \mathcal{F}_2 \phi_{x_{t-1}^*, w_t \mid d_{t-1}, w_{t-1}, d_{t-2}=d} \right) (w_t, x_{t-1}^*)$$

where the characteristic function is defined by

$$\begin{aligned} \phi_{w_t, x_{t-1}^* \mid d_{t-1}, w_{t-1}, d_{t-2}=d}(s_1, s_2) &= \text{E} \left[ \exp \left( i s_1 \frac{1}{\delta^d} x_{t-1} + i s_2 w_t \right) \mid d_{t-1}, w_{t-1}, d_{t-2} = d \right] \times \\ & \frac{\exp \left[ \int_0^{s_1/\delta^d} \frac{\text{E} \left[ i \left( \frac{\delta^d}{\delta^{d'}} x_t - \alpha^{d'} \delta^d - \beta^{d'} \delta^d w_{t-1} \right) \exp(i s'_1 x_{t-1}) \mid d_{t-1}=d', d_{t-2}=d \right]}{\gamma^{d'} \text{E} \left[ \exp(i s'_1 x_{t-1}) \mid d_{t-1}=d', d_{t-2}=d \right]} ds'_1 \right]}{\text{E} \left[ \exp \left( i s_1 \frac{1}{\delta^d} x_{t-1} \right) \mid d_{t-1} = d', d_{t-2} = d \right]}. \end{aligned}$$

Using this conditional density, we nonparametrically identify the transition rule

$$f_2 \left( w_t \mid d_{t-1}, w_{t-1}, x_{t-1}^* \right) = \frac{\sum_d f \left( x_{t-1}^*, w_t \mid d_{t-1}, w_{t-1}, d_{t-2} = d \right) \Pr(d_{t-2} = d \mid d_{t-1}, w_{t-1})}{\int \sum_d f \left( x_{t-1}^*, w_t \mid d_{t-1}, w_{t-1}, d_{t-2} = d \right) \Pr(d_{t-2} = d \mid d_{t-1}, w_{t-1}) dw_t}.$$

In summary, we obtain the closed-form expression

$$\begin{aligned}
f_2(w_t|d_{t-1}, w_{t-1}, x_{t-1}^*) &= \sum_d \mathbb{1}\{d_{t-1} = d\} \times \\
&\frac{\sum_{d'} \left( \mathcal{F}_2 \phi_{x_{t-1}^*, w_t | d_{t-1}=d, w_{t-1}, d_{t-2}=d'} \right) (w_t, x_{t-1}^*) \cdot \Pr(d_{t-2} = d' | d_{t-1} = d, w_{t-1})}{\int \sum_{d'} \left( \mathcal{F}_2 \phi_{x_{t-1}^*, w_t | d_{t-1}=d, w_{t-1}, d_{t-2}=d'} \right) (w_t, x_{t-1}^*) \cdot \Pr(d_{t-2} = d' | d_{t-1} = d, w_{t-1}) dw_t} \\
&= \sum_d \mathbb{1}\{d_{t-1} = d\} \left\{ \sum_{d'} \Pr(d_{t-2} = d' | d_{t-1} = d, w_{t-1}) \int \int \exp(-is_1 w_t - is_2 x_{t-1}^*) \times \right. \\
&\frac{\mathbb{E} \left[ \exp \left( is_1 \frac{1}{\delta^{d'}} x_{t-1} + is_2 w_t \right) | d_{t-1} = d, w_{t-1}, d_{t-2} = d' \right]}{\mathbb{E} \left[ \exp \left( is_1 \frac{1}{\delta^{d'}} x_{t-1} \right) | d_{t-1} = d'', d_{t-2} = d' \right]} \times \\
&\left. \exp \left[ \int_0^{s_1/\delta^{d'}} \frac{\mathbb{E} \left[ i \left( \frac{\delta^{d'}}{\delta^{d''}} x_t - \alpha^{d''} \delta^{d'} - \beta^{d''} \delta^{d'} w_{t-1} \right) \exp(is'_1 x_{t-1}) | d_{t-1} = d'', d_{t-2} = d' \right]}{\gamma^{d''} \mathbb{E} \left[ \exp(is'_1 x_{t-1}) | d_{t-1} = d'', d_{t-2} = d' \right]} ds'_1 \right] ds_1 ds_2 \right\} / \\
&\left\{ \sum_{d'} \int \Pr(d_{t-2} = d' | d_{t-1} = d, w_{t-1}) \int \int \exp(-is_1 w_t - is_2 x_{t-1}^*) \times \right. \\
&\frac{\mathbb{E} \left[ \exp \left( is_1 \frac{1}{\delta^{d'}} x_{t-1} + is_2 w_t \right) | d_{t-1} = d, w_{t-1}, d_{t-2} = d' \right]}{\mathbb{E} \left[ \exp \left( is_1 \frac{1}{\delta^{d'}} x_{t-1} \right) | d_{t-1} = d'', d_{t-2} = d' \right]} \times \\
&\left. \exp \left[ \int_0^{s_1/\delta^{d'}} \frac{\mathbb{E} \left[ i \left( \frac{\delta^{d'}}{\delta^{d''}} x_t - \alpha^{d''} \delta^{d'} - \beta^{d''} \delta^{d'} w_{t-1} \right) \exp(is'_1 x_{t-1}) | d_{t-1} = d'', d_{t-2} = d' \right]}{\gamma^{d''} \mathbb{E} \left[ \exp(is'_1 x_{t-1}) | d_{t-1} = d'', d_{t-2} = d' \right]} ds'_1 \right] ds_1 ds_2 dw_t \right\}
\end{aligned}$$

using any  $d'$  and  $d''$ . This completes Step 3.

**Step 4: Closed-form identification of the CCP  $f_1(d_t|w_t, x_t^*)$ :** Note that we have

$$\begin{aligned}
\mathbb{E} [\mathbb{1}\{d_t = d\} \exp(isx_t) | w_t, d_{t-1} = d'] &= \mathbb{E} \left[ \mathbb{1}\{d_t = d\} \exp \left( is\delta^{d'} x_t^* + is\varepsilon_t^{d'} \right) | w_t, d_{t-1} = d' \right] \\
&= \mathbb{E} \left[ \mathbb{1}\{d_t = d\} \exp \left( is\delta^{d'} x_t^* \right) | w_t, d_{t-1} = d' \right] \mathbb{E} \left[ \exp \left( is\varepsilon_t^{d'} \right) \right] \\
&= \mathbb{E} \left[ \mathbb{E} [\mathbb{1}\{d_t = d\} | w_t, x_t^*, d_{t-1} = d'] \exp \left( is\delta^{d'} x_t^* \right) | w_t, d_{t-1} = d' \right] \mathbb{E} \left[ \exp \left( is\varepsilon_t^{d'} \right) \right]
\end{aligned}$$

by the independence assumption for  $\varepsilon_t^{d'}$  and the law of iterated expectations. Therefore,

$$\begin{aligned}
&\frac{\mathbb{E} [\mathbb{1}\{d_t = d\} \exp(isx_t) | w_t, d_{t-1} = d']}{\mathbb{E} [\exp(is\varepsilon_t^{d'})]} \\
&= \mathbb{E} \left[ \mathbb{E} [\mathbb{1}\{d_t = d\} | w_t, x_t^*, d_{t-1} = d'] \exp \left( is\delta^{d'} x_t^* \right) | w_t, d_{t-1} = d' \right] \\
&= \int \exp \left( is\delta^{d'} x_t^* \right) \mathbb{E} [\mathbb{1}\{d_t = d\} | w_t, x_t^*, d_{t-1} = d'] f(x_t^* | w_t, d_{t-1} = d') dx_t^*
\end{aligned}$$

and rescaling  $s$  yields

$$\begin{aligned} & \frac{\mathbb{E} \left[ \mathbb{1}\{d_t = d\} \exp \left( i s \frac{1}{\delta^{d'}} x_t \right) \mid w_t, d_{t-1} = d' \right]}{\mathbb{E} \left[ \exp \left( i s \frac{1}{\delta^{d'}} \varepsilon_t^{d'} \right) \right]} \\ &= \int \exp (i s x_t^*) \mathbb{E} [\mathbb{1}\{d_t = d\} \mid w_t, x_t^*, d_{t-1} = d'] f (x_t^* \mid w_t, d_{t-1} = d') dx_t^* \end{aligned}$$

This is the Fourier inversion of  $\mathbb{E} [\mathbb{1}\{d_t = d\} \mid w_t, x_t^*, d_{t-1} = d'] f (x_t^* \mid w_t, d_{t-1} = d')$ . On the other hand, the Fourier inversion of  $f (x_t^* \mid w_t, d_{t-1})$  can be found as

$$\mathbb{E} [\exp (i s x_t^*) \mid w_t, d_{t-1} = d'] = \frac{\mathbb{E} \left[ \exp \left( i s \frac{1}{\delta^{d'}} x_t \right) \mid w_t, d_{t-1} = d' \right]}{\mathbb{E} \left[ \exp \left( i s \frac{1}{\delta^{d'}} \varepsilon_t^{d'} \right) \right]}.$$

Therefore, we find the closed-form expression for CCP  $f_1 (d_t \mid w_t, x_t^*)$  as follows.

$$\begin{aligned} \Pr (d_t = d \mid w_t, x_t^*) &= \sum_{d'} \Pr (d_t = d \mid w_t, x_t^*, d_{t-1} = d') \Pr (d_{t-1} = d' \mid w_t, x_t^*) \\ &= \sum_{d'} \mathbb{E} [\mathbb{1}\{d_t = d\} \mid w_t, x_t^*, d_{t-1} = d'] \Pr (d_{t-1} = d' \mid w_t, x_t^*) \\ &= \sum_{d'} \frac{\mathbb{E} [\mathbb{1}\{d_t = d\} \mid w_t, x_t^*, d_{t-1} = d'] f (x_t^* \mid w_t, d_{t-1} = d')}{f (x_t^* \mid w_t, d_{t-1} = d')} \Pr (d_{t-1} = d' \mid w_t, x_t^*) \\ &= \sum_{d'} \frac{(\mathcal{F} \phi_{(d)x_t^* \mid w_t(d')})(x_t^*)}{(\mathcal{F} \phi_{x_t^* \mid w_t(d')})(x_t^*)} \Pr (d_{t-1} = d' \mid w_t, x_t^*) \end{aligned}$$

where the characteristic functions are defined by

$$\begin{aligned} \phi_{(d)x_t^* \mid w_t(d')}(s) &= \frac{\mathbb{E} \left[ \mathbb{1}\{d_t = d\} \exp \left( i s \frac{1}{\delta^{d'}} x_t \right) \mid w_t, d_{t-1} = d' \right]}{\mathbb{E} \left[ \exp \left( i s \frac{1}{\delta^{d'}} \varepsilon_t^{d'} \right) \right]} \\ &= \mathbb{E} \left[ \mathbb{1}\{d_t = d\} \exp \left( i s \frac{1}{\delta^{d'}} x_t \right) \mid w_t, d_{t-1} = d' \right] \\ &\quad \times \frac{\exp \left[ \int_0^{s/\delta^{d'}} \frac{\mathbb{E} \left[ i \left( \frac{\delta^{d'}}{\delta^{d''}} x_{t+1} - \alpha^{d''} \delta^{d'} - \beta^{d''} \delta^{d'} w_t \right) \exp (i s_1 x_t) \mid d_t = d'', d_{t-1} = d' \right]}{\gamma^{d''} \mathbb{E} [\exp (i s_1 x_t) \mid d_t = d'', d_{t-1} = d']} ds_1 \right]}{\mathbb{E} \left[ \exp \left( i s \frac{1}{\delta^{d'}} x_t \right) \mid d_t = d', d_{t-1} = d' \right]} \end{aligned}$$

and

$$\begin{aligned}
\phi_{x_t^*|w_t(d')}(s) &= \frac{\mathbb{E} \left[ \exp \left( is \frac{1}{\delta^{d'}} x_t \right) | w_t \right]}{\mathbb{E} \left[ \exp \left( is \frac{1}{\delta^{d'}} \varepsilon_t^{d'} \right) \right]} \\
&= \frac{\mathbb{E} \left[ \exp \left( is \frac{1}{\delta^{d'}} x_t \right) | w_t \right] \cdot \exp \left[ \int_0^{s/\delta^{d'}} \frac{\mathbb{E} \left[ i \left( \frac{\delta^{d'}}{\delta^{d''}} x_{t+1} - \alpha^{d''} \delta^{d'} - \beta^{d''} \delta^{d'} w_t \right) \exp(is_1 x_t) | d_t = d'', d_{t-1} = d' \right]}{\gamma^{d''} \mathbb{E}[\exp(is_1 x_t) | d_t = d'', d_{t-1} = d']} ds_1 \right]}{\mathbb{E} \left[ \exp \left( is \frac{1}{\delta^{d'}} x_t \right) | d_t = d', d_{t-1} = d'' \right]}
\end{aligned}$$

by (B.2) using any  $d''$ . In summary, we obtain the closed-form expression

$$\begin{aligned}
\Pr(d_t = d | w_t, x_t^*) &= \sum_{d'} \frac{(\mathcal{F}\phi_{(d)x_t^*|w_t(d')})(x_t^*)}{(\mathcal{F}\phi_{x_t^*|w_t(d')})(x_t^*)} \Pr(d_{t-1} = d' | w_t, x_t^*) \\
&= \sum_{d'} \Pr(d_{t-1} = d' | w_t, x_t^*) \int \exp(-isx_t^*) \times \\
&\quad \mathbb{E} \left[ \mathbb{1}\{d_t = d\} \exp \left( is \frac{1}{\delta^{d'}} x_t \right) | w_t, d_{t-1} = d' \right] \times \\
&\quad \exp \left[ \int_0^{s/\delta^{d'}} \frac{\mathbb{E} \left[ i \left( \frac{\delta^{d'}}{\delta^{d''}} x_{t+1} - \alpha^{d''} \delta^{d'} - \beta^{d''} \delta^{d'} w_t \right) \exp(is_1 x_t) | d_t = d'', d_{t-1} = d' \right]}{\gamma^{d''} \mathbb{E}[\exp(is_1 x_t) | d_t = d'', d_{t-1} = d']} ds_1 \right] ds / \\
&\quad \mathbb{E} \left[ \exp \left( is \frac{1}{\delta^{d'}} x_t \right) | d_t = d', d_{t-1} = d'' \right] \\
&\quad \int \exp(-isx_t^*) \cdot \mathbb{E} \left[ \exp \left( is \frac{1}{\delta^{d'}} x_t \right) | w_t \right] \times \\
&\quad \exp \left[ \int_0^{s/\delta^{d'}} \frac{\mathbb{E} \left[ i \left( \frac{\delta^{d'}}{\delta^{d''}} x_{t+1} - \alpha^{d''} \delta^{d'} - \beta^{d''} \delta^{d'} w_t \right) \exp(is_1 x_t) | d_t = d'', d_{t-1} = d' \right]}{\gamma^{d''} \mathbb{E}[\exp(is_1 x_t) | d_t = d'', d_{t-1} = d']} ds_1 \right] ds. \\
&\quad \mathbb{E} \left[ \exp \left( is \frac{1}{\delta^{d'}} x_t \right) | d_t = d', d_{t-1} = d'' \right]
\end{aligned}$$

This completes Step 4. □

### B.3 Proof of Theorem 4

Our closed-form identification includes four steps.

**Step 1: Closed-form identification of the transition rule  $f_2(w_t, x_t^* | d_{t-1}, w_{t-1}, x_{t-1}^*)$ :**

First, we show the identification of the parameters and the distributions in transition of  $x_t^*$ .

Since

$$\begin{aligned}
x_t &= x_t^* + \varepsilon_t = \sum_d \mathbb{1}\{d_{t-1} = d\} [\alpha^d + \beta^d w_{t-1} + \gamma^d x_{t-1}^* + \eta_t^d] + \varepsilon_t \\
&= \sum_d \mathbb{1}\{d_{t-1} = d\} [\alpha^d + \beta^d w_{t-1} + \gamma^d x_{t-1} + \eta_t^d - \gamma^d \varepsilon_{t-1}] + \varepsilon_t
\end{aligned}$$

we obtain the following equalities for each  $d$ :

$$\begin{aligned}
\mathbb{E}[x_t \mid d_{t-1} = d] &= \alpha^d + \beta^d \mathbb{E}[w_{t-1} \mid d_{t-1} = d] + \gamma^d \mathbb{E}[x_{t-1} \mid d_{t-1} = d] \\
&\quad - \mathbb{E}[\gamma^d \varepsilon_{t-1} \mid d_{t-1} = d] + \mathbb{E}[\eta_t^d \mid d_{t-1} = d] + \mathbb{E}[\varepsilon_t \mid d_{t-1} = d] \\
&= \alpha^d + \beta^d \mathbb{E}[w_{t-1} \mid d_{t-1} = d] + \gamma^d \mathbb{E}[x_{t-1} \mid d_{t-1} = d] \\
\mathbb{E}[x_t d_{t-2} \mid d_{t-1} = d] &= \alpha^d \mathbb{E}[d_{t-2} \mid d_{t-1} = d] + \beta^d \mathbb{E}[w_{t-1} d_{t-2} \mid d_{t-1} = d] + \gamma^d \mathbb{E}[x_{t-1} d_{t-2} \mid d_{t-1} = d] \\
&\quad - \mathbb{E}[\gamma^d \varepsilon_{t-1} d_{t-2} \mid d_{t-1} = d] + \mathbb{E}[\eta_t^d d_{t-2} \mid d_{t-1} = d] + \mathbb{E}[\varepsilon_t d_{t-2} \mid d_{t-1} = d] \\
&= \alpha^d \mathbb{E}[d_{t-2} \mid d_{t-1} = d] + \beta^d \mathbb{E}[w_{t-1} d_{t-2} \mid d_{t-1} = d] + \gamma^d \mathbb{E}[x_{t-1} d_{t-2} \mid d_{t-1} = d] \\
\mathbb{E}[x_t d_{t-3} \mid d_{t-1} = d] &= \alpha^d \mathbb{E}[d_{t-3} \mid d_{t-1} = d] + \beta^d \mathbb{E}[w_{t-1} d_{t-3} \mid d_{t-1} = d] + \gamma^d \mathbb{E}[x_{t-1} d_{t-3} \mid d_{t-1} = d] \\
&\quad - \mathbb{E}[\gamma^d \varepsilon_{t-1} d_{t-3} \mid d_{t-1} = d] + \mathbb{E}[\eta_t^d d_{t-3} \mid d_{t-1} = d] + \mathbb{E}[\varepsilon_t d_{t-3} \mid d_{t-1} = d] \\
&= \alpha^d \mathbb{E}[d_{t-3} \mid d_{t-1} = d] + \beta^d \mathbb{E}[w_{t-1} d_{t-3} \mid d_{t-1} = d] + \gamma^d \mathbb{E}[x_{t-1} d_{t-3} \mid d_{t-1} = d]
\end{aligned}$$

by the independence and zero mean assumptions for  $\eta_t^d$  and  $\varepsilon_t$ . From these, we have the linear equation

$$\begin{bmatrix} \mathbb{E}[x_t \mid d_{t-1} = d] \\ \mathbb{E}[x_t d_{t-2} \mid d_{t-1} = d] \\ \mathbb{E}[x_t d_{t-3} \mid d_{t-1} = d] \end{bmatrix} = \begin{bmatrix} 1 & \mathbb{E}[w_{t-1} \mid d_{t-1} = d] & \mathbb{E}[x_{t-1} \mid d_{t-1} = d] \\ \mathbb{E}[d_{t-2} \mid d_{t-1} = d] & \mathbb{E}[w_{t-1} d_{t-2} \mid d_{t-1} = d] & \mathbb{E}[x_{t-1} d_{t-2} \mid d_{t-1} = d] \\ \mathbb{E}[d_{t-3} \mid d_{t-1} = d] & \mathbb{E}[w_{t-1} d_{t-3} \mid d_{t-1} = d] & \mathbb{E}[x_{t-1} d_{t-3} \mid d_{t-1} = d] \end{bmatrix} \begin{bmatrix} \alpha^d \\ \beta^d \\ \gamma^d \end{bmatrix}$$

Provided that the matrix on the right-hand side is non-singular, we can identify the parameters  $(\alpha^d, \beta^d, \gamma^d)$  by

$$\begin{bmatrix} \alpha^d \\ \beta^d \\ \gamma^d \end{bmatrix} = \begin{bmatrix} 1 & \mathbb{E}[w_{t-1} \mid d_{t-1} = d] & \mathbb{E}[x_{t-1} \mid d_{t-1} = d] \\ \mathbb{E}[d_{t-2} \mid d_{t-1} = d] & \mathbb{E}[w_{t-1} d_{t-2} \mid d_{t-1} = d] & \mathbb{E}[x_{t-1} d_{t-2} \mid d_{t-1} = d] \\ \mathbb{E}[d_{t-3} \mid d_{t-1} = d] & \mathbb{E}[w_{t-1} d_{t-3} \mid d_{t-1} = d] & \mathbb{E}[x_{t-1} d_{t-3} \mid d_{t-1} = d] \end{bmatrix}^{-1} \begin{bmatrix} \mathbb{E}[x_t \mid d_{t-1} = d] \\ \mathbb{E}[x_t d_{t-2} \mid d_{t-1} = d] \\ \mathbb{E}[x_t d_{t-3} \mid d_{t-1} = d] \end{bmatrix}$$

Next, we show identification of the distributions of  $\varepsilon_t$  and  $\eta_t^d$  for each  $d$ . Observe that

$$\begin{aligned}
& \mathbb{E} [\exp (i s_1 x_{t-1} + i s_2 x_t) | d_{t-1} = d] \\
&= \mathbb{E} [\exp (i s_1 (x_{t-1}^* + \varepsilon_{t-1}) + i s_2 (\alpha^d + \beta^d w_{t-1} + \gamma^d x_{t-1}^* + \eta_t^d + \varepsilon_t)) | d_{t-1} = d] \\
&= \mathbb{E} [\exp (i (s_1 x_{t-1}^* + s_2 \alpha^d + s_2 \beta^d w_{t-1} + s_2 \gamma^d x_{t-1}^*)) | d_{t-1} = d] \\
&\quad \times \mathbb{E} [\exp (i s_1 \varepsilon_{t-1})] \mathbb{E} [\exp (i s_2 (\eta_t^d + \varepsilon_t))]
\end{aligned}$$

follows from the independence assumptions for  $\eta_t^d$  and  $\varepsilon_t$ . Taking the derivative with respect to  $s_2$  yields

$$\begin{aligned}
& \left[ \frac{\partial}{\partial s_2} \ln \mathbb{E} [\exp (i s_1 x_{t-1} + i s_2 x_t) | d_{t-1} = d] \right]_{s_2=0} \\
&= \frac{\mathbb{E} [i (\alpha^d + \beta^d w_{t-1} + \gamma^d x_{t-1}^*) \exp (i s_1 x_{t-1}^*) | d_{t-1} = d]}{\mathbb{E} [\exp (i s_1 x_{t-1}^*) | d_{t-1} = d]} \\
&= i \alpha^d + \beta^d \frac{\mathbb{E} [i w_{t-1} \exp (i s_1 x_{t-1}^*) | d_{t-1} = d]}{\mathbb{E} [\exp (i s_1 x_{t-1}^*) | d_{t-1} = d]} + \gamma^d \frac{\partial}{\partial s_1} \ln \mathbb{E} [\exp (i s_1 x_{t-1}^*) | d_{t-1} = d] \\
&= i \alpha^d + \beta^d \frac{\mathbb{E} [i w_{t-1} \exp (i s_1 x_{t-1}) | d_{t-1} = d]}{\mathbb{E} [\exp (i s_1 x_{t-1}) | d_{t-1} = d]} + \gamma^d \frac{\partial}{\partial s_1} \ln \mathbb{E} [\exp (i s_1 x_{t-1}^*) | d_{t-1} = d]
\end{aligned}$$

where the switch of the differential and integral operators is permissible provided that there exists  $h \in L^1(F_{w_{t-1} x_{t-1}^* | d_{t-1}=d})$  such that  $|i (\alpha^d + \beta^d w_{t-1} + \gamma^d x_{t-1}^*) \exp (i s_1 x_{t-1}^*)| < h(w_{t-1}, x_{t-1}^*)$  holds for all  $(w_{t-1}, x_{t-1}^*)$ , which follows from the bounded conditional moment given in Assumption 10, and the denominators are nonzero as the conditional characteristic function of  $x_t^*$  given  $d_t$  does not vanish on the real line under Assumption 10. Therefore,

$$\begin{aligned}
\mathbb{E} [\exp (i s x_{t-1}^*) | d_{t-1} = d] &= \exp \left[ \int_0^s \left[ \frac{1}{\gamma^d} \frac{\partial}{\partial s_2} \ln \mathbb{E} [\exp (i s_1 x_{t-1} + i s_2 x_t) | d_{t-1} = d] \right]_{s_2=0} ds_1 \right. \\
&\quad \left. - \int_0^s \frac{i \alpha^d}{\gamma^d} ds_1 - \int_0^s \frac{\beta^d \mathbb{E} [i w_{t-1} \exp (i s_1 x_{t-1}) | d_{t-1} = d]}{\gamma^d \mathbb{E} [\exp (i s_1 x_{t-1}) | d_{t-1} = d]} ds_1 \right] \\
&= \exp \left[ \int_0^s \frac{\mathbb{E} [i (x_t - \alpha^d - \beta^d w_{t-1}) \exp (i s_1 x_{t-1}) | d_{t-1} = d]}{\gamma^d \mathbb{E} [\exp (i s_1 x_{t-1}) | d_{t-1} = d]} ds_1 \right].
\end{aligned}$$

From the proxy model and the independence assumption for  $\varepsilon_t$ ,

$$\mathbb{E} [\exp (i s x_{t-1}) | d_{t-1} = d] = \mathbb{E} [\exp (i s x_{t-1}^*) | d_{t-1} = d] \mathbb{E} [\exp (i s \varepsilon_{t-1})].$$

We then obtain the following result using any  $d$ .

$$\begin{aligned} \mathbb{E}[\exp(is\varepsilon_{t-1})] &= \frac{\mathbb{E}[\exp(isx_{t-1}) | d_{t-1} = d]}{\mathbb{E}[\exp(isx_{t-1}^*) | d_{t-1} = d]} \\ &= \frac{\mathbb{E}[\exp(isx_{t-1}) | d_{t-1} = d]}{\exp\left[\int_0^s \frac{\mathbb{E}[i(x_t - \alpha^d - \beta^d w_{t-1}) \exp(is_1 x_{t-1}) | d_{t-1} = d]}{\gamma^d \mathbb{E}[\exp(is_1 x_{t-1}) | d_{t-1} = d]} ds_1\right]}. \end{aligned}$$

This argument holds for all  $t$  so that we can identify the characteristic function of  $\varepsilon_t$  with

$$\mathbb{E}[\exp(is\varepsilon_t)] = \frac{\mathbb{E}[\exp(isx_t) | d_t = d]}{\exp\left[\int_0^s \frac{\mathbb{E}[i(x_{t+1} - \alpha^d - \beta^d w_t) \exp(is_1 x_t) | d_t = d]}{\gamma^d \mathbb{E}[\exp(is_1 x_t) | d_t = d]} ds_1\right]} \quad (\text{B.3})$$

using any  $d$ .

In order to identify  $f_{w_t, \eta_t^d}$  for each  $d$ , consider

$$x_t + \gamma^d \varepsilon_{t-1} = \alpha^d + \beta^d w_{t-1} + \gamma^d x_{t-1} + \varepsilon_t + \eta_t^d,$$

and thus

$$\begin{aligned} \mathbb{E}[\exp(isx_t) | d_{t-1} = d, w_t] \mathbb{E}[\exp(is\gamma^d \varepsilon_{t-1})] &= \mathbb{E}[\exp(is(\alpha^d + \beta^d w_{t-1} + \gamma^d x_{t-1})) | d_{t-1} = d, w_t] \\ &\quad \times \mathbb{E}[\exp(is\eta_t^d) | w_t] \mathbb{E}[\exp(is\varepsilon_t)] \end{aligned}$$

follows by the independence assumptions for  $\eta_t^d$  and  $\varepsilon_t$ . Therefore, by the formula (B.3), the conditional characteristic function of  $\eta_t^d$  given  $w_t$  can be expressed by

$$\begin{aligned} \mathbb{E}[\exp(is\eta_t^d) | w_t] &= \frac{\mathbb{E}[\exp(isx_t) | d_{t-1} = d, w_t] \cdot \mathbb{E}[\exp(is\gamma^d \varepsilon_{t-1})]}{\mathbb{E}[\exp(is(\alpha^d + \beta^d w_{t-1} + \gamma^d x_{t-1})) | d_{t-1} = d, w_t] \mathbb{E}[\exp(is\varepsilon_t)]} \\ &= \frac{\mathbb{E}[\exp(isx_t) | d_{t-1} = d, w_t] \cdot \exp\left[\int_0^s \frac{\mathbb{E}[i(x_{t+1} - \alpha^d - \beta^d w_t) \exp(is_1 x_t) | d_t = d]}{\gamma^d \mathbb{E}[\exp(is_1 x_t) | d_t = d]} ds_1\right]}{\mathbb{E}[\exp(is(\alpha^d + \beta^d w_{t-1} + \gamma^d x_{t-1})) | d_{t-1} = d, w_t] \cdot \mathbb{E}[\exp(isx_t) | d_t = d]} \times \\ &\quad \frac{\mathbb{E}[\exp(is\gamma^d x_{t-1}) | d_{t-1} = d]}{\exp\left[\int_0^{s\gamma^d} \frac{\mathbb{E}[i(x_t - \alpha^d - \beta^d w_{t-1}) \exp(is_1 x_{t-1}) | d_{t-1} = d]}{\gamma^d \mathbb{E}[\exp(is_1 x_{t-1}) | d_{t-1} = d]} ds_1\right]}. \end{aligned}$$

The denominator on the right-hand side is non-zero, as the conditional and unconditional characteristic functions do not vanish on the real line under Assumption 10. Letting  $\mathcal{F}$  denote

the operator defined by

$$(\mathcal{F}\phi)(\xi) = \frac{1}{2\pi} \int e^{-is\xi} \phi(s) ds \quad \text{for all } \phi \in L^1(\mathbb{R}) \text{ and } \xi \in \mathbb{R},$$

we identify  $f_{w_t, \eta_t^d}$  by

$$f_{w_t, \eta_t^d}(w, \eta) = f_{w_t}(w) \left( \mathcal{F}\phi_{\eta_t^d|w_t}(\cdot | w) \right) (\eta) \quad \text{for all } \eta,$$

where the conditional characteristic function  $\phi_{\eta_t^d|w_t}$  is given by

$$\begin{aligned} \phi_{\eta_t^d|w_t}(s | w) &= \frac{\mathbb{E}[\exp(isx_t) | d_{t-1} = d, w_t = w] \cdot \exp\left[\int_0^s \frac{\mathbb{E}[i(x_{t+1} - \alpha^d - \beta^d w_t) \exp(is_1 x_t) | d_t = d]}{\gamma^d \mathbb{E}[\exp(is_1 x_t) | d_t = d]} ds_1\right]}{\mathbb{E}[\exp(is(\alpha^d + \beta^d w_{t-1} + \gamma^d x_{t-1})) | d_{t-1} = d, w_t = w] \cdot \mathbb{E}[\exp(isx_t) | d_t = d]} \times \\ &\quad \frac{\mathbb{E}[\exp(is\gamma^d x_{t-1}) | d_{t-1} = d]}{\exp\left[\int_0^{s\gamma^d} \frac{\mathbb{E}[i(x_t - \alpha^d - \beta^d w_{t-1}) \exp(is_1 x_{t-1}) | d_{t-1} = d]}{\gamma^d \mathbb{E}[\exp(is_1 x_{t-1}) | d_{t-1} = d]} ds_1\right]}. \end{aligned}$$

We can use this identified joint density in turn to identify the transition rule  $f_2(w_t, x_t^* | d_{t-1}, w_{t-1}, x_{t-1}^*)$  with

$$f_2(w_t, x_t^* | d_{t-1}, w_{t-1}, x_{t-1}^*) = \sum_d \mathbb{1}\{d_{t-1} = d\} f_{w_t, \eta_t^d}(w_t, x_t^* - \alpha^d - \beta^d w_{t-1} - \gamma^d x_{t-1}^*).$$

In summary, we obtain the closed-form expression

$$\begin{aligned} f_2(w_t, x_t^* | d_{t-1}, w_{t-1}, x_{t-1}^*) &= \sum_d \mathbb{1}\{d_{t-1} = d\} f_{w_t}(w_t) \left( \mathcal{F}\phi_{\eta_t^d|w_t}(\cdot | w_t) \right) (x_t^* - \alpha^d - \beta^d w_{t-1} - \gamma^d x_{t-1}^*) \\ &= \sum_d \frac{\mathbb{1}\{d_{t-1} = d\}}{2\pi} f_{w_t}(w_t) \int \exp(-is(x_t^* - \alpha^d - \beta^d w_{t-1} - \gamma^d x_{t-1}^*)) \times \\ &\quad \frac{\mathbb{E}[\exp(isx_t) | d_{t-1} = d, w_t] \cdot \exp\left[\int_0^s \frac{\mathbb{E}[i(x_{t+1} - \alpha^{d'} - \beta^{d'} w_t) \exp(is_1 x_t) | d_t = d']}{\gamma^{d'} \mathbb{E}[\exp(is_1 x_t) | d_t = d']} ds_1\right]}{\mathbb{E}[\exp(is(\alpha^d + \beta^d w_{t-1} + \gamma^d x_{t-1})) | d_{t-1} = d, w_t] \cdot \mathbb{E}[\exp(isx_t) | d_t = d]} \times \\ &\quad \frac{\mathbb{E}[\exp(is\gamma^d x_{t-1}) | d_{t-1} = d']}{\exp\left[\int_0^{s\gamma^d} \frac{\mathbb{E}[i(x_t - \alpha^{d'} - \beta^{d'} w_{t-1}) \exp(is_1 x_{t-1}) | d_{t-1} = d']}{\gamma^{d'} \mathbb{E}[\exp(is_1 x_{t-1}) | d_{t-1} = d']} ds_1\right]} ds. \end{aligned}$$

using any  $d'$ . This completes Step 1.

**Step 2: Closed-form identification of the proxy model  $f_3(x_t | x_t^*)$ :** Given (B.3), we can write the density of  $\varepsilon_t$  by

$$f_{\varepsilon_t}(\varepsilon) = (\mathcal{F}\phi_{\varepsilon_t})(\varepsilon) \quad \text{for all } \varepsilon,$$

where the characteristic function  $\phi_{\varepsilon_t}$  is defined by (B.3) as

$$\phi_{\varepsilon_t}(s) = \frac{\mathbb{E}[\exp(isx_t) | d_t = d]}{\exp\left[\int_0^s \frac{\mathbb{E}[i(x_{t+1} - \alpha^d - \beta^d w_t) \exp(is'x_t) | d_t = d]}{\gamma^d \mathbb{E}[\exp(is'x_t) | d_t = d]} ds'\right]}.$$

Provided this identified density of  $\varepsilon_t$ , we nonparametrically identify the proxy model

$$f_3(x_t | x_t^*) = f_{\varepsilon_t}(x_t - x_t^*)$$

In summary, we obtain the closed-form expression

$$\begin{aligned} f_3(x_t | x_t^*) &= (\mathcal{F}\phi_{\varepsilon_t})(x_t - x_t^*) \\ &= \frac{1}{2\pi} \int \frac{\exp(-is(x_t - x_t^*)) \cdot \mathbb{E}[\exp(isx_t) | d_t = d]}{\exp\left[\int_0^s \frac{\mathbb{E}[i(x_{t+1} - \alpha^d - \beta^d w_t) \exp(is_1x_t) | d_t = d]}{\gamma^d \mathbb{E}[\exp(is_1x_t) | d_t = d]} ds_1\right]} ds \end{aligned}$$

using any  $d$ . This completes Step 2.

**Step 3: Closed-form identification of the CCP  $f_1(d_t | w_t, x_t^*)$ :** Note that we have

$$\begin{aligned} \mathbb{E}[\mathbb{1}\{d_t = d\} \exp(isx_t) | w_t] &= \mathbb{E}[\mathbb{1}\{d_t = d\} \exp(isx_t^* + is\varepsilon_t) | w_t] \\ &= \mathbb{E}[\mathbb{1}\{d_t = d\} \exp(isx_t^*) | w_t] \mathbb{E}[\exp(is\varepsilon_t)] \\ &= \mathbb{E}[\mathbb{E}[\mathbb{1}\{d_t = d\} | w_t, x_t^*] \exp(isx_t^*) | w_t] \mathbb{E}[\exp(is\varepsilon_t)] \end{aligned}$$

by the independence assumption for  $\varepsilon_t$  and the law of iterated expectations. Therefore

$$\begin{aligned} \frac{\mathbb{E}[\mathbb{1}\{d_t = d\} \exp(isx_t) | w_t]}{\mathbb{E}[\exp(is\varepsilon_t)]} &= \mathbb{E}[\mathbb{E}[\mathbb{1}\{d_t = d\} | w_t, x_t^*] \exp(isx_t^*) | w_t] \\ &= \int \exp(isx_t^*) \mathbb{E}[\mathbb{1}\{d_t = d\} | w_t, x_t^*] f(x_t^* | w_t) dx_t^* \end{aligned}$$

This is the Fourier inversion of  $\mathbb{E}[\mathbb{1}\{d_t = d\}|w_t, x_t^*] f(x_t^*|w_t)$ . On the other hand, the Fourier inversion of  $f(x_t^*|w_t)$  can be found as

$$\mathbb{E}[\exp(isx_t^*)|w_t] = \frac{\mathbb{E}[\exp(isx_t)|w_t]}{\mathbb{E}[\exp(is\varepsilon_t)]}.$$

Therefore, we find the closed-form expression for CCP  $f_1(d_t|w_t, x_t^*)$  as follows.

$$\Pr(d_t = d|w_t, x_t^*) = \mathbb{E}[\mathbb{1}\{d_t = d\}|w_t, x_t^*] = \frac{\mathbb{E}[\mathbb{1}\{d_t = d\}|w_t, x_t^*] f(x_t^*|w_t)}{f(x_t^*|w_t)} = \frac{(\mathcal{F}\phi_{(d)x_t^*|w_t})(x_t^*)}{(\mathcal{F}\phi_{x_t^*|w_t})(x_t^*)}$$

where the characteristic functions are defined by

$$\begin{aligned} \phi_{(d)x_t^*|w_t}(s) &= \frac{\mathbb{E}[\mathbb{1}\{d_t = d\} \exp(isx_t)|w_t]}{\mathbb{E}[\exp(is\varepsilon_t)]} \\ &= \frac{\mathbb{E}[\mathbb{1}\{d_t = d\} \exp(isx_t)|w_t] \cdot \exp\left[\int_0^s \frac{\mathbb{E}[i(x_{t+1} - \alpha^{d'} - \beta^{d'} w_t) \exp(is_1 x_t)|d_t = d']}{\gamma^{d'} \mathbb{E}[\exp(is_1 x_t)|d_t = d']} ds_1\right]}{\mathbb{E}[\exp(isx_t)|d_t = d']} \end{aligned}$$

and

$$\begin{aligned} \phi_{x_t^*|w_t}(s) &= \frac{\mathbb{E}[\exp(isx_t)|w_t]}{\mathbb{E}[\exp(is\varepsilon_t)]} \\ &= \frac{\mathbb{E}[\exp(isx_t)|w_t] \cdot \exp\left[\int_0^s \frac{\mathbb{E}[i(x_{t+1} - \alpha^{d'} - \beta^{d'} w_t) \exp(is_1 x_t)|d_t = d']}{\gamma^{d'} \mathbb{E}[\exp(is_1 x_t)|d_t = d']} ds_1\right]}{\mathbb{E}[\exp(isx_t)|d_t = d']} \end{aligned}$$

by (B.3) using any  $d'$ . In summary, we obtain the closed-form expression

$$\begin{aligned} \Pr(d_t = d|w_t, x_t^*) &= \frac{(\mathcal{F}\phi_{(d)x_t^*|w_t})(x_t^*)}{(\mathcal{F}\phi_{x_t^*|w_t})(x_t^*)} \\ &= \int \exp(-isx_t^*) \cdot \mathbb{E}[\mathbb{1}\{d_t = d\} \exp(isx_t)|w_t] \times \\ &\quad \frac{\exp\left[\int_0^s \frac{\mathbb{E}[i(x_{t+1} - \alpha^{d'} - \beta^{d'} w_t) \exp(is_1 x_t)|d_t = d']}{\gamma^{d'} \mathbb{E}[\exp(is_1 x_t)|d_t = d']} ds_1\right]}{\mathbb{E}[\exp(isx_t)|d_t = d']} ds \Big/ \\ &\quad \int \exp(-isx_t^*) \cdot \mathbb{E}[\exp(isx_t)|w_t] \times \\ &\quad \frac{\exp\left[\int_0^s \frac{\mathbb{E}[i(x_{t+1} - \alpha^{d'} - \beta^{d'} w_t) \exp(is_1 x_t)|d_t = d']}{\gamma^{d'} \mathbb{E}[\exp(is_1 x_t)|d_t = d']} ds_1\right]}{\mathbb{E}[\exp(isx_t)|d_t = d']} ds \end{aligned}$$

using any  $d'$ . This completes Step 3.  $\square$

## B.4 Proof of Theorem 5

Our closed-form identification involves three steps.

**Step 1: Closed-form identification of  $f_{2x}(x_t^*|d_{t-1}, x_{t-1}^*)$  and  $f_{2y}(y_t^*|d_{t-1}, y_{t-1}^*)$ :** First, we show the identification of the parameters and the distributions in the transition law of  $x_t^*$ . Since

$$\begin{aligned} x_t &= x_t^* + \varepsilon_t^x = \sum_d \mathbb{1}\{d_{t-1} = d\} [\alpha^{x,d} + \gamma^{x,d} x_{t-1}^* + \eta_t^{x,d}] + \varepsilon_t^x \\ &= \sum_d \mathbb{1}\{d_{t-1} = d\} [\alpha^{x,d} + \gamma^{x,d} x_{t-1} + \eta_t^{x,d} - \gamma^{x,d} \varepsilon_{t-1}^x] + \varepsilon_t^x \end{aligned}$$

is true under Assumption 17, we obtain the following equalities for each  $d$ :

$$\begin{aligned} \mathbb{E}[x_t \mid d_{t-1} = d] &= \alpha^{x,d} + \gamma^{x,d} \mathbb{E}[x_{t-1} \mid d_{t-1} = d] - \\ &\quad \mathbb{E}[\gamma^{x,d} \varepsilon_{t-1}^x \mid d_{t-1} = d] + \mathbb{E}[\eta_t^{x,d} \mid d_{t-1} = d] + \mathbb{E}[\varepsilon_t^x \mid d_{t-1} = d] \\ &= \alpha^{x,d} + \gamma^{x,d} \mathbb{E}[x_{t-1} \mid d_{t-1} = d] \end{aligned}$$

$$\begin{aligned} \mathbb{E}[x_t d_{t-2} \mid d_{t-1} = d] &= \alpha^{x,d} \mathbb{E}[d_{t-2} \mid d_{t-1} = d] + \gamma^{x,d} \mathbb{E}[x_{t-1} d_{t-2} \mid d_{t-1} = d] - \\ &\quad \mathbb{E}[\gamma^{x,d} \varepsilon_{t-1}^x d_{t-2} \mid d_{t-1} = d] + \mathbb{E}[\eta_t^{x,d} d_{t-2} \mid d_{t-1} = d] + \mathbb{E}[\varepsilon_t^x d_{t-2} \mid d_{t-1} = d] \\ &= \alpha^{x,d} \mathbb{E}[d_{t-2} \mid d_{t-1} = d] + \gamma^{x,d} \mathbb{E}[x_{t-1} d_{t-2} \mid d_{t-1} = d] \end{aligned}$$

The independence and zero mean assumptions for  $\eta_t^{x,d}$  and  $\varepsilon_t^x$  stated in Assumption 17 are used above. We thus obtain the linear equation

$$\begin{bmatrix} \mathbb{E}[x_t \mid d_{t-1} = d] \\ \mathbb{E}[x_t d_{t-2} \mid d_{t-1} = d] \end{bmatrix} = \begin{bmatrix} 1 & \mathbb{E}[x_{t-1} \mid d_{t-1} = d] \\ \mathbb{E}[d_{t-2} \mid d_{t-1} = d] & \mathbb{E}[x_{t-1} d_{t-2} \mid d_{t-1} = d] \end{bmatrix} \begin{bmatrix} \alpha^{x,d} \\ \gamma^{x,d} \end{bmatrix}.$$

By the non-singularity of the matrix on the right-hand side stated in Assumption 18, we can identify the parameters  $(\alpha^{x,d}, \gamma^{x,d})$  by

$$\begin{bmatrix} \alpha^{x,d} \\ \gamma^{x,d} \end{bmatrix} = \begin{bmatrix} 1 & \mathbb{E}[x_{t-1} \mid d_{t-1} = d] \\ \mathbb{E}[d_{t-2} \mid d_{t-1} = d] & \mathbb{E}[x_{t-1} d_{t-2} \mid d_{t-1} = d] \end{bmatrix}^{-1} \begin{bmatrix} \mathbb{E}[x_t \mid d_{t-1} = d] \\ \mathbb{E}[x_t d_{t-2} \mid d_{t-1} = d] \end{bmatrix}$$

Next, we show the identification of the distributions of  $\varepsilon_t^x$  and  $\eta_t^{x,d}$  for each  $d$ . Observe that

$$\begin{aligned} & \mathbb{E} [\exp (i s_1 x_{t-1} + i s_2 x_t) | d_{t-1} = d] \\ &= \mathbb{E} \left[ \exp \left( i s_1 \left( x_{t-1}^* + \varepsilon_{t-1}^x \right) + i s_2 \left( \alpha^{x,d} + \gamma^{x,d} x_{t-1}^* + \eta_t^{x,d} + \varepsilon_t^x \right) \right) | d_{t-1} = d \right] \\ &= \mathbb{E} \left[ \exp \left( i \left( s_1 x_{t-1}^* + s_2 \alpha^{x,d} + s_2 \gamma^{x,d} x_{t-1}^* \right) \right) | d_{t-1} = d \right] \mathbb{E} \left[ \exp \left( i s_1 \varepsilon_{t-1}^x \right) \right] \mathbb{E} \left[ \exp \left( i s_2 \left( \eta_t^{x,d} + \varepsilon_t^x \right) \right) \right] \end{aligned}$$

follows from the independence assumptions for  $\eta_t^{x,d}$  and  $\varepsilon_t^x$  stated in Assumption 17. Taking the derivative with respect to  $s_2$  yields

$$\begin{aligned} \left[ \frac{\partial}{\partial s_2} \ln \mathbb{E} [\exp (i s_1 x_{t-1} + i s_2 x_t) | d_{t-1} = d] \right]_{s_2=0} &= \frac{\mathbb{E} \left[ i \left( \alpha^{x,d} + \gamma^{x,d} x_{t-1}^* \right) \exp \left( i s_1 x_{t-1}^* \right) | d_{t-1} = d \right]}{\mathbb{E} \left[ \exp \left( i s_1 x_{t-1}^* \right) | d_{t-1} = d \right]} \\ &= i \alpha^{x,d} + \gamma^{x,d} \frac{\partial}{\partial s_1} \ln \mathbb{E} \left[ \exp \left( i s_1 x_{t-1}^* \right) | d_{t-1} = d \right] \end{aligned}$$

where the switch of the differential and integral operators is permissible provided that there exists  $h \in L^1(F_{x_{t-1}^* | d_{t-1}=d})$  such that  $|i(\alpha^{x,d} + \gamma^{x,d} x_{t-1}^*) \exp(i s_1 x_{t-1}^*)| < h(x_{t-1}^*)$  holds for all  $x_{t-1}^*$ , which follows from the bounded conditional moment condition provided in Assumption 19, and the denominator is nonzero as the conditional characteristic function of  $x_t^*$  given  $d_t$  does not vanish on the real line under Assumption 19. Therefore, we have

$$\mathbb{E} \left[ \exp \left( i s x_{t-1}^* \right) | d_{t-1} = d \right] = \exp \left[ \int_0^s \frac{\mathbb{E} \left[ i \left( x_t - \alpha^{x,d} \right) \exp \left( i s_1 x_{t-1} \right) | d_{t-1} = d \right]}{\gamma^{x,d} \mathbb{E} \left[ \exp \left( i s_1 x_{t-1} \right) | d_{t-1} = d \right]} ds_1 \right].$$

On the other hand, from the proxy model and the independence conditions for  $\varepsilon_t$  stated in Assumption 17, we also have

$$\mathbb{E} [\exp (i s x_{t-1}) | d_{t-1} = d] = \mathbb{E} \left[ \exp \left( i s x_{t-1}^* \right) | d_{t-1} = d \right] \mathbb{E} \left[ \exp \left( i s \varepsilon_{t-1}^x \right) \right].$$

Combining the above two equations, we obtain the following identifying formula using any  $d$ .

$$\mathbb{E} \left[ \exp \left( i s \varepsilon_{t-1}^x \right) \right] = \frac{\mathbb{E} \left[ \exp \left( i s x_{t-1} \right) | d_{t-1} = d \right]}{\mathbb{E} \left[ \exp \left( i s x_{t-1}^* \right) | d_{t-1} = d \right]} = \frac{\mathbb{E} \left[ \exp \left( i s x_{t-1} \right) | d_{t-1} = d \right]}{\exp \left[ \int_0^s \frac{\mathbb{E} \left[ i \left( x_t - \alpha^{x,d} \right) \exp \left( i s_1 x_{t-1} \right) | d_{t-1} = d \right]}{\gamma^{x,d} \mathbb{E} \left[ \exp \left( i s_1 x_{t-1} \right) | d_{t-1} = d \right]} ds_1 \right]}.$$

This argument holds for all  $t$  so that we can identify the characteristic function of  $\varepsilon_t^x$  by

$$\phi_{\varepsilon_t^x}(s) = \mathbb{E} \left[ \exp \left( i s \varepsilon_t^x \right) \right] = \frac{\mathbb{E} \left[ \exp \left( i s x_t \right) | d_t = d \right]}{\exp \left[ \int_0^s \frac{\mathbb{E} \left[ i \left( x_{t+1} - \alpha^{x,d} \right) \exp \left( i s_1 x_t \right) | d_t = d \right]}{\gamma^{x,d} \mathbb{E} \left[ \exp \left( i s_1 x_t \right) | d_t = d \right]} ds_1 \right]} \quad (\text{B.4})$$

using any  $d$ .

In order to identify the distribution of  $\eta_t^{x,d}$  for each  $d$ , consider

$$x_t + \gamma^{x,d} \varepsilon_{t-1} = \alpha^{x,d} + \gamma^{x,d} x_{t-1} + \varepsilon_t^x + \eta_t^{x,d}$$

which holds under Assumption 17. From this equality,

$$\begin{aligned} \mathbb{E} [\exp(isx_t) | d_{t-1} = d] \mathbb{E} [\exp(is\gamma^{x,d} \varepsilon_{t-1})] &= \mathbb{E} [\exp(is(\alpha^{x,d} + \gamma^{x,d} x_{t-1})) | d_{t-1} = d] \\ &\times \mathbb{E} [\exp(is\eta_t^{x,d})] \mathbb{E} [\exp(is\varepsilon_t^x)] \end{aligned}$$

follows by the independence assumptions for  $\eta_t^{x,d}$  and  $\varepsilon_t^x$  stated in Assumption 17. Therefore, by the identifying formula (B.4) for  $\phi_{\varepsilon_t^x}$ , the characteristic function of  $\eta_t^{x,d}$  can be expressed by

$$\begin{aligned} \phi_{\eta_t^{x,d}}(s) &= \mathbb{E} [\exp(is\eta_t^{x,d})] = \frac{\mathbb{E} [\exp(isx_t) | d_{t-1} = d] \cdot \mathbb{E} [\exp(is\gamma^{x,d} \varepsilon_{t-1}^x)]}{\mathbb{E} [\exp(is(\alpha^{x,d} + \gamma^{x,d} x_{t-1})) | d_{t-1} = d] \mathbb{E} [\exp(is\varepsilon_t^x)]} \\ &= \frac{\mathbb{E} [\exp(isx_t) | d_{t-1} = d] \cdot \exp \left[ \int_0^s \frac{\mathbb{E}[(x_{t+1} - \alpha^{x,d}) \exp(is_1 x_t) | d_t = d]}{\gamma^{x,d} \mathbb{E}[\exp(is_1 x_t) | d_t = d]} ds_1 \right]}{\mathbb{E} [\exp(is(\alpha^{x,d} + \gamma^{x,d} x_{t-1})) | d_{t-1} = d] \cdot \mathbb{E} [\exp(isx_t) | d_t = d]} \\ &\times \frac{\mathbb{E} [\exp(is\gamma^{x,d} x_{t-1}) | d_{t-1} = d]}{\exp \left[ \int_0^{s\gamma^d} \frac{\mathbb{E}[(x_t - \alpha^{x,d}) \exp(is_1 x_{t-1}) | d_{t-1} = d]}{\gamma^{x,d} \mathbb{E}[\exp(is_1 x_{t-1}) | d_{t-1} = d]} ds_1 \right]}. \end{aligned} \quad (\text{B.5})$$

The denominator on the right-hand side is non-zero, as the conditional and unconditional characteristic functions do not vanish on the real line under Assumption 19. Letting  $\mathcal{F}$  denote the Fourier transform operator defined by

$$(\mathcal{F}\phi)(\xi) = \frac{1}{2\pi} \int e^{-is\xi} \phi(s) ds \quad \text{for all } \phi \in L^1(\mathbb{R}) \text{ and } \xi \in \mathbb{R},$$

we identify  $f_{\eta_t^{x,d}}$  by

$$f_{\eta_t^{x,d}}(\eta) = \left( \mathcal{F}\phi_{\eta_t^{x,d}} \right) (\eta) \quad \text{for all } \eta,$$

under Assumption 19, where the characteristic function  $\phi_{\eta_t^{x,d}}$  is identified in (B.5). We can use this identified density function  $f_{\eta_t^{x,d}}$  in turn to identify the transition rule  $f_{2x}(x_t^* | d_{t-1}, x_{t-1}^*)$  with

$$f_{2x}(x_t^* | d_{t-1}, x_{t-1}^*) = \sum_d \mathbb{1}\{d_{t-1} = d\} f_{\eta_t^{x,d}}(x_t^* - \alpha^{x,d} - \gamma^{x,d} x_{t-1}^*).$$

In summary, we obtain the closed-form identifying formula for the law of state transition  $f_{2x}(x_t^* | d_{t-1}, x_{t-1}^*)$ :

$$f_{2x}(x_t^* | d_{t-1}, x_{t-1}^*) = \sum_d \frac{\mathbb{1}\{d_{t-1} = d\}}{2\pi} \int \exp(-is(x_t^* - \alpha^{x,d} - \gamma^{x,d}x_{t-1}^*)) \times \\ \frac{\mathbb{E}[\exp(isx_t) | d_{t-1} = d] \cdot \exp\left[\int_0^s \frac{\mathbb{E}[i(x_{t+1} - \alpha^{x,d'}) \exp(is_1 x_t) | d_t = d']}{\gamma^{x,d'} \mathbb{E}[\exp(is_1 x_t) | d_t = d']} ds_1\right]}{\mathbb{E}[\exp(is(\alpha^{x,d} + \gamma^{x,d}x_{t-1}^*)) | d_{t-1} = d] \cdot \mathbb{E}[\exp(isx_t) | d_t = d]} \times \\ \frac{\mathbb{E}[\exp(is\gamma^{x,d}x_{t-1}^*) | d_{t-1} = d']}{\exp\left[\int_0^s \frac{\mathbb{E}[i(x_t - \alpha^{x,d'}) \exp(is_1 x_{t-1}) | d_{t-1} = d']}{\gamma^{x,d'} \mathbb{E}[\exp(is_1 x_{t-1}) | d_{t-1} = d']} ds_1\right]} ds$$

using any  $d'$ . By similar lines of argument, we also obtain the closed-form identifying formula for the law of state transition  $f_{2y}(y_t^* | d_{t-1}, y_{t-1}^*)$ :

$$f_{2y}(y_t^* | d_{t-1}, y_{t-1}^*) = \sum_d \frac{\mathbb{1}\{d_{t-1} = d\}}{2\pi} \int \exp(-is(y_t^* - \alpha^{y,d} - \gamma^{y,d}y_{t-1}^*)) \times \\ \frac{\mathbb{E}[\exp(isy_t) | d_{t-1} = d] \cdot \exp\left[\int_0^s \frac{\mathbb{E}[i(y_{t+1} - \alpha^{y,d'}) \exp(is_1 y_t) | d_t = d']}{\gamma^{y,d'} \mathbb{E}[\exp(is_1 y_t) | d_t = d']} ds_1\right]}{\mathbb{E}[\exp(is(\alpha^{y,d} + \gamma^{y,d}y_{t-1}^*)) | d_{t-1} = d] \cdot \mathbb{E}[\exp(isy_t) | d_t = d]} \times \\ \frac{\mathbb{E}[\exp(is\gamma^{y,d}y_{t-1}^*) | d_{t-1} = d']}{\exp\left[\int_0^s \frac{\mathbb{E}[i(y_t - \alpha^{y,d'}) \exp(is_1 y_{t-1}) | d_{t-1} = d']}{\gamma^{y,d'} \mathbb{E}[\exp(is_1 y_{t-1}) | d_{t-1} = d']} ds_1\right]} ds$$

using any  $d'$ . This completes Step 1.

**Step 2: Closed-form identification of  $f_{3x}(x_t | x_t^*)$  and  $f_{3y}(y_t | y_t^*)$ :** We can write the density function of  $\varepsilon_t^x$  by

$$f_{\varepsilon_t^x}(\varepsilon) = (\mathcal{F}\phi_{\varepsilon_t^x})(\varepsilon^x) \quad \text{for all } \varepsilon^x,$$

where the characteristic function  $\phi_{\varepsilon_t^x}$  is identified in (B.4) with a closed-form formula. Provided this identified density function  $f_{\varepsilon_t^x}$ , we identify the proxy model  $f_{3x}(x_t | x_t^*)$  by

$$f_{3x}(x_t | x_t^*) = f_{\varepsilon_t^x}(x_t - x_t^*).$$

In summary, we obtain the closed-form identifying formula for the proxy model  $f_{3x}(x_t | x_t^*)$ :

$$f_{3x}(x_t | x_t^*) = \frac{1}{2\pi} \int \frac{\exp(-is(x_t - x_t^*)) \cdot \mathbb{E}[\exp(isx_t) | d_t = d]}{\exp\left[\int_0^s \frac{\mathbb{E}[i(x_{t+1} - \alpha^{d,x}) \exp(is_1 x_t) | d_t = d]}{\gamma^{d,x} \mathbb{E}[\exp(is_1 x_t) | d_t = d]} ds_1\right]} ds$$

using any  $d$ . By similar lines of argument, we also obtain the closed-form identifying formula for the proxy model  $f_{3y}(y_t | y_t^*)$ :

$$f_{3y}(y_t | y_t^*) = \frac{1}{2\pi} \int \frac{\exp(-is(y_t - y_t^*)) \cdot \mathbb{E}[\exp(isy_t) | d_t = d]}{\exp\left[\int_0^s \frac{\mathbb{E}[i(y_{t+1} - \alpha^{d,y}) \exp(is_1 y_t) | d_t = d]}{\gamma^{d,y} \mathbb{E}[\exp(is_1 y_t) | d_t = d]} ds_1\right]} ds$$

using any  $d$ . This completes Step 2.

**Step 3: Closed-form identification of the CCP  $f_1(d_t | x_t^*, y_t^*)$ :** We can write

$$\begin{aligned} \mathbb{E}[\mathbb{1}\{d_t = d\} \exp(is^x x_t + is^y y_t)] &= \mathbb{E}[\mathbb{1}\{d_t = d\} \exp(is^x x_t^* + is^y y_t^* + is^x \varepsilon_t^x + is^y \varepsilon_t^y)] \\ &= \mathbb{E}[\mathbb{1}\{d_t = d\} \exp(is^x x_t^* + is^y y_t^*)] \mathbb{E}[\exp(is^x \varepsilon_t^x)] \mathbb{E}[\exp(is^y \varepsilon_t^y)] \\ &= \mathbb{E}[\mathbb{E}[\mathbb{1}\{d_t = d\} | x_t^*, y_t^*] \exp(is^x x_t^* + is^y y_t^*)] \mathbb{E}[\exp(is^x \varepsilon_t^x)] \mathbb{E}[\exp(is^y \varepsilon_t^y)] \end{aligned}$$

by the independence assumption for  $\varepsilon_t^x$  and  $\varepsilon_t^y$  stated in Assumption 17 and the law of iterated expectations. Therefore, we obtain

$$\begin{aligned} \frac{\mathbb{E}[\mathbb{1}\{d_t = d\} \exp(is^x x_t + is^y y_t)]}{\mathbb{E}[\exp(is^x \varepsilon_t^x)] \mathbb{E}[\exp(is^y \varepsilon_t^y)]} &= \mathbb{E}[\mathbb{E}[\mathbb{1}\{d_t = d\} | x_t^*, y_t^*] \exp(is^x x_t^* + is^y y_t^*)] \\ &= \int \int \exp(is^x x^* + is^y y^*) \mathbb{E}[\mathbb{1}\{d_t = d\} | (x_t^*, y_t^*) = (x^*, y^*)] f_{x_t^*, y_t^*}(x^*, y^*) dx^* dy^* \end{aligned}$$

This is the two-dimensional Fourier inversion of  $\mathbb{E}[\mathbb{1}\{d_t = d\} | (x_t^*, y_t^*) = (\cdot, \cdot)] f_{x_t^*, y_t^*}(\cdot, \cdot)$ .

On the other hand, the two-dimensional Fourier inversion of  $f_{x_t^*, y_t^*}$  can be found as

$$\mathbb{E}[\exp(is^x x_t^* + is^y y_t^*)] = \frac{\mathbb{E}[\exp(is^x x_t + is^y y_t)]}{\mathbb{E}[\exp(is^x \varepsilon_t^x)] \mathbb{E}[\exp(is^y \varepsilon_t^y)]}.$$

Letting  $\mathcal{F}^2$  denote the two-dimensional Fourier transform operator defined by

$$(\mathcal{F}^2 \phi)(\xi_1, \xi_2) = \frac{1}{4\pi^2} \int \int e^{-is_1 \xi_1 - is_2 \xi_2} \phi(s_1, s_2) ds_1 ds_2 \quad \text{for all } \phi \in L^1(\mathbb{R}^2) \text{ and } (\xi_1, \xi_2) \in \mathbb{R}^2,$$

we find the closed-form expression for CCP  $f_1(d_t | x_t^*, y_t^*)$  as follows.

$$\Pr(d_t = d | x_t^*, y_t^*) = \frac{\mathbb{E}[\mathbb{1}\{d_t = d\} | x_t^*, y_t^*] f_{x_t^*, y_t^*}(x_t^*, y_t^*)}{f_{x_t^*, y_t^*}(x_t^*, y_t^*)} = \frac{(\mathcal{F}^2 \phi_{(d)x_t^*, y_t^*})(x_t^*, y_t^*)}{(\mathcal{F}^2 \phi_{x_t^*, y_t^*})(x_t^*, y_t^*)}$$

where the ‘phi’ functions in the last expression are

$$\begin{aligned} \phi_{(d)x_t^*, y_t^*}(s^x, s^y) &= \frac{\mathbb{E}[\mathbb{1}\{d_t = d\} \exp(is^x x_t + is^y y_t)]}{\mathbb{E}[\exp(is\varepsilon_t^x)] + \mathbb{E}[\exp(is\varepsilon_t^y)]} \\ &= \mathbb{E}[\mathbb{1}\{d_t = d\} \exp(is^x x_t + is^y y_t)] \frac{\exp\left[\int_0^{s^x} \frac{\mathbb{E}[i(x_{t+1} - \alpha^{x,d'}) \exp(is_1 x_t) | d_t = d']}{\gamma^{x,d'} \mathbb{E}[\exp(is_1 x_t) | d_t = d']} ds_1\right]}{\mathbb{E}[\exp(is^x x_t) | d_t = d']} \\ &\quad \times \frac{\exp\left[\int_0^{s^y} \frac{\mathbb{E}[i(y_{t+1} - \alpha^{y,d'}) \exp(is_1 y_t) | d_t = d']}{\gamma^{y,d'} \mathbb{E}[\exp(is_1 y_t) | d_t = d']} ds_1\right]}{\mathbb{E}[\exp(is^y y_t) | d_t = d']} \end{aligned}$$

and

$$\begin{aligned} \phi_{x_t^*, y_t^*}(s^x, s^y) &= \frac{\mathbb{E}[\exp(is^x x_t + is^y y_t)]}{\mathbb{E}[\exp(is\varepsilon_t^x)] \mathbb{E}[\exp(is\varepsilon_t^y)]} = \mathbb{E}[\exp(is^x x_t + is^y y_t)] \times \\ &\quad \frac{\exp\left[\int_0^{s^x} \frac{\mathbb{E}[i(x_{t+1} - \alpha^{x,d'}) \exp(is_1 x_t) | d_t = d']}{\gamma^{x,d'} \mathbb{E}[\exp(is_1 x_t) | d_t = d']} ds_1\right]}{\mathbb{E}[\exp(is^x x_t) | d_t = d']} \frac{\exp\left[\int_0^{s^y} \frac{\mathbb{E}[i(y_{t+1} - \alpha^{y,d'}) \exp(is_1 y_t) | d_t = d']}{\gamma^{y,d'} \mathbb{E}[\exp(is_1 y_t) | d_t = d']} ds_1\right]}{\mathbb{E}[\exp(is^y y_t) | d_t = d']} \end{aligned}$$

from (B.4) using any  $d'$ . In summary, we obtain the closed-form identifying formula for the CCP  $f_1(d_t | x_t^*, y_t^*)$ :

$$\begin{aligned} \Pr(d_t = d | x_t^*, y_t^*) &= \frac{(\mathcal{F}^2 \phi_{(d)x_t^*, y_t^*})(x_t^*, y_t^*)}{(\mathcal{F}^2 \phi_{x_t^*, y_t^*})(x_t^*, y_t^*)} \\ &= \int \int \exp(-is^x x_t^* - is^y y_t^*) \cdot \mathbb{E}[\mathbb{1}\{d_t = d\} \exp(is^x x_t + is^y y_t)] \times \\ &\quad \frac{\exp\left[\int_0^{s^x} \frac{\mathbb{E}[i(x_{t+1} - \alpha^{x,d'}) \exp(is_1 x_t) | d_t = d']}{\gamma^{x,d'} \mathbb{E}[\exp(is_1 x_t) | d_t = d']} ds_1\right]}{\mathbb{E}[\exp(is^x x_t) | d_t = d']} \frac{\exp\left[\int_0^{s^y} \frac{\mathbb{E}[i(y_{t+1} - \alpha^{y,d'}) \exp(is_1 y_t) | d_t = d']}{\gamma^{y,d'} \mathbb{E}[\exp(is_1 y_t) | d_t = d']} ds_1\right]}{\mathbb{E}[\exp(is^y y_t) | d_t = d']} ds^x ds^y / \\ &\quad \int \int \exp(-is^x x_t^* - is^y y_t^*) \cdot \mathbb{E}[\exp(is^x x_t + is^y y_t)] \times \\ &\quad \frac{\exp\left[\int_0^{s^x} \frac{\mathbb{E}[i(x_{t+1} - \alpha^{x,d'}) \exp(is_1 x_t) | d_t = d']}{\gamma^{x,d'} \mathbb{E}[\exp(is_1 x_t) | d_t = d']} ds_1\right]}{\mathbb{E}[\exp(is^x x_t) | d_t = d']} \frac{\exp\left[\int_0^{s^y} \frac{\mathbb{E}[i(y_{t+1} - \alpha^{y,d'}) \exp(is_1 y_t) | d_t = d']}{\gamma^{y,d'} \mathbb{E}[\exp(is_1 y_t) | d_t = d']} ds_1\right]}{\mathbb{E}[\exp(is^y y_t) | d_t = d']} ds^x ds^y \end{aligned}$$

using any  $d'$ . This completes Step 3.  $\square$

## C Consistency of $\widehat{f}_2(x_t^* \mid d_{t-1}, x_{t-1}^*)$

This section discusses consistency of the analog estimator  $\widehat{f}_2(x_t^* \mid d_{t-1}, x_{t-1}^*)$  for the law of state transition proposed in Section 2.3. For each time period  $t$ , consider the sequence  $\{r_{N,t}\}_{N=1}^\infty$  of positive numbers such that

$$\sup_{s \in [-h_N^{-1}, h_N^{-1}]} \left| \widehat{\phi}_{\varepsilon_t}(s) - \phi_{\varepsilon_t}(s) \right| = o_p(r_{N,t}) \quad (\text{C.1})$$

holds as  $N \rightarrow \infty$ . Li and Vuong (1998) derive admissible combinations of  $\{r_{N,t}\}_N$  and  $\{h_N\}_N$  under various cases of smoothness assumptions for the distributions of  $x_t^*$  and  $\varepsilon_t$ . Also let  $\{c_{N,t}\}_{N=1}^\infty$  be the sequence of positive numbers defined by

$$c_{N,t} = \inf_{s \in [-h_N^{-1}, h_N^{-1}]} |\phi_{\varepsilon_t}(s)| \quad (\text{C.2})$$

Note that this sequence also depends on smoothness assumptions for the distribution of  $\varepsilon_t$ . We make the following assumptions.

**Assumption 20.** *The following conditions hold. (i)  $|\gamma^d| < 1$ . (ii)  $|\widehat{\gamma}^d - \gamma^d| = O_p(N^{-1/2})$ .*

*(iii)  $\left| \widehat{E}_N[\exp(isx_t) \mid d_{t-1} = d] - E[\exp(isx_t) \mid d_{t-1} = d] \right| = O_p(N^{-1/2})$ .*

*(iv)  $\left| \widehat{E}_N[\exp(is(\widehat{\alpha}^d + \widehat{\gamma}^d x_{t-1})) \mid d_{t-1} = d] - E[\exp(is(\alpha^d + \gamma^d x_{t-1})) \mid d_{t-1} = d] \right| = O_p(N^{-1/2})$ .*

*(v)  $c_{N,t} N^{1/2} \rightarrow \infty$  as  $N \rightarrow \infty$ . (vi)  $\frac{c_{N,t}}{r_{N,t}} \rightarrow \infty$  as  $N \rightarrow \infty$ . (vii)  $\phi_K(s) = \frac{1}{2} \mathbb{1}[-1 \leq s \leq 1]$ .*

Part (i) requires that the dynamics of  $x_t^*$  follows a sub-unit-root process. Parts (ii), (iii), and (vi) require parametric convergence rates for the respective parametric estimators, which is satisfied under the standard assumptions for central limit theorems. Parts (v) and (vi) require  $h_N$  to be chosen to vanish slow enough to have  $c_{N,t}$  dominate  $r_{N,t}$  as well as  $N^{-1/2}$ . Part (v) and (vi) are consistent with the restrictions imposed by Li and Vuong (1998) under various smoothness cases. Finally, part (vii) assumes that  $\phi_K$  is an indicator function to have the setting equivalent to that of Li and Vuong (1998). We obtain the following consistency results.

**Lemma 1.** *If Assumption 20 (i)–(vi) are satisfied, then*

$$\sup_{s \in [-h_N^{-1}, h_N^{-1}]} \left| \widehat{\phi}_{\eta_t^d}(s) - \phi_{\eta_t^d}(s) \right| = O_P \left( \frac{r_{N,t} \vee r_{N,t-1}}{c_{N,t}^2} \right)$$

as  $N \rightarrow \infty$ .

*Proof.* By Assumption 20 (i) and (ii) and the uniform continuity of characteristic functions, we have

$$\sup_{s \in [-h_N^{-1}, h_N^{-1}]} \left| \widehat{\phi}_{\varepsilon_{t-1}}(\widehat{\gamma}^d s) - \phi_{\varepsilon_{t-1}}(\gamma^d s) \right| \leq \sup_{s \in [-h_N^{-1}, h_N^{-1}]} \left| \widehat{\phi}_{\varepsilon_{t-1}}(s) - \phi_{\varepsilon_{t-1}}(s) \right|$$

with probability approaching one as  $N \rightarrow \infty$ , and hence

$$\sup_{s \in [-h_N^{-1}, h_N^{-1}]} \left| \widehat{\phi}_{\varepsilon_{t-1}}(\widehat{\gamma}^d s) - \phi_{\varepsilon_{t-1}}(\gamma^d s) \right| = o_p(r_{N,t-1}) \quad (\text{C.3})$$

holds as  $N \rightarrow \infty$ . Also, we can write

$$\begin{aligned} \widehat{\phi}_{\eta_t^d}(s) - \phi_{\eta_t^d}(s) &= \left[ \left( \widehat{\mathbb{E}}_N[\exp(isx_t) \mid d_{t-1} = d] - \mathbb{E}[\exp(isx_t) \mid d_{t-1} = d] \right) \cdot \left( \widehat{\phi}_{\varepsilon_{t-1}}(\widehat{\gamma}^d s) - \phi_{\varepsilon_{t-1}}(\gamma^d s) \right) \right. \\ &\quad \left. + \left( \widehat{\mathbb{E}}_N[\exp(isx_t) \mid d_{t-1} = d] - \mathbb{E}[\exp(isx_t) \mid d_{t-1} = d] \right) \cdot \phi_{\varepsilon_{t-1}}(\gamma^d s) + \mathbb{E}[\exp(isx_t) \mid d_{t-1} = d] \right. \\ &\quad \left. \times \left( \widehat{\phi}_{\varepsilon_{t-1}}(\widehat{\gamma}^d s) - \phi_{\varepsilon_{t-1}}(\gamma^d s) \right) \right] \cdot \mathbb{E}[\exp(is(\alpha^d + \gamma^d x_{t-1})) \mid d_{t-1} = d] \cdot \phi_{\varepsilon_t}(s) + \mathbb{E}[\exp(isx_t) \mid d_{t-1} = d] \\ &\quad \times \phi_{\varepsilon_{t-1}}(\gamma^d s) \cdot \left\{ \left( \widehat{\mathbb{E}}_N[\exp(is(\widehat{\alpha}^d + \widehat{\gamma}^d x_{t-1})) \mid d_{t-1} = d] - \mathbb{E}[\exp(is(\alpha^d + \gamma^d x_{t-1})) \mid d_{t-1} = d] \right) \right. \\ &\quad \left. \times \left( \widehat{\phi}_{\varepsilon_t}(s) - \phi_{\varepsilon_t}(s) \right) + \left( \widehat{\mathbb{E}}_N[\exp(is(\widehat{\alpha}^d + \widehat{\gamma}^d x_{t-1})) \mid d_{t-1} = d] - \mathbb{E}[\exp(is(\alpha^d + \gamma^d x_{t-1})) \mid d_{t-1} = d] \right) \right. \\ &\quad \left. \times \phi_{\varepsilon_t}(s) + \mathbb{E}[\exp(is(\alpha^d + \gamma^d x_{t-1})) \mid d_{t-1} = d] \cdot \left( \widehat{\phi}_{\varepsilon_t}(s) - \phi_{\varepsilon_t}(s) \right) \right\} \Big/ \\ &\quad \left[ \mathbb{E}[\exp(is(\alpha^d + \gamma^d x_{t-1})) \mid d_{t-1} = d]^2 \cdot \phi_{\varepsilon_t}(s)^2 + \mathbb{E}[\exp(is(\alpha^d + \gamma^d x_{t-1})) \mid d_{t-1} = d] \cdot \phi_{\varepsilon_t}(s) \right. \\ &\quad \left. \times \left\{ \left( \widehat{\mathbb{E}}_N[\exp(is(\widehat{\alpha}^d + \widehat{\gamma}^d x_{t-1})) \mid d_{t-1} = d] - \mathbb{E}[\exp(is(\alpha^d + \gamma^d x_{t-1})) \mid d_{t-1} = d] \right) \cdot \left( \widehat{\phi}_{\varepsilon_t}(s) - \phi_{\varepsilon_t}(s) \right) \right. \right. \\ &\quad \left. \left. + \left( \widehat{\mathbb{E}}_N[\exp(is(\widehat{\alpha}^d + \widehat{\gamma}^d x_{t-1})) \mid d_{t-1} = d] - \mathbb{E}[\exp(is(\alpha^d + \gamma^d x_{t-1})) \mid d_{t-1} = d] \right) \cdot \phi_{\varepsilon_t}(s) \right. \right. \\ &\quad \left. \left. + \mathbb{E}[\exp(is(\alpha^d + \gamma^d x_{t-1})) \mid d_{t-1} = d] \cdot \left( \widehat{\phi}_{\varepsilon_t}(s) - \phi_{\varepsilon_t}(s) \right) \right\} \right] \end{aligned}$$

Therefore, using (C.1), (C.2), (C.3), and Assumption 20 (iii)–(vi) yields

$$\sup_{s \in [-h_N^{-1}, h_N^{-1}]} \left| \widehat{\phi}_{\eta_t^d}(s) - \phi_{\eta_t^d}(s) \right| = O_P \left( \frac{r_{N,t} \vee r_{N,t-1}}{c_{N,t}^2} \right)$$

as  $N \rightarrow \infty$ . □

**Theorem 6.** *If Assumption 20 is satisfied, then*

$$\sup_{(x^*, x_-^*)} \left| \widehat{f}_2(x^* \mid d, x_-^*) - f_2(x^* \mid d, x_-^*) \right| = O_P \left( \frac{r_{N,t} \vee r_{N,t-1}}{h_N \cdot c_{N,t}^2} \right) + O \left( \int_{h_N^{-1}}^{\infty} \left| \phi_{\eta_t^d}(s) \right| ds \right)$$

as  $N \rightarrow \infty$ .

*Proof.* Assumption 20 (vii) allows us to write

$$\begin{aligned} \sup_{(x^*, x_-^*)} \left| \widehat{f}_2(x^* \mid d, x_-^*) - f_2(x^* \mid d, x_-^*) \right| &= O_P \left( \int_{-h_N^{-1}}^{h_N^{-1}} \left| \widehat{\phi}_{\eta_t^d}(s) - \phi_{\eta_t^d}(s) \right| ds \right) + O \left( \int_{h_N^{-1}}^{\infty} \left| \phi_{\eta_t^d}(s) \right| ds \right) \\ &= O_P \left( \frac{r_{N,t} \vee r_{N,t-1}}{h_N \cdot c_{N,t}^2} \right) + O \left( \int_{h_N^{-1}}^{\infty} \left| \phi_{\eta_t^d}(s) \right| ds \right) \end{aligned}$$

as  $N \rightarrow \infty$ , where the last equality follows from Lemma 1. □