## Relations of the estimates in the text to maximum likelihood estimates

The observations on the different types of polymorphism can be partitioned as $n_{1}, n_{2}$ and $n_{3}$, corresponding to type 1 , type $2 / 3$ and type $4 / 5 / 6$ events, respectively; their sum is $n$. Let $P_{a}$ be the probability that a polymorphism is ancestral. The estimate of $t$ given in the text is equivalent to equating the observed fraction of 1 polymorphisms among 1/2/3 polymorphisms to their expected value minus 0.5 , which is independent of $P_{a}$. The estimate of $P_{a}$ is obtained by comparing the overall frequency of $1 / 2 / 3$ polymorphisms among all polymorphisms to its expected value, which is dependent on $t$. We can parameterize the likelihood model for the data in terms of $\theta_{1}$ and $\theta_{2}$, where $\theta_{1}$ is the expected fraction of type 1 polymorphisms among types $1 / 2 / 3$, and $\theta_{2}$ is the expected fraction of type $1 / 2 / 3$ polymorphisms among the total. The likelihood of the data can be written as:

$$
\begin{equation*}
L=\operatorname{Pr}\left\{n_{1} \mid\left(n_{1}+n_{2}\right)\right\} \operatorname{Pr}\left\{\left(n_{1}+n_{2}\right)\right\} \tag{S1a}
\end{equation*}
$$

so that

$$
\begin{equation*}
\ln L=\ln \left(\operatorname{Pr}\left\{n_{1} \mid\left(n_{1}+n_{2}\right)\right\}\right)+\ln \left(\operatorname{Pr}\left\{\left(n_{1}+n_{2}\right)\right\}\right) \tag{S1b}
\end{equation*}
$$

where

$$
\begin{gather*}
\ln \left(\operatorname{Pr}\left\{n_{1} \mid\left(n_{1}+n_{2}\right)\right\}\right)=C_{1}+n_{1} \ln \left(\theta_{1}\right)+n_{2} \ln \left(1-\theta_{1}\right)  \tag{S2a}\\
\ln \left(\operatorname{Pr}\left\{\left(n_{1}+n_{2}\right)\right\}\right)=C_{2}+\left(n_{1}+n_{2}\right) \ln \left(\theta_{2}\right)+n_{3} \ln \left(1-\theta_{2}\right) \tag{S2b}
\end{gather*}
$$

and $\theta_{1}=P_{1} / P_{2+3}, \theta_{2}=\left(P_{1}+P_{2+3}\right) P_{\mathrm{a}}$

Differentiating with respect to $t$ and $P_{a}$, we have:

$$
\begin{gather*}
\frac{\ln L}{t}=\left(\frac{n_{1}}{1}-\frac{n_{2}}{\left(11_{1}\right)}\right) \frac{1}{t}+\left(\frac{\left(n_{1}+n_{2}\right)}{2}-\frac{n_{3}}{\left(1-2_{2}\right)}\right) \frac{2}{t}  \tag{S3a}\\
\frac{\ln L}{P_{\mathrm{a}}}=\left(\frac{\left(n_{1}+n_{2}\right)}{2}-\frac{n_{3}}{\left(1-{ }_{2}\right)}\right) \frac{2}{P_{\mathrm{a}}} \tag{S3b}
\end{gather*}
$$

For a maximum of $\ln L$, both derivatives must be zero. It follows that the multiplicand of the partial derivative of $\theta_{2}$ in equation (S3a) is zero, since this is required for equation (S3b) to equal zero. Hence, the multiplicand of the partial derivative of $\theta_{1}$ in Eqn. (S3a) is also zero.

The last result is equivalent to equating $\left(1-\theta_{1}\right) / \theta_{1}$ to $n_{2} / n_{1}$; this in turn is equivalent to 0.5 plus equation (2) of the text, so that this estimate of $t$ is also the ML estimate.

Similarly, setting equation (S3a) to zero with this estimate of $t$ yields the estimate of $P_{a}$ as $r_{\mathrm{T}}=\left(n_{1}+n_{2}\right) /\left\{\left(n_{1}+n_{2}+n_{3}\right)\left(P_{1}+P_{2+3}\right)\right\}$, the expression that was used in the text .

## Variances of the estimates

From eqn. (11b) of Charlesworth et al. (2005), the estimated a priori probability of an ancestral polymorphism is given approximately for a reasonably large sample size by:

$$
\begin{equation*}
P_{d}=\frac{1}{3}+\frac{1}{2} \exp (-t)+\frac{1}{6} \exp (-3 t) \tag{S4}
\end{equation*}
$$

so that

$$
\begin{align*}
\frac{\mathrm{d} P_{d}}{\mathrm{~d} t} & =-\frac{1}{2}\{\exp (-t)+\exp (-3 t)\} \\
& =-\frac{1}{2}\left(x^{-1}+x^{-3}\right) \tag{S5}
\end{align*}
$$

where $x=\exp (t), t=\ln (x)$.

Using the delta-method formula for the variance of a function, we have:

$$
\begin{equation*}
V\left\{P_{d}\right\} \quad \frac{1}{4}\left(x^{-1}+x^{-3}\right)^{2} V\{t\} \tag{S6}
\end{equation*}
$$

By differentiating equation (S4) above, we find that:

$$
\begin{equation*}
\frac{\mathrm{d} x}{\mathrm{~d} a}=\frac{x}{(x-2 a)} \tag{S7}
\end{equation*}
$$

so that

$$
\begin{equation*}
V\{x\} \frac{x^{2} V\{a\}}{(x-2 a)^{2}} \tag{S8}
\end{equation*}
$$

where $a=\left(0.5 f_{1}+f_{2+3}\right) / f_{1}$ and

$$
\begin{equation*}
V\{a\}=V\left\{f_{2+3} / f_{1}\right\} \tag{S9a}
\end{equation*}
$$

From equations (S4) and (S5) we find that:

$$
\begin{equation*}
V\{t\} \quad \frac{1}{x^{2}} V\{x\}=\frac{V\{a\}}{(x-2 a)^{2}} \tag{S9b}
\end{equation*}
$$

Again using the delta method, we have:

$$
\begin{equation*}
V\left\{f_{2+3}\left\{f_{1}\right\} \quad \frac{V\left\{f_{2+3}\right\}}{f_{1}^{2}}+\frac{f_{2+3}^{2} V\left\{f_{1}\right\}}{f_{1}^{4}}-\frac{2 f_{2+3} \operatorname{Cov}\left\{f_{1}, f_{2+3}\right\}}{f_{1}^{3}}\right. \tag{S10}
\end{equation*}
$$

With $k$ independent SNPs, the variances and covariances in equation (S10) are given by their multinomial expressions:

$$
\begin{equation*}
V\left\{f_{1}\right\}=f_{1}\left(1-f_{1}\right) / k, \quad V\left\{f_{2+3}\right\}=f_{2+3}\left(1-f_{2+3}\right) / k, \operatorname{Cov}=-f_{1} f_{2+3} / k \tag{S11}
\end{equation*}
$$

Substituting these into equations (S8) and (S9) thus gives us estimates of the variances of the estimates $x$ and $t$.

The estimated proportion of ancestral polymorphisms is obtained as $r_{T}$, given by the ratio of $f_{1}+f_{2+3}$ to the estimate of $P_{\mathrm{d}}$.

The variance of $r_{\mathrm{T}}$ is thus given approximately by the delta method as:

$$
\begin{equation*}
\frac{V\left\{f_{1}\right\}+V\left\{f_{2+3}\right\}+2 \operatorname{Cov}\left\{f_{1}, f_{2+3}\right\}}{P_{\mathrm{d}}^{2}}+\frac{\left(f_{1}+f_{2+3}\right) V\left\{P_{\mathrm{d}}\right\}}{P_{\mathrm{d}}^{4}}-\frac{2\left(f_{1}+f_{2+3}\right) \operatorname{Cov}\left\{\left(f_{1}+f_{2+3}\right), P_{d}\right\}}{P_{\mathrm{d}}^{3}} \tag{S12}
\end{equation*}
$$

where the variance of the estimate of $P_{\mathrm{d}}$ is obtained from equations (S4), (S9) and (S10). Using equation (S4) and the delta method, we have

$$
\begin{equation*}
\operatorname{Cov}\left\{f_{1}+f_{2+3}\right\} \quad-\frac{1}{2}\left(x^{-1}+x^{-3}\right) \operatorname{Cov}\left\{f_{1}+f_{2+3}, t\right\} \tag{S13a}
\end{equation*}
$$

where

$$
\begin{equation*}
\operatorname{Cov}\left\{f_{1}+f_{2+3}, t\right\} \quad \frac{1}{(x-2 a)} \operatorname{Cov}\left\{f_{1}+f_{2+3}, a\right\} \tag{S13b}
\end{equation*}
$$

and
$\operatorname{Cov}\left\{f_{1}+f_{2+3}, a\right\}=\frac{V\left\{f_{2+3}\right\}+\operatorname{Cov}\left\{f_{1}, f_{2+3}\right\}}{f_{1}}-\frac{f_{2+3}\left[V\left\{f_{1}\right\}+\operatorname{Cov}\left\{f_{1}, f_{2+3}\right\}\right]}{f_{1}^{2}} \quad(\mathrm{~S} 13 \mathrm{c})$

These expressions allow the approximate variances of the estimates $t$ and $P_{a}$ to be estimated.

## Supplementary Table S1: Locus-by-locus breakdown of the (uncorrected) number of ancestral

polymorphisms in 34 autosomal and $33 X$-linked genes in $D$. pseudoobscura and D. miranda.

| Locus | Type 1 (Shared) | D. pseudoobscura |  |  | D. miranda |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $n$ | Type 2/3 (Ancestral) | Type $4 / 5$ \& 6 <br> ('de novo') | $n$ | Type 2/3 (Ancestral) | Type $4 / 5 \& 6$ <br> ('de novo’) |
| Autosomal |  |  |  |  |  |  |  |
| ade3 | 0 | 16 | 2 | 7 | 15 | 0 | 4 |
| crinkled | 0 | 15 | 3 | 14 | 15 | 1 | 0 |
| GA10135 | 0 | 16 | 0 | 10 | 15 | 0 | 3 |
| GA10344 | 0 | 16 | 0 | 7 | 15 | 0 | 0 |
| GA12147 | 1 | 15 | 0 | 10 | 16 | 0 | 4 |
| GA12512 | 0 | 16 | 1 | 10 | 15 | 0 | 0 |
| GA12664 | 0 | 16 | 1 | 13 | 16 | 0 | 3 |
| GA12722 | 0 | 16 | 1 | 9 | 16 | 0 | 1 |
| GA13578 | 0 | 15 | 1 | 6 | 15 | 0 | 1 |
| GA14694 | 0 | 15 | 1 | 6 | 16 | 0 | 0 |
| GA14715 | 0 | 15 | 0 | 2 | 14 | 1 | 1 |
| GA15377 | 0 | 16 | 0 | 8 | 16 | 0 | 3 |
| GA16473 | 0 | 15 | 1 | 7 | 15 | 0 | 2 |
| GA17300 | 0 | 15 | 3 | 10 | 15 | 0 | 1 |
| GA17553 | 0 | 16 | 1 | 7 | 16 | 0 | 2 |
| GA17997 | 1 | 16 | 1 | 7 | 15 | 0 | 1 |
| GA17998 | 0 | 16 | 1 | 5 | 16 | 0 | 1 |
| GA18219 | 0 | 15 | 5 | 17 | 16 | 0 | 1 |
| GA18470 | 0 | 16 | 0 | 4 | 15 | 0 | 1 |
| GA18654 | 0 | 16 | 0 | 7 | 15 | 1 | 2 |
| GA19119 | 0 | 16 | 3 | 7 | 15 | 0 | 3 |
| GA19169 | 0 | 16 | 0 | 8 | 15 | 0 | 2 |
| GA19326 | 0 | 16 | 1 | 1 | 15 | 0 | 1 |
| GA19427 | 0 | 16 | 1 | 4 | 15 | 0 | 3 |
| GA19649 | 0 | 15 | 2 | 10 | 16 | 0 | 5 |
| GA19678 | 0 | 16 | 0 | 5 | 16 | 0 | 3 |
| GA20117 | 0 | 16 | 2 | 7 | 16 | 0 | 1 |
| GA20218 | 0 | 16 | 0 | 9 | 16 | 0 | 1 |
| GA20407 | 0 | 16 | 0 | 4 | 15 | 0 | 0 |
| GA22152 | 1 | 15 | 0 | 9 | 15 | 1 | 2 |
| GA25303 | 1 | 16 | 3 | 11 | 14 | 0 | 5 |
| GA25341 | 0 | 14 | 0 | 3 | 14 | 1 | 1 |
| GA25997 | 0 | 16 | 1 | 9 | 16 | 0 | 2 |
| lamin | 0 | 16 | 1 | 5 | 16 | 0 | 1 |
| Sub-total | 4 | 532 | 36 | 258 | 521 | 5 | 61 |
| X-linked |  |  |  |  |  |  |  |
| GA10596 | 0 | 16 | 0 | 8 | 14 | 0 | 3 |
| GA10819 | 1 | 16 | 1 | 11 | 13 | 0 | 0 |
| GA11209 | 1 | 16 | 2 | 10 | 16 | 0 | 1 |
| GA11389 | 0 | 16 | 3 | 6 | 16 | 1 | 6 |
| GA12234 | 0 | 16 | 0 | 7 | 16 | 0 | 1 |
| GA12817 | 0 | 16 | 1 | 6 | 16 | 0 | 2 |
| GA12834 | 0 | 15 | 3 | 5 | 14 | 1 | 1 |


| GA12844 | 0 | 15 | 3 | 16 | 14 | 0 | 0 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| GA13242 | 0 | 16 | 0 | 1 | 16 | 0 | 0 |
| GA13913 | 0 | 15 | 0 | 0 | 16 | 0 | 0 |
| GA14058 | 1 | 15 | 1 | 4 | 16 | 1 | 3 |
| GA14176 | 0 | 16 | 0 | 10 | 14 | 0 | 1 |
| GA14572 | 0 | 16 | 1 | 4 | 15 | 0 | 3 |
| GA14705 | 1 | 16 | 0 | 4 | 16 | 1 | 0 |
| GA15328 | 0 | 16 | 0 | 8 | 12 | 0 | 5 |
| GA15393 | 0 | 16 | 0 | 0 | 12 | 1 | 1 |
| GA15436 | 0 | 15 | 0 | 10 | 16 | 0 | 0 |
| GA15687 | 0 | 16 | 0 | 6 | 16 | 0 | 0 |
| GA16943 | 0 | 16 | 3 | 3 | 14 | 0 | 0 |
| GA17354 | 0 | 16 | 0 | 1 | 16 | 1 | 1 |
| GA17564 | 0 | 16 | 0 | 1 | 15 | 0 | 1 |
| GA17590 | 1 | 16 | 1 | 3 | 16 | 2 | 0 |
| GA17594 | 0 | 16 | 0 | 1 | 15 | 0 | 1 |
| GA17599 | 0 | 16 | 0 | 1 | 16 | 1 | 0 |
| GA17782 | 0 | 16 | 0 | 4 | 16 | 0 | 1 |
| GA18107 | 0 | 16 | 0 | 11 | 16 | 0 | 0 |
| GA19139 | 0 | 16 | 1 | 2 | 16 | 0 | 0 |
| GA19265 | 1 | 16 | 1 | 9 | 16 | 0 | 4 |
| GA19408 | 0 | 16 | 1 | 13 | 13 | 1 | 0 |
| GA20067 | 0 | 16 | 2 | 3 | 16 | 0 | 1 |
| GA20648 | 0 | 16 | 0 | 4 | 16 | 0 | 0 |
| GA20842 | 0 | 16 | 0 | 5 | 16 | 0 | 5 |
| GA21237 | 0 | 16 | 1 | 9 | 16 | 0 | 1 |
| GA21946 | 0 | 15 | 1 | 12 | 14 | 0 | 0 |
| Sub-total | 6 | 538 | 26 | 198 | 515 | 10 | 42 |
| TOTAL | 10 | 15.74* | 62 | 456 | 15.24* | 15 | 103 |

$\left(^{*}\right)$ denotes an average, rather than a total.

Supplementary Table S2: Counts of transitions and transversions across both species for $X$ linked (X) and autosomal (A) datasets. The ratio of transitions to transversions did not differ significantly between $X$ and $A\left(\chi^{2}{ }_{[1]}=0.246, p>0.05\right)$; thus, the datasets were combined to give empirical rates of mutation for transitions and transversions.

|  | D. pseudoobscura |  | D. miranda |  | TOTAL |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | X | A | X | A |  |
| Transitions | 170 | 221 | 46 | 59 | 496 |
| Transversions | 79 | 113 | 17 | 22 | 231 |
| TOTAL | 249 | 334 | 63 | 81 | 727 |

