# American Political Science Review Turnout and Amendment 4: Mobilizing Eligible Voters Close to Formerly Incarcerated Floridians --Manuscript Draft--

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Corresponding Author:	Kevin T Morris Brennan Center for Justice New York, New York UNITED STATES
Corresponding Author's Institution:	Brennan Center for Justice
Corresponding Author's Secondary Institution:	
Order of Authors:	Kevin T Morris, MUP
Order of Authors Secondary Information:	
Abstract:	Recent scholarship shows that eligible voters in neighborhoods home to many arrested and incarcerated individuals vote at lower rates than those in less-impacted neighborhoods. Little work, however, has interrogated how this turnout gap might be counteracted. This paper uses Amendment 4, a 2018 Florida ballot initiative that promised to re-enfranchise most individuals whose voting rights had been revoked due to a felony conviction, to investigate whether this turnout disparity can be narrowed by a ballot initiative of particular significance to communities most impacted by incarceration. Using prison release records, I identify the neighborhoods and households where formerly incarcerated individuals live and assess the voting history of their neighbors and housemates. I find no evidence that Amendment 4 increased these voters' turnout in 2018 relative to other voters. While ending felony disenfranchisement is necessary, closing the turnout gap resulting from histories of policing and incarceration will require greater investment and engagement.
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# Turnout and Amendment 4: Mobilizing Eligible Voters Close to Formerly Incarcerated Floridians\*

Kevin Morris $^{\dagger,\ddagger,\$}$ 

February 09, 2021

#### Abstract

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<sup>&</sup>lt;sup>†</sup>ORCID ID: 0000-0001-7725-6723

<sup>&</sup>lt;sup>‡</sup>Researcher, Brennan Center for Justice at NYU School of Law, 120 Broadway Ste 1750, New York, NY 10271 (kevin.morris@nyu.edu)

<sup>&</sup>lt;sup>§</sup>PhD Student, Sociology Program, CUNY Graduate Center

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#### <sup>19</sup> Introduction

On November 6<sup>th</sup>, 2018, Floridians voted to amend their state constitution to re-enfranchise 20 individuals with felony convictions in their past (Taylor 2018). The move was hailed as 21 transformative for Floridian — and American — democracy; Uggen, Larson, and Shannon 22 (2016) had estimated a few years earlier that some 1.5 million Floridians were disenfranchised 23 and had finished serving their sentences, making the amendment the largest expansion of 24 the franchise in the United States since the Twenty-sixth Amendment lowered the voting age 25 to 18. The amendment received broad support. Although it needed just 60 percent of the 26 vote to pass, 64.5 percent of voters supported the ballot initiative. This support contrasts 27 sharply with other statewide races: Ron DeSantis won the gubernatorial race with only 49.6 28 percent of the vote, while winning just 50.1 percent sent Rick Scott to the United States 29 Senate. 30

Prior to 2018, Floridians convicted of felony offenses were permanently disenfranchised unless 31 they applied for and received an individual pardon from the state's clemency board. This 32 was characterized by a "low success rate, cumbersome process, and lengthy amount of time" 33 (B. L. Miller and Spillane 2012b, 432) and was driven in part by gubernatorial discretion: 34 although Charlie Crist restored voting rights to roughly 150 thousand individuals over a 4 35 year period, Rick Scott did so for fewer than 3 thousand people over 8 years (Schlakman 36 2018). At the time Amendment 4 was passed, it was widely reported that the backlog of 37 applications was nearly 10,000 and the wait stretched for as long as a decade (Ramadan, 38 Stucka, and Washington 2018). Over the years, Florida's procedure was subject to numerous 39 lawsuits, and was ruled unconstitutional in early 2018 with Judge Mark Walker describing it 40 as "a gauntlet of constitutionally infirm hurdles."<sup>1</sup> Amendment 4 promised to automatically 41 restore voting rights once individuals had completed their sentence, though it did not apply 42 to individuals convicted of murder or sexual offenses. 43

<sup>&</sup>lt;sup>1</sup>Hand et al. v. Scott et al., 4:17cv128-MW/CAS (U.S. District Court for the Northern District of Florida 2018).

In recent years scholars have leveraged administrative records and sophisticated statistical 44 techniques to study the actual political effects of felony disenfranchisement in the United 45 States (e.g. Meredith and Morse 2014, 2015; Colgan 2019; Morris 2021b). With the notable 46 exception of White (2019a), however, the behavior of voters who live with individuals who 47 have been convicted of felony offenses — but have not themselves been convicted — has gone 48 unstudied. This article brings together these analytical approaches and an interdisciplinary 49 body of literature to understand the political behavior of citizens whose family members 50 have been incarcerated due to a felony conviction. 51

This study explores whether the opportunity to vote on Amendment 4 increased the (relative) 52 participation among eligible voters who lived with or near individuals disenfranchised due 53 to a period of felony incarceration. Americans' political knowledge is deeply shaped by the 54 incarceration of a loved one (Lee, Porter, and Comfort 2014), and exposure to the carceral 55 state chills political involvement even among individuals who are not convicted. The criminal 56 justice system can leave even would-be voters without a criminal record feeling as though 57 political involvement is not for "people like me," often despite having considerable political 58 knowledge (Lerman and Weaver 2014). A growing body of quantitative research captures 59 these "spillover" effects, demonstrating that neighborhoods with high levels of incarceration 60 and disenfranchisement vote at markedly lower rates than other similar neighborhoods (e.g. 61 Burch 2014; Morris 2020). 62

Amendment 4 in Florida offers a unique opportunity to investigate whether these chilling 63 effects can be overcome by one ballot initiative. As I explain in the section that follows, 64 Amendment 4 offered individuals living with or near formerly incarcerated individuals an 65 opportunity to redefine their relationship with the government in positive ways. Although 66 this made the ballot initiative perhaps particularly salient for these individuals, it took place 67 against the backdrop of an entrenched carceral state that negatively structured many facets 68 of their lives (see, for instance, Travis and Waul 2003). Ultimately, I do not find evidence 69 that Amendment 4 mobilized individuals living with or near formerly incarcerated Floridians 70

<sup>71</sup> in 2018 above-and-beyond turnout increases observed in other, similar voters.

#### <sup>72</sup> Theory and Literature

In recent years scholars have documented the effect of the American criminal legal system 73 on the lives of those who come under its purview, even once they are no longer under 74 formal supervision. The growth of the criminal legal system has resulted in what Monica 75 Bell calls legal estrangement, which reflects both legal cynicism — a cultural orientation 76 that views the law and its enforcers as "illegitimate, unresponsive, and ill equipped to ensure 77 public safety" (Kirk and Papachristos 2011, 1191; see also Sampson and Bartusch 1998; Kirk 78 and Matsuda 2011; Morenoff and Harding 2014) — and the objective structural conditions 79 (such as policing practices and criminal law) that give rise to this orientation (Bell 2017, 80 2066 – 2067). Legal estrangement has also been linked with "institutional" or "system 81 avoidance." Brayne (2014, 385), for instance, documents that "individuals who have been 82 stopped, arrested, convicted, or incarcerated are less likely to interact with institutions 83 that keep formal records, such as hospitals, banks, employment, and schools." Haskins 84 and Jacobsen (2017) finds that institutional avoidance explains formerly incarcerated men's 85 reduced willingness to be involved with their children's schools, and Remster and Kramer 86 (2018) shows that this avoidance explains the behavior of Black and non-Black individuals 87 alike. 88

Institutional avoidance is especially clear when it comes to democratic participation, particularly in the voting booth. It is well established that a criminal conviction — and, more specifically, a period of incarceration — decreases turnout even when individuals are no longer legally disenfranchised (Weaver and Lerman 2010; Burch 2011; White 2019b; but see Gerber et al. 2017). The effect of disenfranchisement policy on the political behavior of individuals who experience the criminal justice system indirectly via the conviction of a family or community member, however, is somewhat mixed. Most research finds that turnout

is measurably lower in states with stricter voter disenfranchisement policies or more disen-96 franchised citizens (e.g. Bowers and Preuhs 2009; King and Erickson 2016), though Miles 97 (2004) argues that these effects are small. The little research that has explored the spillover 98 effects of disenfranchisement policy at the *neighborhood* level has similarly found evidence 99 that incarceration and disenfranchisement demobilizes eligible voters in impacted communi-100 ties (Burch 2014; Morris 2020; but see White 2019a). Understanding whether Amendment 101 4 was likely to recoup the lost turnout of eligible voters who lived with or near the disen-102 franchised requires understanding how their indirect exposure to the criminal justice system 103 (or "proximal contact" (Walker 2014)) depressed turnout to begin with. 104

Work from Vesla Weaver and Amy Lerman (2010; 2014) describes in great detail how legal 105 estrangement ruptures individuals' willingness to engage in electoral politics. They argue 106 that a felony conviction serves as "a durable constraint and marker of their citizenship" 107 (Lerman and Weaver 2014, 133), and that custodial citizens — individuals in communities 108 with aggressive crime control who may or may not have a criminal history themselves — 109 "become less likely to believe that they (and those like them) can change the system, a 110 reduction in external efficacy" (Lerman and Weaver 2014, 137, emphasis in the original). 111 Their work is replete with examples of individuals who know much about politics yet choose 112 to "stay below the radar" because "'they're [government officials] not interested in what I 113 have to say'" (Lerman and Weaver 2014, 210). 114

Importantly, these demobilizing consequences are not limited to those who are convicted; 115 rather, "the sense of alienation in a carceral regime emanates not only from what police 116 might do to 'you,' but from what they might do to your friends, your intimate partners, 117 your parents, your children; to people of your race or social class; and to people who live 118 in the neighborhood or the city where you live" (Bell 2017, 2058). Put differently, the legal 119 system serves as a site of political socialization even for those who are not formally convicted 120 of a crime (Lee, Porter, and Comfort 2014; Comfort 2016; Kirk 2016). There is, however, 121 some evidence that these chilling effects on political participation can be overcome. Recent 122

work demonstrates that direct and indirect contact with the criminal justice system can
be mobilizing when these experiences are linked with narratives of injustice (Walker and
García-Castañon 2017; Walker 2020).

Of course, there is no bright line dividing individuals with *indirect* exposure to the criminal 126 justice system from individuals with their own, *direct* exposure to the carceral state. The 127 geographic concentration of policing and incarceration patterns (e.g. Gelman, Fagan, and 128 Kiss 2007) mean that individuals in community with the formerly incarcerated — that is, 129 people living with or near formerly incarcerated residents — might also have other, direct 130 relationships with the criminal justice system. In 2017 there were 711,831 arrests in Florida 131 but just 134,554 guilty felonious dispositions.<sup>2</sup> Although individuals who were arrested but 132 not convicted of felonies were not legally disenfranchised, even low-level interactions can 133 have a chilling effect on one's relationship with the government, a relationship Amendment 134 4 could have led them to reconsider. 135

Based on these literatures, I hypothesized that both the substance of the proposed consti-136 tutional amendment and the messaging used by the campaign supporting its passage would 137 increase the relative turnout of individuals living with and near formerly incarcerated Florid-138 ians. Restoring voting rights to individuals who had been convicted of felony offenses would 139 end the "civil death" of felony disenfranchisement (Ewald 2002; B. L. Miller and Spillane 140 2012a), nullifying one of the durable badges identified by Lerman and Weaver. Amendment 141 4 offered those in community with the formerly incarcerated the chance to affirm that their 142 family and community members deserved to have their voices heard in the democratic arena, 143 a chance I anticipated would disproportionately spur them to participate. 144

<sup>145</sup> Moreover, the public messaging employed by the Amendment 4 campaign was explicitly <sup>146</sup> designed to change how voters understood the citizenship of disenfranchised individuals. <sup>147</sup> The campaign cast the ballot initiative as an issue of fairness, criticizing Florida's existing <sup>148</sup> disenfranchisement policy for creating two tiers of citizenship. The organization leading the

<sup>&</sup>lt;sup>2</sup>See http://edr.state.fl.us/Content/resource-demand/criminal-justice/reports/criminal-justice/cj7.pdf.

campaign leveraged the notion that disenfranchised citizens deserved to be re-incorporated 149 into the body politic in its very name — "Second Chances Florida." The framing was 150 effective: the editorial boards of each of Florida's three biggest newspapers endorsed the 151 amendment, all using language related to fairness and civic redemption. The Tampa Bay 152 Times told readers they had a "remarkable opportunity to remedy that unfairness" (Tampa 153 Bay Times 2018); the Sun Sentinel informed voters "[t]here may never be an opportunity 154 to do a better thing than to vote yes on this reform" (Sun Sentinel 2018); and the Orlando 155 Sentinel said that Florida's then-policy "denie[d] our fellow citizens a second chance. It 156 denie[d] redemption" (Orlando Sentinel 2018). Insofar as the campaign was successful at 157 helping these individuals understand the experiences of their formerly incarcerated family 158 and community members in the context of a broader narrative of (racial) injustice, I expected 159 this framing would mobilize them to vote at higher rates than other voters. 160

In addition to newspapers across the state, the campaign deployed "volunteers from a broad 161 coalition that included advocacy groups, Christian organizations, the League of Women Vot-162 ers, criminal justice experts and, of course, those who had been convicted of felonies" (Robles 163 2018). Andrew Gillum, the Democratic gubernatorial candidate, also vocally supported the 164 amendment, openly discussing his family's relationship with the criminal justice system and 165 his own sibling's disenfranchisement (Smith 2018). Voters were thus getting cues from all 166 sorts of messengers that Amendment 4 deserved to be passed, and that individuals with 167 convictions in their past should be allowed to vote. I expected that these cues, plus the 168 descriptive representation (Merolla, Sellers, and Fowler 2013) promised by Gillum, would 169 have proved especially mobilizing for the population in closest contact with disenfranchised 170 Floridians. 171

At the same time, there was some reason to think the ballot initiative would not disproportionately increase turnout among voters in close contact with formerly incarcerated, disenfranchised individuals. Legal estrangement runs deep: the "hidden curriculum" of the criminal justice system (Justice and Meares 2014; Meares 2017) teaches individuals their place in this system over a very long period, through both incarceration and day-to-day interactions with government representatives such as the police. It is perhaps naive to expect
that a single ballot initiative could overcome these negative forces.

Moreover, the individuals in these neighborhoods were perhaps less familiar with the content 179 of Amendment 4 than others: Bowler and Donovan (1994), for instance, demonstrates that 180 education and polarization are strong predictors of individuals' familiarity with ballot ini-181 tiatives. Shaker (2012) also finds that higher-educated individuals are more knowledgeable 182 about local politics. Given that formerly incarcerated individuals leave prison for neigh-183 borhoods with less access to higher education (see Table 2 below), their neighbors and 184 housemates may have been less aware of the amendment in the first place, in which case it 185 obviously could not heighten motivation to cast a ballot. 186

#### <sup>187</sup> Research Design and Expectations

I begin by testing whether a neighborhood's formerly incarcerated population influenced its turnout in 2018. Because statewide felony probation records are not available, this analysis is based on only the subset of disenfranchised individuals who were imprisoned for a felony conviction. Neighborhoods that are home to formerly incarcerated individuals are identified by geocoding release records from the Florida Department of Corrections, and I offer two definitions of neighborhoods.

Neighborhoods are first defined as voting precincts. The Florida Division of Elections makes election results available at this level, which allows me to test turnout specifically on Amendment 4 and neighborhood-level support for the amendment. I can also assess how salient the amendment was for participants by estimating the share of voters who "rolled off" (or chose not to vote) for Amendment 4. Unfortunately, the use of precinct-level data leaves us with a major drawback: when doing analysis at this level, bias-free turnout denominators are hard to come by. Because the Census Bureau does not produce population estimates for individual voting precincts, turnout cannot be calculated by dividing the number of ballots cast by the eligible population (that is, citizens over the age of 17 without a felony conviction); rather, it must be constructed as a share of registered voters. If there is a relationship between the number of formerly incarcerated residents and the registration rate of a neighborhood, our estimates will be biased.

That could be the case in the study at hand. Some political organizers supporting Amendment 4 focused on canvassing neighborhoods with many formerly incarcerated individuals (Speri 2018), potentially raising the registration rate in these areas. If relatively few of these newly-registered individuals voted, the net effect would be higher turnout among *eligible residents* but lower turnout among *registered voters*. For further discussion of how improper denominators can bias turnout estimates, see Amos, McDonald, and Watkins (2017) and Amos and McDonald (2020).

To address this potential problem, I also define neighborhoods as Census block groups. The 213 Census Bureau makes estimates of the citizen voting-age population (a better denominator 214 for turnout) available at this level. In this case, however, I must use a geocoded voter file 215 to determine turnout. Because I aggregate the number of participants in a block group 216 from individual-level data, I cannot determine whether an individual actually participated 217 in the contest for Amendment 4 or they rolled off. Similarly, I am unable to interrogate the 218 relationship between block group characteristics and support for Amendment 4. Although 219 each definition of neighborhood presents some drawbacks, the two definitions together paint 220 a full picture. 221

After examining whether the presence of formerly incarcerated residents was related with neighborhoods' voting behavior, I ask whether voters who lived with formerly incarcerated individuals turned out at higher rates in 2018. For this analysis, I use the release addresses of formerly incarcerated individuals (the most recent address available, according to the Department of Corrections) and voter file data to identify registered voters who lived with formerly incarcerated individuals. Voters are considered "treated" if they lived with a formerly incarcerated individual, and "untreated" otherwise. I then use a variety of individualand neighborhood-level characteristics to match treated and untreated voters using what methodologists call a "genetic" process (Sekhon 2011).

After matching these voters, I employ a difference-in-differences specification to determine whether treated voters participated at higher rates in the 2018 election. These analyses are run for all voters who lived with a formerly incarcerated individual, as well as only the subset of households whose members have not been to prison for many years. This final specification allows me to disentangle the depressive effect of indirect exposure to the criminal justice system from the mobilizing effect of Amendment 4 in 2018 by incorporating any depressive effect into the pre-2018 baseline.

<sup>238</sup> Table 1 summarizes the specific hypotheses this article tests.

	Hypothesis	Approach
Neigh	borhood Level	
1a.	Each additional formerly incarcerated resident in a voting precinct	OLS regression
	is associated with increased turnout among registered voters in that	
	precinct.	
1b.	Each additional formerly incarce rated resident in a Census block	OLS regression
	group is associated with increased turnout among eligible citizens in	
	that block group.	
2.	Each additional formerly incarcerated resident in a voting precinct is	OLS regression
	associated with increased support for Amendment 4 in that precinct.	
3.	Each additional formerly incarce rated resident in a voting precinct	OLS regression
	is associated with decreased roll-off in that precinct.	
House	ehold Level	
4.	Amendment 4 increased turnout in 2018 among household members	Difference-in-differences
	of formerly incarce rated individuals relative to their controls. This	comparing turnout of voters in
	treatment effect was especially large among households whose	treated households to voters in
	members have not been to prison for many years.	untreated households.

Table 1: Hypotheses

# 239 Data

I leverage multiple data sources to investigate whether individuals in community with formerly incarcerated Floridians were more likely to vote in the 2018 election. Replication materials can be found in the *APSR Dataverse* (Morris 2021a). Although this study relies on voter file data and publicly-available prison release records, I anonymize the neighborhoods and households home to formerly incarcerated individuals in order to protect privacy.

#### 245 Department of Corrections Data

Felony incarceration records come from the Florida Department of Corrections' Offender 246 Based Information System (OBIS). The OBIS includes all individuals released from prison 247 following a felony conviction since October 1, 1997. There were approximately 390,000 such 248 individuals. I retain only the record associated with an individual's most recent incarceration 240 according to the release date, and identify all formerly incarcerated individuals who were 250 finished with their sentence as of the 2018 election by cross-referencing these records against 251 imprisonment and parole records. Roughly 38,000 individuals were either re-incarcerated or 252 on parole as of the 2018 election and are thus removed. The 10,000 or so individuals who 253 died or absconded before their sentence was completed are also removed from the dataset, 254 leaving us with about 343,000 individuals who had finished their sentence by the time of the 255 2018 midterm election. 256

The OBIS provides the "release plan address" for individuals who were formerly incarcerated. 257 As noted above, this is the most recent address available for individuals who are no longer 258 under supervision.<sup>3</sup> The address data are messy and require substantial cleaning. In some 259 cases, the address field is left blank; in others, the record simply notes the road or the town 260 of the individual's residence, without providing full address information. I assume that any 261 record that does not begin with an integer does not have a full address and cannot be used 262 (this results in the exclusion of just under 3 percent of records). The remaining addresses 263 are geocoded. Individuals whose addresses were geocoded outside of Florida (10.9 percent) 264 or for whom the geocoder failed (3.2 percent) are dropped. After completing the geocoding 265 process we are left with some 286,000 individuals who were finished with their sentence as of 266 the 2018 midterm, were released to Florida addresses, and reported an address that could be 267 geocoded. In other words, at least 94 percent of individuals released to addresses in Florida 268

<sup>&</sup>lt;sup>3</sup>The OBIS lists current addresses for individuals currently under community supervision, which may differ from the release plan addresses. However, according to a response to a public records request filed by the author with the Department of Corrections, these historical data are not maintained once an individual has been discharged.

<sup>269</sup> were successfully geocoded.

The successfully geocoded, formerly incarcerated individuals are then mapped to their home Census block groups using shapefiles from the Census Bureau, and to their home voter precincts using shapefile data collected by Kelso and Migurski (2018).

#### 273 Caveats with the DOC Data

Using the release plan address for individuals last released from prison many years ago presents some potential problems. Some of these individuals surely died or moved after completing their sentence. In the Supplementary Information I show the results presented in the body of this article when I limit the pool of formerly incarcerated people to individuals released from prison during or after 2015. Because these individuals were released more recently, their addresses are probably more accurate. The primary findings of this study hold when the sample is thus limited.

Many formerly incarcerated individuals leave prison not for homes with family members, but 281 rather to homeless shelters or other sites of incarceration. Of the five most commonly listed 282 addresses, three were Immigration and Customs Enforcement properties, one was owned by 283 the Salvation Army, and one was a rescue mission. The body of this article excludes formerly 284 incarcerated individuals whose address was listed by five or more individuals, as institutions 285 for returning citizens may have uniquely structured responses to Amendment 4 (see, for 286 instance, Henig 1994). The Supplementary Information shows that the primary findings in 287 the article hold when I include all formerly incarcerated individuals. Just over 15 percent of 288 formerly incarcerated individuals listed these sorts of addresses as their post-incarceration 289 residence. 290

Neither the OBIS nor any other statewide database makes records available for individuals
 sentenced to felony probation. Between 75 and 80 percent of individuals found guilty of

felonies in recent years in Florida have been sentenced to probation.<sup>4</sup> This may pose a problem: neighborhoods with residents disenfranchised due to felony probation are also "treated," as are housemates of these individuals. However, not all individuals who serve a term of felony probation actually lose their voting rights. Florida judges are allowed to "withhold adjudication" (Tragos and Sartes 2008), meaning defendants are not formally convicted of a felony, but consent to pay fines and restitution and to serve a term of probation. Individuals whose adjudication is withheld are not disenfranchised.

As discussed in the Supplementary Information, probation records with residential addresses 300 are available for Hillsborough County, the Florida county with the third-highest number 301 of formerly incarcerated individuals according to the OBIS records. Within Hillsborough 302 County, the correlation coefficient between the number of felony probationers and formerly 303 incarcerated residents (scaled by population) is 0.92 at the block group level. The evidence 304 from Hillsborough County therefore indicates that number of formerly incarcerated individ-305 uals in a neighborhood should be a reasonable proxy for the total number of disenfranchised 306 residents. 307

In the Supplementary Information, the neighborhood- and individual-level models presented in the body of this article are re-estimated using only neighborhoods and individuals in Hillsborough County, with individuals sentenced both to felony incarceration *and* probation included in the models. Their incorporation does not meaningfully impact the primary results. Although this study relies only on formerly incarcerated individuals, the data available for robustness checks indicate that the relationships detailed here probably extend to the full disenfranchised population.

<sup>&</sup>lt;sup>4</sup>See http://edr.state.fl.us/Content/resource-demand/criminal-justice/reports/criminal-justice/index. cfm.

#### <sup>315</sup> Voter File Data and Census Data

I primarily use Florida voter file data from the data vendor L2 Political which includes publicly-available information on individuals such as their home address, their age and gender, their participation history, and their political affiliation. In addition to the L2 data I use self-identified race and ethnicity information from the raw Florida voter file. I also use the raw Florida file to provide the gender for voters for whom L2 did not have data, as well as voters' home counties and precincts.

Precinct and block group demographics are constructed by aggregating up from the voter file data. Neighborhood characteristics such as average age are the averages of all registered voters in that neighborhood. For characteristics such as income that are unavailable at the individual level, voters are assigned the value associated with their home block group from the American Community Survey's 2014 – 2018 5-year estimates; the precinct average income, therefore, is effectively the average of all the block groups within that precinct, weighted by the number of registered voters.

#### 329 Matched Department of Corrections and Voter File Data

I identify registered voters who lived with formerly incarcerated individuals by matching on 330 residential addresses. As discussed above, these addresses are often in different formats. To 331 increase the quality of the matches, I standardize common street and address abbreviations 332 as well as capitalization. "Boulevard," for instance, becomes "BLVD" in each instance in the 333 DOC and voter file data. These standardizations are taken from Appendix C of the USPS 334 Postal Addressing Standards (2015). Exact matching for the entire residential address is 335 required. Formerly incarcerated individuals who were registered to vote are removed (as 336 noted in the Introduction, some individuals were able to have their voting rights restored). 337

#### 338 Potential Confounders

Voters with indirect exposure to the criminal justice system might have been uniquely mo-339 tivated to turn out through avenues other than the ballot initiative. For instance, Andrew 340 Gillum was poised to become the state's first Black governor, which could increase the 341 turnout of Black voters who are over-represented in the treatment group (e.g. Washing-342 ton 2006; Fairdosi and Rogowski 2015; P. Miller and Chaturvedi 2018). By controlling for 343 neighborhood demographics (and, in the matching exercise, forcing control voters to mirror 344 treated voters on key demographics such as race and party affiliation), I minimize the dif-345 ferences between the treatment and control groups along characteristics known to influence 346 turnout. 347

There is little reason to believe that changes to electoral rules would have differently influ-348 enced the turnout for individuals in close proximity to the formerly incarcerated than other, 349 similar voters. The number of early voting days was cut for the 2012 general election, but 350 the longer period was restored for the 2014 - 2018 period.<sup>5</sup> Early voting was not allowed on 351 college campuses in the 2014 and 2016 elections, though it was allowed in 2018 (Bousquet 352 2018). If voters who lived near the formerly incarcerated had better or worse access to college 353 campuses than other voters, this could influence their turnout. I include neighborhood-level 354 estimates of collegiate education in each of the regressions to mitigate the potential effects of 355 this change. Florida did not enact other reforms such as same-day registration or automatic 356 voter registration over the period, nor did its absentee voting rules change. We can therefore 357 be confident that any turnout effects observed are not being driven by the treatment group 358 responding to rules changes in different ways than other voters. 359

<sup>&</sup>lt;sup>5</sup>See https://ballotpedia.org/Voting\_in\_Florida.

#### <sup>360</sup> Neighborhood-Level Results

Before presenting the results of the econometric modeling, I examine whether — and to 361 what extent — block groups with formerly incarcerated individuals differ from block groups 362 elsewhere in the state. A simple comparison of block groups with and without formerly 363 incarcerated individuals, however, proves unhelpful: 97.1 percent of block groups in the 364 state are home to someone who has been to prison, though formerly incarcerated individuals 365 are clearly concentrated in some block groups. Column 1 of Table 2 presents the statewide 366 mean of all block groups, weighted by their population. In Column 2, I re-weight the block 367 groups by the number of formerly incarcerated residents. 368

Measure	Average Block Group All Floridians	Average Block Group Formerly Inc. Floridians
Median Income <sup>*</sup>	\$59,988	\$45,484
$Median Age^*$	42.5	39.9
% Unemployed*	6.4%	8.9%
% with Some College*	73.0%	65.2%
% Non-Hispanic White*	54.4%	44.5%
%Non-Hispanic Black*	15.4%	30.5%
% Latino*	25.2%	20.7%
Count	20,590,223	279,324

 Table 2:
 Block Group Demographics

\* Difference is significant at 95 percent confidence level.

Although nearly all parts of the state are impacted by the criminal justice system (and, more specifically, mass incarceration), Table 2 makes clear that formerly incarcerated individuals are concentrated in neighborhoods with lower incomes, higher levels of unemployment, and where a much larger share of the population is Black.

<sup>373</sup> I next assess whether the presence of formerly incarcerated residents was associated with

higher turnout in 2018 using ordinary least squares regressions. In the precinct-level model, 374 turnout is calculated by dividing the number of ballots cast for or against Amendment 4 375 by the number of actively registered voters in the precinct.<sup>6</sup> while block group turnout is 376 calculated by dividing the number of voters marked as participants in the voter file by the 377 adjusted citizen voting age population (ACVAP).<sup>7</sup> Formerly Incarcerated Residents is the 378 primary independent variable. Models 2 and 4 also include a measure of how long the 379 average formerly incarcerated resident has been out of prison (Av. Years since Most Recent 380 *Incarceration*) to test whether recently incarcerated residents impact turnout differently 381 than those who were released many years ago. Neighborhoods with no formerly incarcerated 382 residents are excluded from models 2 and 4. I also control for other covariates known to 383 influence turnout such as age and income. There is just one observation per neighborhood 384 in each model, but I control for neighborhood-level turnout from the 2010 - 2016 general 385 elections. Finally, I include fixed effects for congressional districts, and robust standard 386 errors are clustered at this level.<sup>8</sup> 387

 $<sup>^{6}</sup>$ The 35 precincts where calculated turnout exceeds 100 percent have been dropped from the analysis, though their inclusion does not affect the results.

<sup>&</sup>lt;sup>7</sup>I define ACVAP by subtracting the number of all formerly incarcerated individuals from the Census Bureau's estimated citizen voting age population (including the individuals who are excluded from the primary independent variable count because they returned to common post-release residences). My definition of ACVAP is similar to the voting eligible population estimated by McDonald (2002), though I do not have estimates of the number of individuals disenfranchised for a felony probation at the neighborhood-level.

<sup>&</sup>lt;sup>8</sup>Where neighborhoods cross congressional district boundaries they are assigned to the district in which most of their voters live.

	Precinc	t-Level	Block Group-Level	
	(1)	(2)	(3)	(4)
Formerly Incarcerated Residents	-0.0002***	-0.0002***	-0.0002***	-0.0002***
	(0.00004)	(0.00003)	(0.00004)	(0.00004)
Av. Years since Most Recent Incarceration		0.0001		0.0002*
		(0.001)		(0.0001)
Percent White	0.017	-0.088	0.017	0.017
	(0.110)	(0.123)	(0.014)	(0.014)
Percent Black	0.027	-0.086	0.041**	0.040**
	(0.109)	(0.121)	(0.017)	(0.017)
Percent Latino	-0.081	-0.175	-0.007	-0.008
	(0.116)	(0.125)	(0.016)	(0.016)
Percent Asian	0.082	-0.006	0.040*	0.039*
	(0.128)	(0.166)	(0.022)	(0.022)
Percent Male	0.302	0.376**	0.095	0.102
	(0.188)	(0.179)	(0.086)	(0.089)
Percent Democrats	0.059	0.161**	0.067***	0.067***
	(0.082)	(0.073)	(0.020)	(0.020)
Percent Republicans	0.015	0.105	0.007	0.004
1	(0.081)	(0.070)	(0.024)	(0.024)
Average Age	0.0001	0.0001	0.001***	0.001***
	(0.0005)	(0.001)	(0.0003)	(0.0003)
Average Income (\$10,000s)	0.002**	0.001**	0.002***	0.002***
	(0.001)	(0.001)	(0.0003)	(0.0003)
Percent With Some College	0.183***	0.188***	0.082***	0.082***
0	(0.016)	(0.020)	(0.005)	(0.005)
Percent Unemployed	-0.032	-0.033	-0.005	-0.004
I J I	(0.025)	(0.028)	(0.006)	(0.006)
Constant	$-0.211^{*}$	$-0.235^{*}$	$-0.188^{**}$	-0.200**
	(0.114)	(0.127)	(0.083)	(0.087)
Congressional District FEs	X	X	X	X
Turnout in 2010 – 2016	Х	Х	Х	Х
Observations $\mathbf{D}^2$	5,797	5,477	10,817	10,550
$\pi$ Adjusted $B^2$	0.782 0.781	0.814	0.979 0.979	0.979 0.979

### Table 3: Neighborhood Turnout in 2018

 $^{***}p<0.01,\,^{**}p<0.05,\,^*p<0.1.$  Robust standard errors (clustered by congressional district) in parentheses.

Table 3 indicates that 2018 turnout was lower in neighborhoods with more formerly incarcerated residents, and the average length of time since formerly incarcerated residents' most recent incarceration is not related to turnout. The block group models have nearly twice as many observations as the precinct-level ones and their  $R^2$ s are considerably higher, perhaps indicating a better fit. Nevertheless, the estimated coefficient for *Formerly Incarcerated Residents* is the same (when rounded to one hundredth of a percentage point) for both neighborhood definitions.

The primary coefficients in Table 3 are small and perhaps difficult to interpret without 395 context. Figure 1 shows the marginal effect of each additional formerly incarcerated resident 396 on precinct-level turnout for Amendment 4 from model 1. All other covariates are held at 397 their means. Although the number of formerly incarcerated residents in a precinct reaches 398 a maximum of 594, there are 300 or fewer such residents in 99.2 percent of precincts, and I 399 limit the figures to this range. Predicted turnout in precincts with zero formerly incarcerated 400 residents is just over 66 percent; in precincts with 300 such residents, predicted turnout was 401 below 61 percent, implying a five-point decrease over the effective range of observed values. 402



Figure 1: Marginal Effect of Formerly Incarcerated Residents on Precinct Turnout Among Registered Voters

In Table 4 I present the results of OLS models that test whether the number of formerly incarcerated community members influenced a neighborhood's support for Amendment 4 or Amendment 4 roll-off. Roll-off is calculated as  $1 - \frac{Ballots Cast for Amendment 4}{Ballots Cast in Contest with the Most Votes}$ . It ranges from zero (if everyone who cast a ballot made a decision on the Amendment 4 question) to one (if no participants voted for or against Amendment 4). A lower number represents lower roll-off, indicating that the issue was more salient for participants.

	Support	for Am. 4	Roll-Off	
	(1)	(2)	(3)	(4)
Formerly Incarcerated Residents	0.0001**	0.0001**	-0.00004***	-0.00004***
	(0.00003)	(0.00003)	(0.00001)	(0.00001)
Av. Years since Most Recent Incarceration		0.002**		0.0004**
		(0.001)		(0.0002)
Percent White	0.069	-0.051	$-0.071^{*}$	$-0.076^{*}$
	(0.122)	(0.093)	(0.042)	(0.046)
Percent Black	0.188*	0.026	-0.042	-0.048
	(0.107)	(0.084)	(0.040)	(0.042)
Percent Latino	0.049	-0.101	-0.050	-0.052
	(0.114)	(0.092)	(0.043)	(0.045)
Percent Asian	0.244	0.133	$-0.101^{*}$	$-0.117^{*}$
	(0.177)	(0.170)	(0.052)	(0.061)
Percent Male	$-0.383^{**}$	$-0.299^{*}$	$-0.204^{*}$	$-0.193^{*}$
	(0.185)	(0.170)	(0.113)	(0.117)
Percent Democrats	0.192	0.197	0.031	0.024
	(0.143)	(0.191)	(0.021)	(0.029)
Percent Republicans	$-0.396^{***}$	$-0.429^{***}$	$0.039^{*}$	0.037
	(0.120)	(0.151)	(0.020)	(0.027)
Average Age	-0.0003	0.00005	0.001***	0.001***
	(0.0004)	(0.0004)	(0.0002)	(0.0002)
Average Income (\$10,000s)	$-0.003^{***}$	$-0.002^{**}$	-0.00003	-0.00004
	(0.001)	(0.001)	(0.0002)	(0.0002)
Percent With Some College	$0.155^{***}$	$0.158^{***}$	$-0.029^{***}$	$-0.032^{***}$
	(0.034)	(0.029)	(0.006)	(0.008)
Percent Unemployed	-0.015	-0.024	$-0.019^{*}$	-0.011
	(0.018)	(0.021)	(0.011)	(0.010)
Constant	1.023***	1.055***	0.220**	0.212**
	(0.165)	(0.197)	(0.095)	(0.105)
Congressional District FEs	Х	Х	Х	X
Turnout in 2010 – 2016	Х	Х	Х	Х
Observations $\mathbf{D}^2$	5,797	5,477	5,797	5,477
$R^-$ Adjusted $R^2$	0.788 0.787	0.869	0.315 0.309	$\begin{array}{c} 0.385\\ 0.380\end{array}$
	0.101	0.000	0.000	0.000

### Table 4: Precinct Engagement with Amendment 4

 $^{***}p<0.01,\,^{**}p<0.05,\,^*p<0.1.$  Robust standard errors (clustered by congressional district) in parentheses.

Table 4 demonstrates that precincts with more formerly incarcerated residents supported Amendment 4 at slightly higher rates. Similarly, roll-off was lower in neighborhoods with more formerly incarcerated residents. Figures 2 and 3 plot the marginal effect of each additional formerly incarcerated resident on a precinct's support for Amendment 4 (model 1), and the precinct's roll-off on Amendment 4 (model 3). These figures make clear that the number of formerly incarcerated residents has a relatively small impact on precinct support for its passage, and a relatively large impact on precinct level roll-off.



Figure 2: Marginal Effect of Formerly Incarcerated Residents on Support for Amendment 4



Figure 3: Marginal Effect of Formerly Incarcerated Residents on Amendment 4 Roll-Off

Why the relationship between formerly incarcerated residents and support is less strong 416 (though positive and statistically significant) than salience is not clear, perhaps pointing to 417 a variety of individual responses to crime and criminal justice policy in these neighborhoods. 418 Leverentz (2011) argues that punitiveness is positively correlated with the salience of crime. 419 The recently incarcerated residents might activate both punitiveness and support for the 420 amendment, with support winning out slightly. The coefficients for Av. Years since Most 421 *Recent Incarceration* indicate that neighborhoods where the formerly incarcerated residents 422 have been out of prison for longer saw both higher support for Amendment 4 and higher 423 roll-off. Future work ought to interrogate how support for criminal justice reforms and the 424 salience of those reforms change as community members' incarcerations recede into the past. 425 These neighborhood-level models demonstrate that neighborhoods with many formerly in-426 carcerated residents did not turn out at higher rates than other, similar neighborhoods in 427 2018 even though Amendment 4 was on the ballot. However, while formerly incarcerated 428 neighbors were not associated with getting people into the voting booth, they were associated 429

with how voters cast their ballots once there.

#### 431 Individual-Level Results

Neighborhood turnout rates could be obscuring underlying patterns. Inducements to vote 432 at the household level might be too small to register at the neighborhood level, and it is 433 possible that Amendment 4 shaped turnout differently for individuals who live with formerly 434 incarcerated individuals than for their neighbors. A neighborhood may have disengaged from 435 the political process thanks to exposure to the carceral state. Household members of the 436 formerly incarcerated may have had a similar historical response, and yet be more susceptible 437 to mobilization from Amendment 4; they are, after all, the voters whose identities are most 438 likely shaped by indirect exposure to felony disenfranchisement. 439

This section directly examines the turnout of individuals who lived with formerly incarcerated individuals in 2018, relative to other, similar voters. As discussed above, I identify individuals who live with formerly incarcerated individuals by matching addresses listed in the Department of Corrections release data to the registered voter file. All registered voters who live at an address reported by a formerly incarcerated individual are considered "treated."

Each treated individual is then matched (Sekhon 2011) with five untreated registered voters elsewhere in her congressional district.<sup>9</sup> I use five matches in order to increase the sample size of the study; the large pool of potential controls means this can be done without sacrificing the quality of the matches. Voters' block group median income and share with some collegiate education come from the ACS 2018 5-year estimates, while all other characteristics come from the voter file. Matching is done with replacement and ties are randomly broken. Table 5 presents the results of the matching exercise for each of the characteristics used.

<sup>&</sup>lt;sup>9</sup>Due to computing constraints, a random 5 percent random sample stratified by treatment status is used to calculate the genetic weights. The full sample is used for matching.

	Means: Uni	matched Data	Means: Ma	Means: Matched Data		s: Matched Data Perce		Percent In	nprovement	
	Treated	Control	Treated	Control	Mean Diff	eQQ Med	eQQ Mean	eQQ Max		
%White	41.5%	63.2%	41.5%	41.5%	100.00	100.00	100.00	100.00		
% Black	38.8%	12.7%	38.8%	38.8%	100.00	100.00	100.00	100.00		
% Latino	12.8%	16.9%	12.8%	12.8%	100.00	100.00	100.00	100.00		
% Asian	0.8%	2.0%	0.8%	0.8%	100.00	100.00	100.00	100.00		
% Female	55.2%	52.4%	55.2%	55.2%	100.00	100.00	100.00	100.00		
% Male	41.5%	45.0%	41.5%	41.5%	99.99	99.99	99.99	99.99		
Registration Date	2004-01-28	2004-09-24	2004-01-28	2004-02-11	94.03	38.85	27.88	19.19		
Age	48.95	52.45	48.95	48.77	94.71	94.34	92.44	90.89		
% Democrat	53.7%	36.9%	53.7%	53.7%	99.99	99.99	99.99	99.99		
% Republican	21.0%	35.4%	21.0%	21.0%	100.00	100.00	100.00	100.00		
% with Some College	66.5%	75.3%	66.5%	66.5%	99.92	99.95	99.92	99.62		
Median Income	\$47,389	\$62,995	\$47,389	\$47,402	99.92	99.82	99.70	99.22		

Table 5:Balance Table

As Table 5 makes clear, the treated registered voters differ in meaningful ways from the rest of the electorate: three times as many are Black, a larger share are registered Democrats, and they live in neighborhoods with lower incomes. The matching process, however, results in a control group that is very similar to the treatment group with at least a 94 percent improvement in the mean difference for each measure.

Figure 4 demonstrates that the parallel trends assumption is satisfied: although the treatment group has lower turnout rates in general, the gap between the treatment and control groups is largely constant between 2010 and 2016. Turnout in each year is measured as a function of voters registered in 2018, which partially explains why observed turnout is higher later in the period. Of course, some of the increase in turnout observed in later years in Figure 4 can be attributed to higher "real" turnout as a share of eligible citizens.



Figure 4: General Election Turnout for Treated and Control Voters, 2010 – 2018

The trends presented in Figure 4 offer preliminary visual corroboration of what I find at the 464 neighborhood level — namely, that 2018 turnout was not higher for voters in close contact 465 with formerly incarcerated individuals. Table 6 formalizes these trends into an ordinary least 466 squares regression.<sup>10</sup> A treatment dummy distinguishes treated from control voters. The 467 treatment dummy is interacted with another dummy identifying the 2018 election. Robust 468 standard errors are clustered at the level of the match (Abadie and Spiess 2020). Model 1 469 presents the model output without the other controls used for matching; model 2 includes 470 these covariates. 471

<sup>472</sup> In models 3 and 4 of Table 6 I consider the possibility that the negative spillover effects

<sup>&</sup>lt;sup>10</sup>Although the dependent variable here is binary — it takes the value 0 if a voter does not participate, and 1 if she does — the coefficients produced by logistic regressions in the difference-in-differences context are largely uninterpretable. I thus use a linear specification here. When the models are estimated using a logistic specification, the treatment effect is virtually identical.

of incarceration dissipate over time. In these models, the dummies indicating treatment
and the 2018 election are interacted with the number of years since the most recent release
of a household member from prison (*Years Since Latest Incarceration*, shortened to *Years Since* in interactions). Matched control observations are assigned the value associated with
their treated observation. Model 3 includes no other covariates, while model 4 includes the
matched variables.

Formerly incarcerated individuals who were released from prison many years ago may no longer live at the same address they reported when leaving prison. Models 5 – 8 therefore include only the treated individuals (and their matches) whose registration dates predate the latest prison release date of a household member, who we can be relatively sure lived with an incarcerated individual. The treatment effects in these models tell the same general story.

	All Matched Observations			Registration Date prior to Release Date				
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
2018	$\begin{array}{c} 0.094^{***} \\ (0.0004) \end{array}$	$\begin{array}{c} 0.094^{***} \\ (0.0004) \end{array}$	$\begin{array}{c} 0.095^{***} \\ (0.001) \end{array}$	$0.095^{***}$ (0.001)	$\begin{array}{c} 0.055^{***} \\ (0.0005) \end{array}$	$0.055^{***}$ (0.0005)	$\begin{array}{c} 0.081^{***} \\ (0.001) \end{array}$	$\begin{array}{c} 0.081^{***} \\ (0.001) \end{array}$
Treated	$-0.058^{***}$ (0.001)	$-0.060^{***}$ (0.001)	$-0.073^{***}$ (0.001)	$-0.075^{***}$ (0.001)	$-0.056^{***}$ (0.001)	$-0.064^{***}$ (0.001)	$-0.065^{***}$ (0.001)	$-0.068^{***}$ (0.001)
Years Since Latest Incarceration			$\begin{array}{c} 0.00000\\ (0.0001) \end{array}$	-0.00004 (0.0001)			$0.013^{***}$ (0.0001)	$0.003^{***}$ (0.0001)
2018 $\times$ Treated	$-0.022^{***}$ (0.001)	$-0.022^{***}$ (0.001)	$-0.038^{***}$ (0.001)	$-0.038^{***}$ (0.001)	$-0.033^{***}$ (0.001)	$-0.033^{***}$ (0.001)	$-0.048^{***}$ (0.002)	$-0.048^{***}$ (0.002)
2018 $\times$ Years Since			-0.0001 (0.0001)	-0.0001 (0.0001)			$-0.004^{***}$ (0.0001)	$-0.004^{***}$ (0.0001)
Treated $\times$ Years Since			$0.002^{***}$ (0.0001)	$0.002^{***}$ (0.0001)			$0.001^{***}$ (0.0002)	$0.001^{***}$ (0.0002)
2018 $\times$ Treated $\times$ Years Since			$0.002^{***}$ (0.0002)	$0.002^{***}$ (0.0002)			$0.002^{***}$ (0.0002)	$0.002^{***}$ (0.0002)
Constant	$\begin{array}{c} 0.478^{***} \\ (0.001) \end{array}$	$\begin{array}{c} 0.011^{***} \\ (0.004) \end{array}$	$\begin{array}{c} 0.478^{***} \\ (0.001) \end{array}$	$\begin{array}{c} 0.012^{***} \\ (0.004) \end{array}$	$\begin{array}{c} 0.575^{***} \\ (0.001) \end{array}$	$-0.047^{***}$ (0.005)	$\begin{array}{c} 0.494^{***} \\ (0.001) \end{array}$	$-0.059^{***}$ (0.005)
Includes covariates from matching Congressional District fixed effects		X X		X X		X X		X X
Observations R <sup>2</sup> Adjusted R <sup>2</sup>	7,388,640 0.008 0.008	$\overline{7,388,640}$ 0.199 0.199	7,388,640 0.009 0.009	$\overline{7,388,640}$ 0.199 0.199	4,915,920 0.005 0.005	4,915,920 0.157 0.157	$\begin{array}{r} 4,915,920\\ 0.023\\ 0.023\end{array}$	4,915,920 0.157 0.157

	Table 6:	General	Election	Turnout,	2010 -	2018
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\*\*\*p < 0.01, \*\*p < 0.05, \*p < 0.1. Robust standard errors (clustered at level of match) in paren-

Each model in Table 6 identifies a negative treatment effect. The coefficients on  $2018 \times$ 485 Treated in models 1 and 2 indicate that turnout among treated voters was about 2.2 per-486 centage points below what it would have been if the gap between treated and control voters 487 in 2018 had conformed to prior years. This mirrors the findings from the neighborhood-level 488 analyses, where the number of formerly incarcerated residents is not associated with higher 489 turnout. 490

There is some indication that spillover effects lessen with time. In each model,  $2018 \times Treated$ 491  $\times$  Years Since and Treated  $\times$  Years Since is positive and statistically significant. In other 492 words, individuals whose housemates had not been imprisoned for many years were more 493 likely to vote than other treated voters, and this was especially true in 2018. Models 3 and 4 494 estimate that the treatment effect for an individual whose household member returned from 495

<sup>496</sup> prison within one year of the election was about -3.8 percentage points. For each year the <sup>497</sup> most recent incarceration recedes into the past, the treatment effect decreases by about 0.2 <sup>498</sup> points in years other than 2018, and by 0.4 points in 2018. That the spillover effects "decay" <sup>499</sup> is a positive sign, and indicates that the negative socialization induced by a housemate's <sup>500</sup> incarceration might not be permanent.

It is unsurprising that the effect is moderated by time. Individuals whose household members 501 went to and were released from prison between the 2016 and 2018 elections, for instance, 502 received two treatments: they both were "negatively" treated by the incarceration of their 503 housemate and potentially "positively" treated by Amendment 4. What is surprising, how-504 ever, is the continued negative treatment effect even for the households furthest removed 505 from the incarceration of a household member. Table 7 presents the results of models 5 506 and 6 from Table 6, but limits the pool to households where someone last returned home 507 from prison prior to 2010. The "negative" treatment for these individuals should be reflected 508 in the base years of the difference-in-differences models. In these models,  $2018 \times Treated$ 509 remains significant and negative. The neighborhood-level analyses indicate that the amount 510 of time that has elapsed since an individual's incarceration is also related to support for and 511 the salience of Amendment 4; similar processes may be at play here, but the individual-level 512 data does not allow us to explore them. 513

	(1)	(2)
2018	0.031***	0.031***
	(0.001)	(0.001)
Treated	$-0.048^{***}$	$-0.057^{***}$
	(0.002)	(0.002)
$2018 \times \text{Treated}$	$-0.020^{***}$	-0.020***
	(0.002)	(0.002)
Constant	0.656***	-0.011
	(0.001)	(0.012)
Includes covariates from matching		X
Congressional District fixed effects		Х
Observations	$1,\!524,\!000$	1,524,000
$\mathbb{R}^2$	0.003	0.102
Adjusted $\mathbb{R}^2$	0.003	0.102

Table 7: General Election Turnout, 2010 – 2018

p < 0.01, p < 0.05, p < 0.1.

Robust standard errors (clustered at level of match) in parentheses.

These negative, statistically significant findings at the individual and neighborhood level 514 should probably not be interpreted to mean that Amendment 4 had a demobilizing effect 515 on individuals whose family and community members would be re-enfranchised by its pas-516 sage. Rather, it likely highlights that these individuals are less susceptible to other broadly 517 mobilizing phenomena. The 2018 election saw higher participation than any midterm in 518 a century as many infrequent voters turned out. It appears that voters whose household 519 members have been to prison were less mobilized by the factors that encouraged other de-520 mographically similar voters to participate in 2018. This analysis cannot determine whether 521 their indirect exposure to the criminal justice system caused this imperviousness, or if they 522 would have remained on the sidelines in 2018 even if their household members had not been 523 imprisoned. Nevertheless, that their turnout in 2018 did not increase relative to other voters 524 – even with Amendment 4 on the ballot – underscores just how difficult their political 525

<sup>526</sup> (re)integration is.

#### 527 Discussion and Conclusion

Turnout in 2018 hit historic levels for a midterm election as infrequent voters participated 528 and made their voices heard. In addition to hotly contested Congressional, senate, and 529 gubernatorial races, Floridians were presented with the opportunity to restore voting rights 530 to well over a million permanently disenfranchised individuals who had been convicted of 531 felony offenses. Amendment 4 and its organizers were hugely successful — in a year where 532 both statewide winners won by less than 0.5 percentage points, nearly two-thirds of Floridians 533 supported expanding the franchise. Neighborhoods and voters most directly impacted by 534 felony disenfranchisement gained meaningful political representation from the passage of the 535 amendment, and one of the "durable markers" of their civil death was nullified. However, I 536 fail to uncover evidence that Amendment 4 itself increased the turnout of neighborhoods and 537 individuals in close proximity to the formerly incarcerated above-and-beyond the increases 538 observed among other voters and in other communities. 539

It is not immediately apparent why Amendment 4 did not disproportionately heighten mobilization among these voters. The current study cannot tell whether it was an issue of lower political knowledge, or because the legal estrangement of the carceral state runs too deep for a single ballot initiative to overcome. However, if estrangement was the reason that the ballot initiative failed to mobilize these voters, this was likely only reinforced in the aftermath of the 2018 election. After the state constitution was amended to re-enfranchise their family members and neighbors, legislators rewrote the law to exclude them anew.

Just months after the 2018 election the Florida legislature passed a bill requiring disenfranchised individuals to pay off all court-ordered financial obligations before registering to vote, despite the fact that the state was incapable of determining how much any individual actually owed (Stern 2019). A federal judge ruled the law unconstitutional in May of 2020,

arguing that conditioning voting rights on the repayment of obligations that individuals can-551 not afford amounted to a poll tax and violation of the 24th Amendment.<sup>11</sup> That September. 552 however, an en bank ruling by the U.S. Court of Appeals for the 11th Circuit overturned 553 that decision,<sup>12</sup> upholding the constitutionality of the law. In his dissent from the Eleventh 554 Circuit's ruling, Judge Adalberto Jordan noted that "[h]ad Florida wanted to create a sys-555 tem to obstruct, impede, and impair the ability of felons to vote under Amendment 4, it 556 could not have come up with a better one" and that "Florida cannot tell felons — the great 557 majority of whom are indigent — how much they owe... and has come up with conflicting 558 (and uncodified) methods for determining how LFO [legal financial obligation] payments by 559 felons should be credited." That Florida legislators would condition voting on criteria that 560 cannot be verified, or cannot be afforded, has understandably been described as "unfair [and] 561 heartbreaking" by one disenfranchised individual who said the amendment had promised to 562 "give me a voice in my own future" (Harris 2020). It remains to be seen how such legis-563 lation and litigation will inform how criminal justice-involved individuals understand their 564 relationship with the state and structure their future democratic participation. 565

The results of this study point to the next chapter of the fight for political integration and rep-566 resentation for advocates in the Sunshine State. The relatively lower turnout in 2018 for the 567 communities most impacted by the carceral state indicates that formal re-enfranchisement is 568 not enough. If Floridian and American democracy wants to *actually* incorporate voices from 569 these communities — and not simply legally *allow* for their incorporation — the advocacy 570 movement cannot consider its work done once the formal barriers to the ballot box have been 571 torn down. Re-enfranchisement is clearly necessary, but it is not sufficient. Researchers must 572 continue exploring why the political re-incorporation of these communities is so difficult, and 573 organizers on the ground must do the hard work of reknitting them to our body politic. 574

 $<sup>^{11} \</sup>rm Jones$  et al. v. DeSantis et al., 4:19cv300-RH/MJF (U.S. District Court for the Northern District of Florida 2020).

 $<sup>^{12}</sup>$  Jones et al. v. DeSantis et al., 4:19cv300-RH/MJF (United States Court of Appeals for the Eleventh Circuit).

## 575 Declarations

The author affirms this research did not directly involve human subjects. It was conducted in accordance with the Brennan Center for Justice's Quantitative Research Protocol.

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<sup>580</sup> Research documentation and data that support the findings of this study are openly available

<sup>581</sup> in the APSR Dataverse at https://doi.org/10.7910/DVN/A81LPK. Limitations on data <sup>582</sup> availability are discussed in the text.

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# Supplementary Information

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# **Re-Estimation with Hillsborough County**

As discussed in the body of this article, statewide data on the residential addresses of individuals sentenced to felony probation are not available. These data are, however, available in Hillsborough County, the county in Florida with the third-highest number of formerly incarcerated individuals.<sup>1</sup> These records go back to 1988, though I have restricted them to individuals sentenced since October 1, 1997, so that they mirror the incarceration records. I follow the same geocoding and address cleaning procedures as for the incarceration records discussed above. These data do not include unique identifiers. To avoid double-counting, only the most recent record for each unique first name, middle name, last name, and date of birth is retained. This potentially excludes different people whose names and dates of birth are identical. Individuals whose adjudication was withheld are excluded, as are individuals whose names, dates of birth, and addresses match individuals who were formerly incarcerated. This avoids double counting individuals both incarcerated and sentenced to probation.

<sup>&</sup>lt;sup>1</sup>See https://www.hillsclerk.com/Records-and-Reports/Public-Data-Files.

Figure 1 plots the relationship between the number of formerly incarcerated residents and residents who have been sentenced to felony probation in each block group in Hillsborough County (scaled by population). As the figure makes clear, individuals who have been sentenced to felony probation are concentrated in the same neighborhoods where individuals live after a period of incarceration (the  $R^2$  of the bivariate regression is 0.92). As with the marginal effects plots in the body of this article, the figure does not show outlier neighborhoods but the line of best fit and  $R^2$  are calculated using all observations.



Figure 1: Relationship Between Formerly Incarcerated and Probationed Residents, Hillsborough County

Table 1 replicates the models from Tables 3 and 4 in the main body of this article. In each pair of models in the table, I begin by re-fitting the exact models presented in the body of this article but limiting the precincts and block groups to Hillsborough County. In the second model in each pair, the primary dependent variable includes both formerly incarcerated residents *and* the number of residents who have been convicted of a felony probation.

	Precinct-Level Turnout		Block Group-Level Turnout		Am. 4 Support		Roll-off	
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Formerly Incarcerated Residents	$\begin{array}{c} 0.00002 \\ (0.00004) \end{array}$		$-0.0002^{***}$ (0.00001)		-0.00003 (0.00004)		$-0.00005^{***}$ (0.00001)	
Total Disenfranchised Individuals		-0.00000 (0.00001)		$-0.0001^{***}$ (0.00001)		-0.00001 (0.00001)		$-0.00002^{***}$ (0.00000)
Percent White	$-0.528^{*}$ (0.316)	-0.514 (0.324)	$\begin{array}{c} 0.013\\ (0.011) \end{array}$	$\begin{array}{c} 0.013 \\ (0.011) \end{array}$	$\begin{array}{c} 0.124 \\ (0.491) \end{array}$	$\begin{array}{c} 0.114 \\ (0.492) \end{array}$	$\begin{array}{c} 0.029 \\ (0.039) \end{array}$	$\begin{array}{c} 0.025 \\ (0.038) \end{array}$
Percent Black	$-0.690^{***}$ (0.227)	$-0.669^{***}$ (0.239)	$0.006 \\ (0.006)$	$0.007 \\ (0.005)$	$\begin{array}{c} 0.122\\ (0.442) \end{array}$	$\begin{array}{c} 0.107 \\ (0.443) \end{array}$	$\begin{array}{c} 0.012\\ (0.071) \end{array}$	$\begin{array}{c} 0.005 \\ (0.070) \end{array}$
Percent Latino	$-0.721^{**}$ (0.296)	$-0.708^{**}$ (0.302)	$-0.039^{***}$ (0.012)	$-0.040^{***}$ (0.011)	-0.043 (0.442)	-0.052 (0.443)	$\begin{array}{c} 0.017\\ (0.036) \end{array}$	$\begin{array}{c} 0.013\\ (0.035) \end{array}$
Percent Asian	-0.560 (0.408)	-0.547 (0.412)	$0.046^{*}$ (0.024)	$0.044^{*}$ (0.024)	-0.076 (0.543)	-0.085 (0.545)	$0.098 \\ (0.077)$	0.093 (0.077)
Percent Male	$\begin{array}{c} 0.386\\ (0.343) \end{array}$	$\begin{array}{c} 0.370 \\ (0.351) \end{array}$	$\begin{array}{c} 0.217^{***} \\ (0.040) \end{array}$	$\begin{array}{c} 0.224^{***} \\ (0.042) \end{array}$	-0.174 (0.315)	-0.162 (0.320)	$-0.149^{**}$ (0.061)	$-0.142^{**}$ (0.062)
Percent Democrats	$\begin{array}{c} 0.497^{***} \\ (0.121) \end{array}$	$\begin{array}{c} 0.499^{***} \\ (0.121) \end{array}$	$0.117^{**}$ (0.054)	$\begin{array}{c} 0.114^{**} \\ (0.051) \end{array}$	$\begin{array}{c} 0.121 \\ (0.165) \end{array}$	$0.120 \\ (0.166)$	$\begin{array}{c} 0.155\\ (0.145) \end{array}$	$0.156 \\ (0.147)$
Percent Republicans	$\begin{array}{c} 0.395^{***} \\ (0.076) \end{array}$	$\begin{array}{c} 0.398^{***} \\ (0.077) \end{array}$	$\begin{array}{c} 0.051 \\ (0.033) \end{array}$	$\begin{array}{c} 0.047\\ (0.031) \end{array}$	$-0.851^{***}$ (0.077)	$-0.853^{***}$ (0.079)	$0.142 \\ (0.122)$	$0.140 \\ (0.123)$
Average Age	-0.003 (0.002)	-0.003 (0.002)	-0.001 (0.001)	-0.001 (0.001)	$0.001 \\ (0.001)$	$\begin{array}{c} 0.001 \\ (0.001) \end{array}$	$0.002^{***}$ (0.001)	$0.002^{***}$ (0.001)
Average Income (\$10,000s)	$59.268^{***}$ (21.581)	$58.970^{***}$ (21.417)	$0.001^{**}$ (0.001)	$0.001^{**}$ (0.001)	-3.836 (8.885)	-3.568 (9.017)	2.738 (4.638)	3.004 (4.688)
Percent With Some College	$\begin{array}{c} 0.127^{***} \\ (0.011) \end{array}$	$\begin{array}{c} 0.124^{***} \\ (0.010) \end{array}$	$0.066^{***}$ (0.010)	$0.063^{***}$ (0.010)	$0.088^{*}$ (0.047)	$0.091^{*}$ (0.047)	-0.009 (0.014)	-0.009 (0.013)
Percent Unemployed	$-0.175^{***}$ (0.029)	$-0.170^{***}$ (0.029)	-0.019 (0.015)	-0.017 (0.014)	$-0.117^{*}$ (0.066)	$-0.120^{*}$ (0.066)	$0.064 \\ (0.040)$	$\begin{array}{c} 0.065 \\ (0.040) \end{array}$
Constant	-0.024 (0.116)	-0.020 (0.119)	$-0.223^{**}$ (0.090)	$-0.221^{**}$ (0.088)	$\begin{array}{c} 0.883^{***} \\ (0.169) \end{array}$	$0.880^{***}$ (0.168)	-0.051 (0.037)	-0.052 (0.037)
Congressional District FEs Turnout in 2010 – 2016	X X	X X	X X	X X	X X	X X	X X	X X
Observations $\mathbf{R}^2$ Adjusted $\mathbf{R}^2$	$390 \\ 0.881 \\ 0.875$	390 0.881 0.874	812 0.976 0.975	812 0.976 0.975	$390 \\ 0.944 \\ 0.941$	$390 \\ 0.944 \\ 0.941$	$390 \\ 0.483 \\ 0.455$	390 0.482 0.454

Table 1: Neighborhood Turnout, Support for Am. 4, and Roll-Off in 2018

 $^{***}p<0.01,\,^{**}p<0.05,\,^*p<0.1.$  Robust standard errors (clustered by congressional district) in parentheses.

The relationship between disenfranchised residents and precinct-level support for Amendment 4, and precinct-level turnout, are nonsignificant in Table 1 despite being significant statewide. Block group-level turnout and roll-off remain negatively associated with the presence of disenfranchised individuals. Importantly, in no model does moving from measuring only formerly incarcerated individuals to measuring all disenfranchised individuals change the sign on a statistically significant relationship. This provides corroboration for the argument that the neighborhood-level results presented in the body of this article, measured using only formerly incarcerated residents, apply to the formerly disenfranchised population more generally.

I next interrogate whether the use of only incarceration records is likely impacting the individual-level analyses presented in the body of the article. I re-run the matching procedure described above, where a registered voter is considered treated if they lived with *any* disenfranchised individual. Potential controls for this matching procedure are limited to Hillsborough County, where we can be sure registered voters do not live with individuals sentenced to felony probation. The matching procedure is successful at reducing differences between treated and control voters in Hillsborough County.

In Table 2, models 1 - 4 re-estimate models 1 - 4 from Table 6 from the main paper, where the pool is limited to treated voters who live in Hillsborough County and their matches. Models 5 - 8 present the results using the broader treatment definition.

	Lives with Formerly Incarcerated				Lives with Disenfranchised			
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
2018	$0.094^{***}$ (0.001)	$0.094^{***}$ (0.001)	$0.096^{***}$ (0.002)	$0.096^{***}$ (0.002)	$\begin{array}{c} 0.104^{***} \\ (0.001) \end{array}$	$\begin{array}{c} 0.104^{***} \\ (0.001) \end{array}$	$\begin{array}{c} 0.104^{***} \\ (0.002) \end{array}$	$\begin{array}{c} 0.104^{***} \\ (0.002) \end{array}$
Treated	$-0.060^{***}$ (0.002)	$-0.063^{***}$ (0.002)	$-0.071^{***}$ (0.004)	$-0.072^{***}$ (0.004)	$-0.066^{***}$ (0.001)	$-0.066^{***}$ (0.001)	$-0.079^{***}$ (0.002)	$-0.078^{***}$ (0.002)
Years Since Latest Incarceration			$0.001^{**}$ (0.0003)	$0.0005^{**}$ (0.0002)			$0.001^{***}$ (0.0002)	$0.001^{***}$ (0.0001)
2018 $\times$ Treated	$-0.021^{***}$ (0.003)	$-0.021^{***}$ (0.003)	$-0.037^{***}$ (0.005)	$-0.037^{***}$ (0.005)	$-0.029^{***}$ (0.002)	$-0.029^{***}$ (0.002)	$-0.047^{***}$ (0.003)	$-0.047^{***}$ (0.003)
2018 $\times$ Years Since			-0.0002 (0.0002)	-0.0002 (0.0002)			-0.00004 (0.0002)	-0.00004 (0.0002)
Treated $\times$ Years Since			$0.001^{***}$ (0.0004)	$0.001^{***}$ (0.0004)			$0.002^{***}$ (0.0003)	$0.001^{***}$ (0.0002)
2018 $\times$ Treated $\times$ Years Since			$0.002^{***}$ (0.001)	$0.002^{***}$ (0.001)			$0.002^{***}$ (0.0003)	$0.002^{***}$ (0.0003)
Constant	$\begin{array}{c} 0.449^{***} \\ (0.002) \end{array}$	$0.038 \\ (0.026)$	$\begin{array}{c} 0.444^{***} \\ (0.003) \end{array}$	$0.035 \\ (0.026)$	$\begin{array}{c} 0.441^{***} \\ (0.001) \end{array}$	$0.073^{***}$ (0.018)	$\begin{array}{c} 0.432^{***} \\ (0.002) \end{array}$	$\begin{array}{c} 0.070^{***} \\ (0.018) \end{array}$
Includes covariates from matching Congressional District fixed effects		X X		X X		X X		X X
Observations R <sup>2</sup> Adjusted R <sup>2</sup>	650,250 0.009 0.009	650,250 0.213 0.213	$ \begin{array}{c} 650,250\\ 0.009\\ 0.009 \end{array} $	$\begin{array}{c} 650,\!250 \\ 0.214 \\ 0.214 \end{array}$	$\begin{array}{c} 1,410,870 \\ 0.011 \\ 0.011 \end{array}$	$\begin{array}{c} 1,410,870 \\ 0.211 \\ 0.211 \end{array}$	$1,410,870 \\ 0.012 \\ 0.012$	$\begin{array}{c} 1,410,870 \\ 0.212 \\ 0.212 \end{array}$

Table 2:	General	Election	Turnout,	2010 -	2018
			,		

 $^{***}p < 0.01, \,^{**}p < 0.05, \,^{*}p < 0.1.$  Robust standard errors (clustered at level of match) in paren-

theses.

In Hillsborough County, the magnitude of the treatment effect grows when we broaden the treatment group to include anyone who lives with a formerly disenfranchised individual. This raises interesting questions about the potential differential spillover effects of living with a formerly incarcerated individual versus with an individual sentenced to felony probation. This may also be due to some housemates of probationed individuals serving as controls in the main analysis, collapsing the distinction between treated and control and producing conservative estimates. Nonetheless, Table 2 provides evidence that the negative treatment effects identified among voters living with formerly incarcerated individuals in the body of this article are likely generalizable to all voters living with disenfranchised individuals.

#### **Re-Estimation with All Formerly Incarcerated Individuals**

When discussing the impact of formerly incarcerated residents on neighborhood turnout and support for Amendment 4 in the body of this paper, I include only a subset of formerly incarcerated residents. I exclude individuals who returned from prison to institutions listed by four or more other formerly incarcerated individuals. I choose to exclude these individuals because I am most interested in the relationship between Amendment 4 and the turnout of individuals in proximal contact with the criminal justice system. Walker and García-Castañon (2017) defines proximal contact "as having a loved one who is a custodial citizen without yourself having had contact" (542). Because much of the literature focuses on the mechanisms linking personal relationships, proximal contact, and political participation, I limit the sample to formerly incarcerated individuals who are likely returning to neighborhoods with social and familial ties.

Nevertheless, living in a neighborhood with a large number of formerly incarcerated individuals who reside in institutions like half-way houses or shelters might structure voting behavior. Here I re-estimate the models presented in Tables 3 and 4 in the body of this paper, but now including *all* formerly incarcerated residents. Table 3 presents the results of these estimations. Model 1 presents the turnout regression estimated at the block group level, while Models 2 - 4 are estimated using precinct level data.

	Block Group		Precinct	
	Turnout	Turnout	Support for Am. 4	Roll-Off
	(1)	(2)	(3)	(4)
Formerly Incarcerated Residents	$-0.0001^{***}$ (0.00001)	$-0.00004^{***}$ (0.00001)	0.00003*** (0.00001)	$\begin{array}{c} -0.00001^{***} \\ (0.00000) \end{array}$
Percent White	$0.020^{**}$ (0.008)	$0.004 \\ (0.036)$	$0.072^{*}$ (0.041)	$-0.074^{***}$ (0.015)
Percent Black	$\begin{array}{c} 0.040^{***} \\ (0.008) \end{array}$	-0.005 (0.036)	$\begin{array}{c} 0.196^{***} \\ (0.041) \end{array}$	$-0.049^{***}$ (0.015)
Percent Latino	-0.005 (0.008)	$-0.091^{**}$ (0.036)	$0.052 \\ (0.041)$	$-0.052^{***}$ (0.015)
Percent Asian	$\begin{array}{c} 0.046^{***} \\ (0.011) \end{array}$	$0.092^{*}$ (0.052)	$\begin{array}{c} 0.243^{***} \\ (0.059) \end{array}$	$-0.099^{***}$ (0.021)
Percent Male	$0.092^{***}$ (0.023)	$\begin{array}{c} 0.319^{***} \\ (0.055) \end{array}$	$-0.389^{***}$ (0.063)	$-0.200^{***}$ (0.023)
Percent Democrats	$0.063^{***}$ (0.008)	$0.067^{***}$ (0.020)	$\begin{array}{c} 0.191^{***} \\ (0.023) \end{array}$	$0.033^{***}$ (0.008)
Percent Republicans	$0.006 \\ (0.008)$	$0.023 \\ (0.019)$	$-0.397^{***}$ (0.021)	$0.041^{***}$ (0.008)
Average Age	$0.001^{***}$ (0.0001)	0.00005 (0.0002)	-0.0003 (0.0002)	$0.001^{***}$ (0.0001)
Average Income (\$10,000s)	$0.002^{***}$ (0.0001)	$0.002^{***}$ (0.0004)	$-0.003^{***}$ (0.0004)	-0.00002 (0.0002)
Percent With Some College	$\begin{array}{c} 0.086^{***} \\ (0.003) \end{array}$	$0.196^{***}$ (0.008)	$\begin{array}{c} 0.151^{***} \\ (0.010) \end{array}$	$-0.027^{***}$ (0.003)
Percent Unemployed	-0.006 (0.005)	$-0.039^{**}$ (0.018)	-0.014 (0.021)	$-0.020^{***}$ (0.007)
Constant	$-0.189^{***}$ (0.023)	$-0.236^{***}$ (0.049)	$\frac{1.030^{***}}{(0.056)}$	$\begin{array}{c} 0.216^{***} \\ (0.020) \end{array}$
Congressional District FEs Turnout in 2010 – 2016	X X	X X	X X	X X
Observations $R^2$ Adjusted $R^2$	$10,817 \\ 0.979 \\ 0.979$	5,797 0.779 0.777	5,797 0.788 0.786	5,797 0.312 0.307

 Table 3:
 Including All Formerly Incarcerated Residents

 $^{***}p<0.01,\,^{**}p<0.05,\,\,^*p<0.1.$  Robust standard errors (clustered by congressional district) in parentheses.

The inclusion of all formerly incarcerated residents substantially shrinks the size of the estimated coefficients of interest with respect to the estimates presented in the body of the article. Nevertheless, turnout (measured at the block group and precinct level) and roll-off are significantly and negatively related with the formerly incarcerated population in a neighborhood, and support for Amendment 4 remains positively (and significantly) related.

#### **Re-Estimation with Recently Released Individuals**

The body of the article also acknowledges that the use of release plan address data may be unreliable considering the fact that many individuals may have moved or died since their discharge from parole. This is especially possible for individuals who have not had contact with the state incarceration agency for many years. To account for this possibility, Table 4 re-estimates the models presented in Tables 3 and 4 from the main paper, but limits the formerly incarcerated individuals to those residents who were last released from prison between 2015 and the 2018 election. These individuals are the least likely to have died or moved, simply because their information is the most recent. These models include only individuals who returned to non-institutions, as presented in the body of the article.

	Block Group		Precinct	
	Turnout	Turnout	Support for Am. 4	Roll-Off
	(1)	(2)	(3)	(4)
Formerly Incarcerated Residents	-0.001***	-0.001***	0.0002***	-0.0001***
·	(0.0001)	(0.0001)	(0.0001)	(0.00002)
Percent White	$0.019^{**}$	$-0.142^{***}$	-0.024	$-0.028^{**}$
	(0.009)	(0.035)	(0.033)	(0.014)
Percent Black	0.040***	$-0.131^{***}$	0.069**	-0.011
	(0.009)	(0.035)	(0.033)	(0.014)
Percent Latino	-0.007	$-0.238^{***}$	$-0.083^{**}$	-0.005
	(0.009)	(0.034)	(0.033)	(0.014)
Percent Asian	0.045***	-0.096	0.150**	-0.012
	(0.012)	(0.062)	(0.059)	(0.025)
Percent Male	0.041	0.392***	$-0.285^{***}$	$-0.155^{***}$
	(0.026)	(0.059)	(0.056)	(0.024)
Percent Democrats	0.073***	0.182***	0.088***	0.043***
	(0.009)	(0.022)	(0.021)	(0.009)
Percent Republicans	0.006	0.118***	$-0.533^{***}$	0.043***
	(0.009)	(0.021)	(0.020)	(0.008)
Average Age	0.001***	$0.0003^{*}$	0.0002	0.001***
	(0.0001)	(0.0002)	(0.0002)	(0.0001)
Average Income (\$10,000s)	0.002***	0.002***	$-0.002^{***}$	-0.0001
	(0.0002)	(0.0004)	(0.0004)	(0.0002)
Percent With Some College	0.081***	0.163***	$0.161^{***}$	$-0.030^{***}$
	(0.003)	(0.008)	(0.007)	(0.003)
Percent Unemployed	0.0001	$-0.028^{*}$	$-0.040^{***}$	-0.0002
	(0.005)	(0.016)	(0.015)	(0.006)
Constant	$-0.148^{***}$	$-0.268^{***}$	$1.104^{***}$	0.114***
	(0.026)	(0.053)	(0.050)	(0.021)
Congressional District FEs	Х	Х	Х	Х
Turnout in 2010 – 2016	Х	Х	Х	Х
Observations	8,967	4,905	4,905	4,905
$\mathbb{R}^2$	0.979	0.839	0.897	0.407
Adjusted $R^2$	0.979	0.837	0.896	0.401

 Table 4:
 Formerly Incarcerated Residents Released Since 1/1/2015

 $^{***}p<0.01,\,^{**}p<0.05,\,^*p<0.1.$  Robust standard errors (clustered by congressional district) in parentheses.

In each of the models presented in Table 4, the independent variable of interest is statistically significant at the 99 percent level. Moreover, the estimated coefficient is in each case larger than that presented in the body of the article. This could be because using more recent data better identifies communities that are currently home, not just historically home, to formerly incarcerated individuals. On the other hand, the primary analyses in this article indicate that a community member's incarceration may be more salient in places where residents were more recently incarcerated. Proximal contact, in other words, might shape voters' behavior more strongly if that contact was recent. The "decaying" spillover effects in the individual-level difference-in-differences regressions presented later in the paper would seem to corroborate this as well.

#### References

Walker, Hannah L., and Marcela García-Castañon. 2017. "For Love and Justice: The Mobilizing of Race, Gender, and Criminal Justice Contact." *Politics & Gender* 13 (4): 541–68. https://doi.org/10.1017/S1743923X17000198.